

Evaluating and Selecting a Bayesian MCMC Model

Casualty Actuarial Society Spring 2023 Meeting

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Goals of Presentation

- Demonstrate steps in model selection
 - Exploratory Data Analysis
 - Identify potential model forms
 - Build alternative models
 - Check integrity of model estimates
 - Evaluate model fit
 - Compare predictive power
- Highlight software that makes this practical
 - Ggplot2
 - STAN
 - Brms
 - Bayesplot & ShinyStan:
 - Tidybayes & ggdist
 - Loo and loo_compare

Modeling Environment

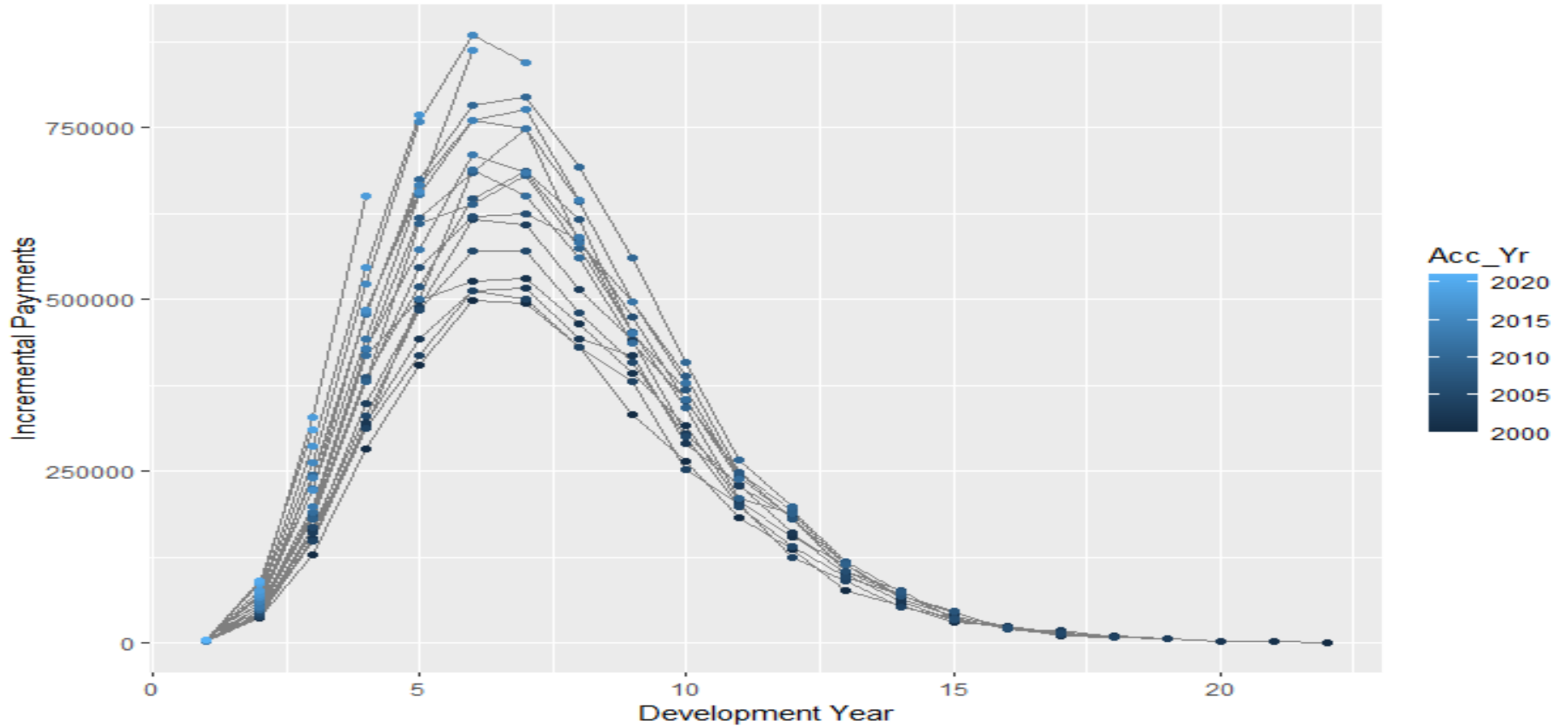


Exploratory Data Analysis

Exploratory Data Analysis

- Goals for this presentation
 - Display patterns as first step to writing formulas for models
 - Information on likely predictive variable distribution
- Additional steps for real life analysis
 - Search for anomalous behavior in claim handling by region or business unit
 - Look for changes in behavior in claims activity over time caused by change in underwriting practice
 - Check for claim recoding effects
- Data Source
 - Simulated data using Poisson for counts & Lognormal for severity
 - Underlying distributions displayed in appendix

Plot of Development Year Incremental Payments By Accident Year

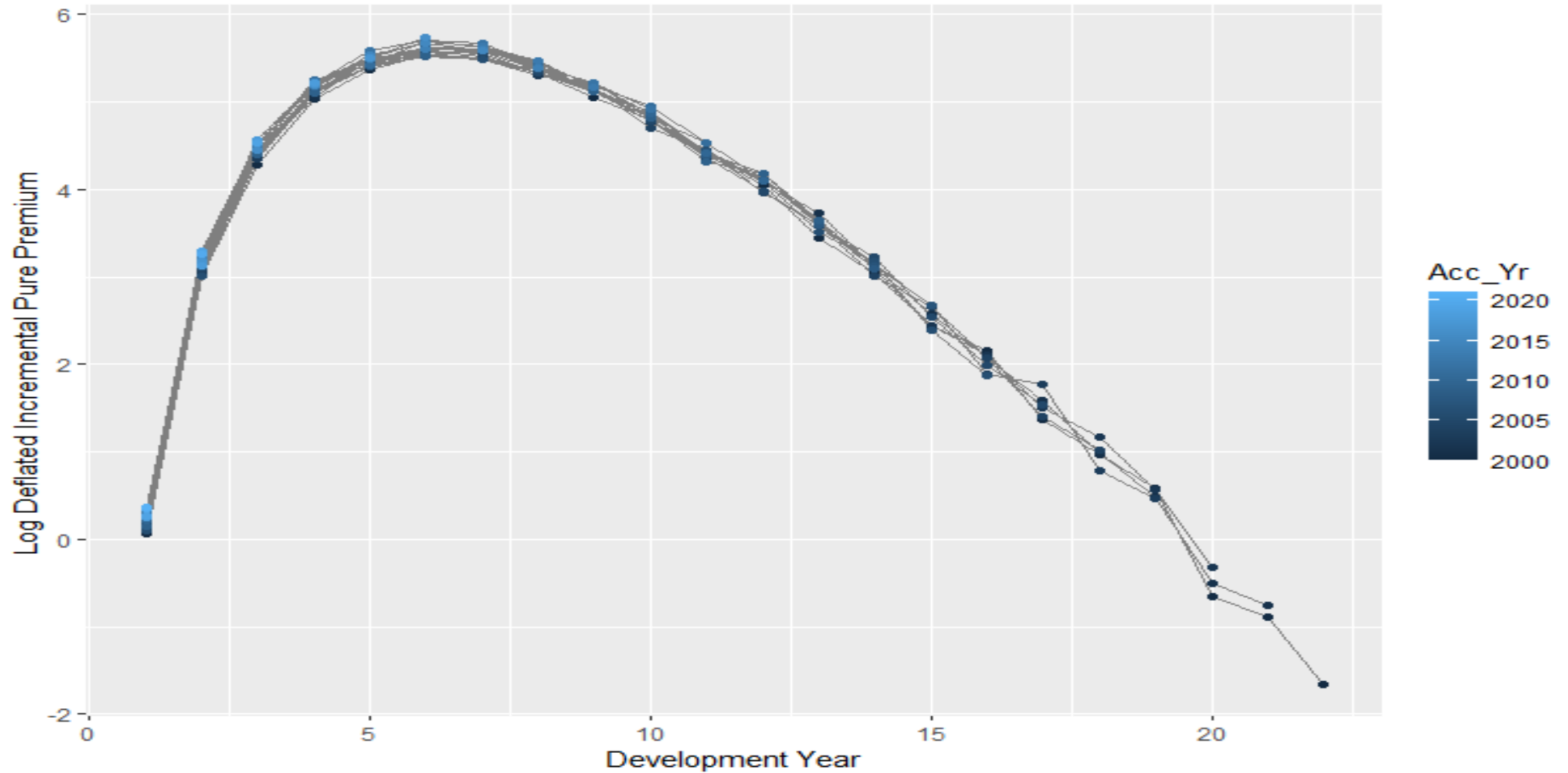


Ggplot2 Example

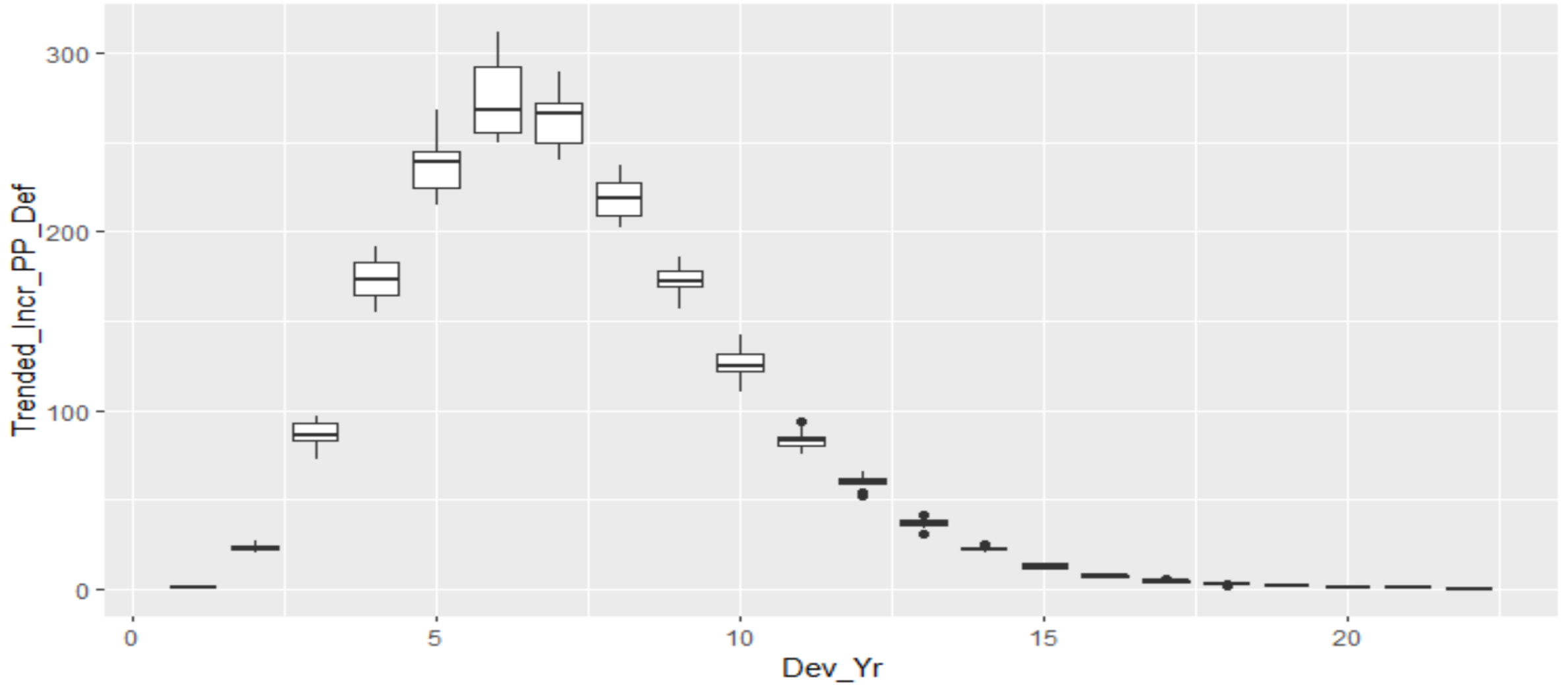
The screenshot displays the RStudio interface with the following components:

- Source Editor:** Contains R code for creating three ggplot2 plots. The first plot, titled "Plot of Development Year Incremental Payments \n By Accident Year", uses `geom_line` (grey50), `geom_point` (grey50), and `labs` to create a line plot of incremental payments by development year, grouped by accident year (Acc_Yr).
- Console:** Shows the execution of the code, including the loading of the `shiny` package and the execution of the `ggplot` function.
- Environment/Plots Panel:** Displays the resulting plot, titled "Plot of Development Year Incremental Payments By Accident Year". The plot shows multiple lines representing different accident years (2000, 2005, 2010, 2015, 2020) plotted against Development Year (x-axis, 0 to 20) and Incremental Payments (y-axis, 0 to 750,000). The lines show a peak around development year 5-10 and then decline.

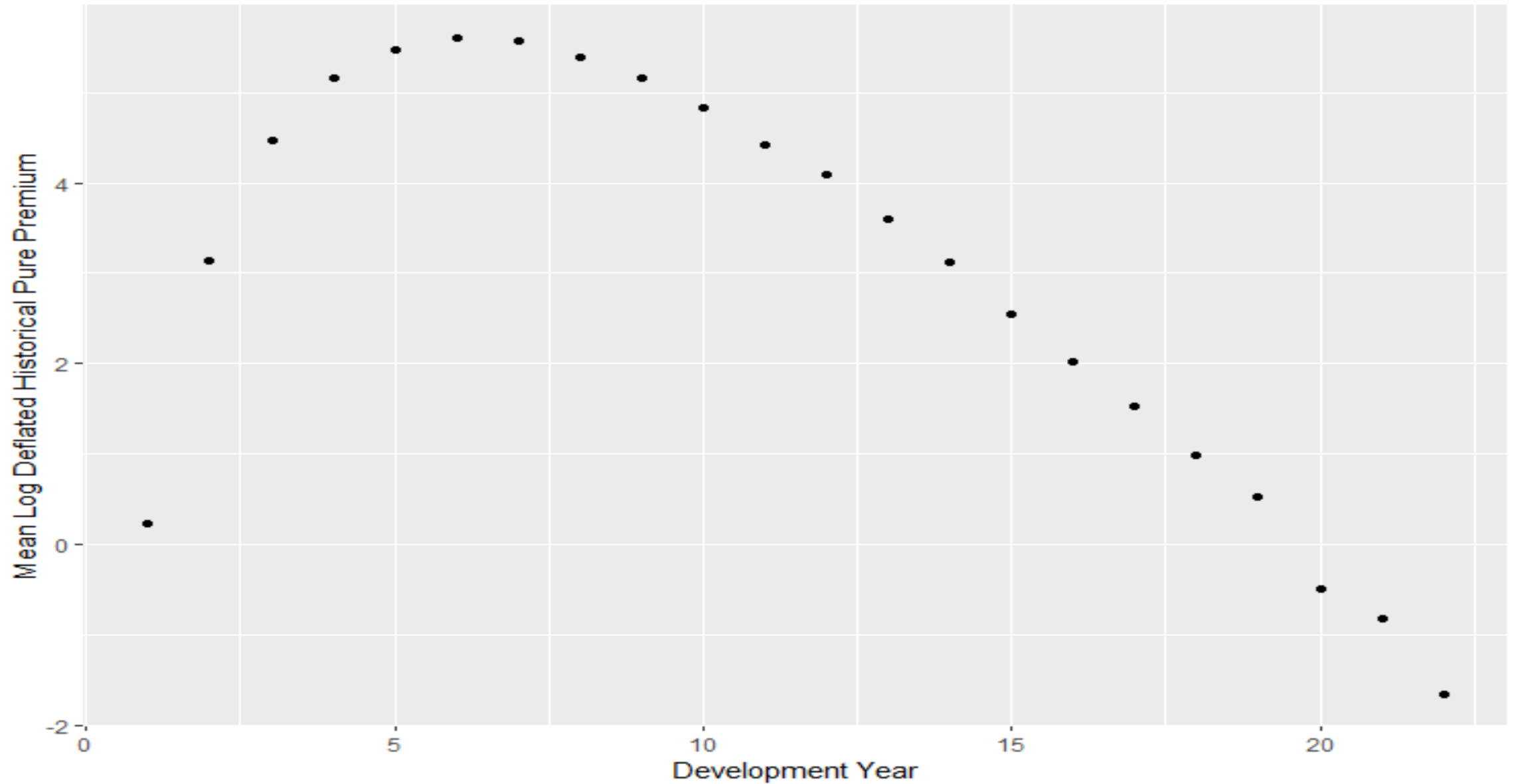
Plot of Development Year Log Deflated Incremental Pure Premium
By Accident Year



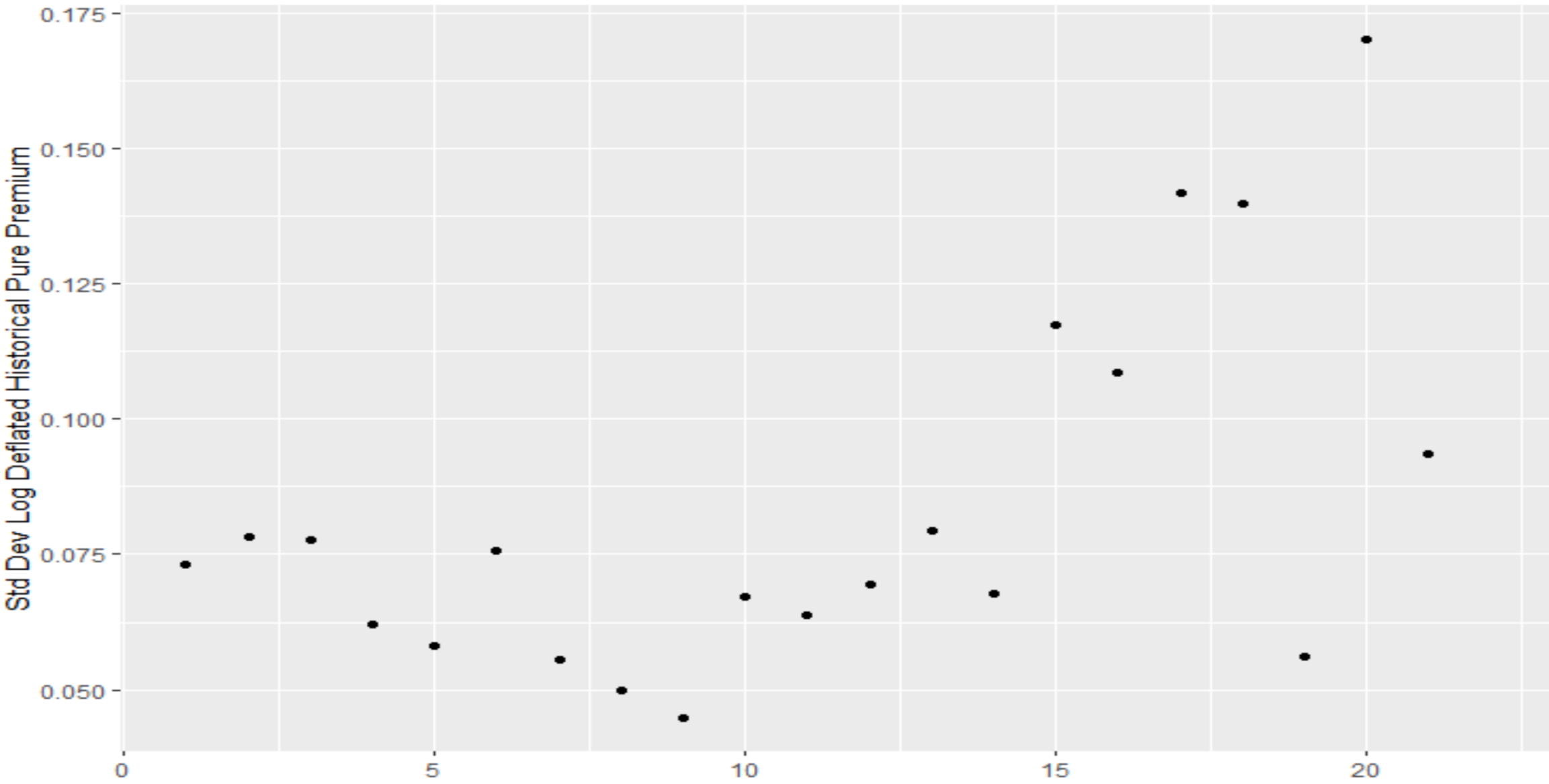
Box Plot of Deflated Incremental Pure Premium Across Accident Year



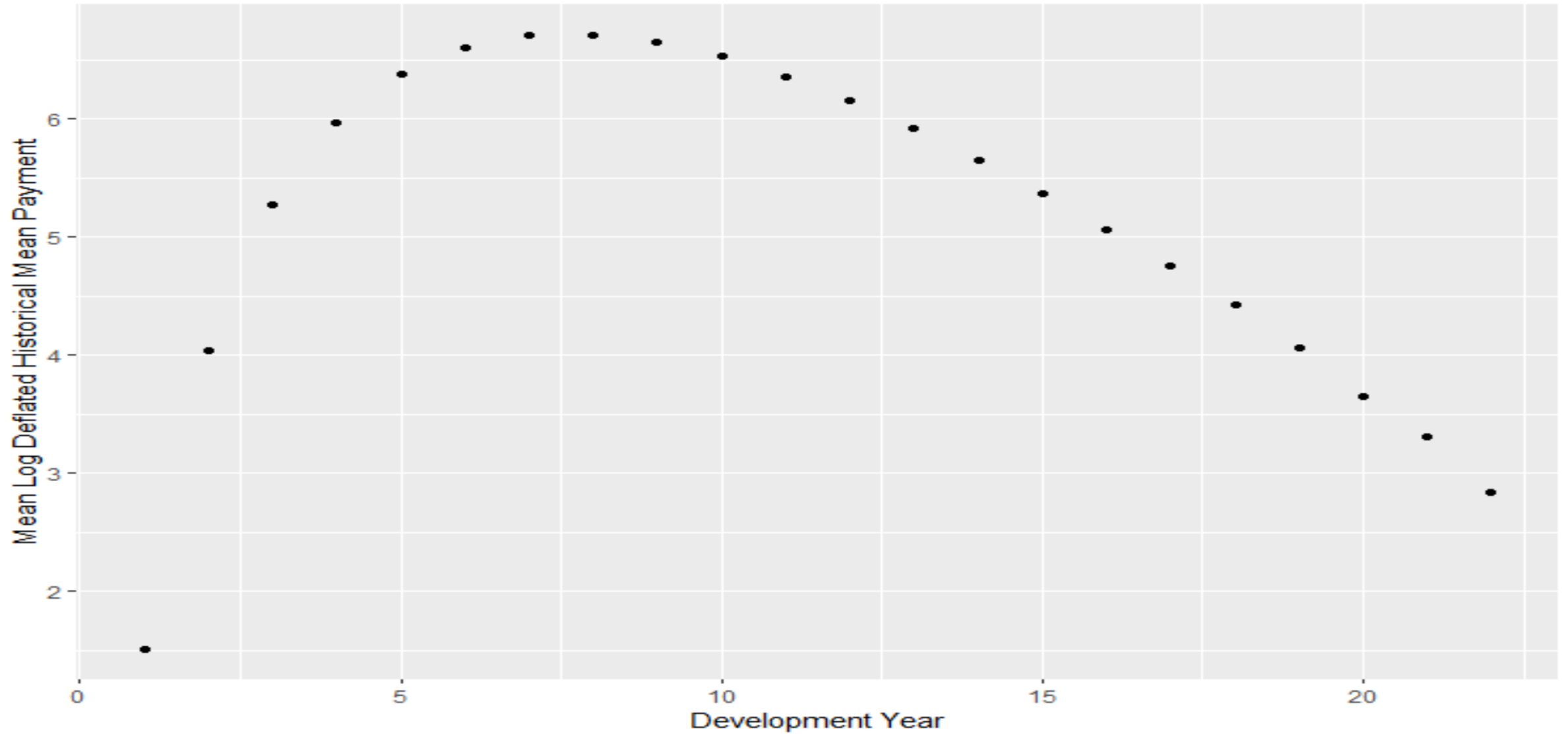
Mean Natural Log of Incremental PP Deflated Across Accident Years



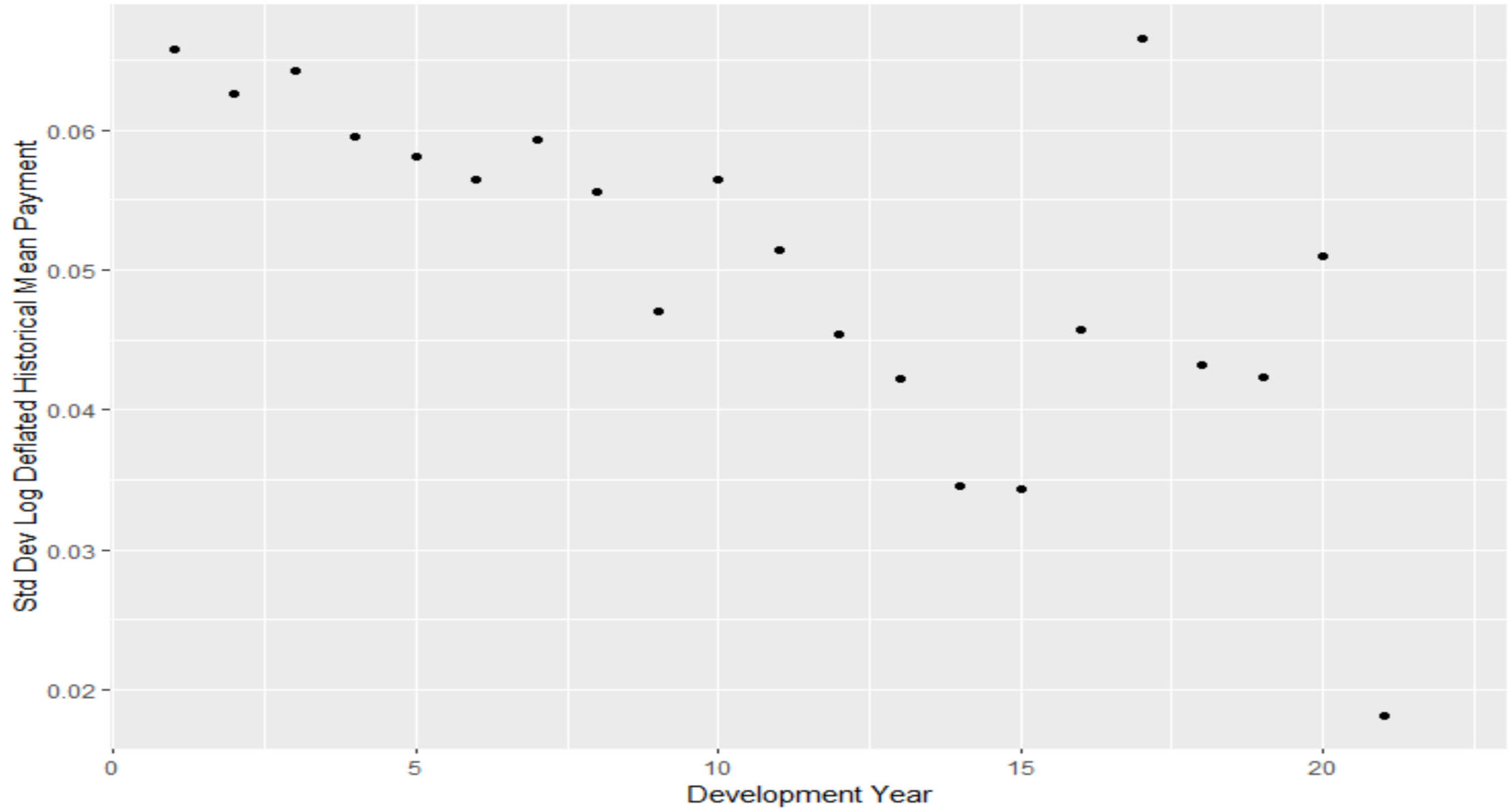
Standard Deviation Natural Log of Incremental PP Deflated Across Accident Years



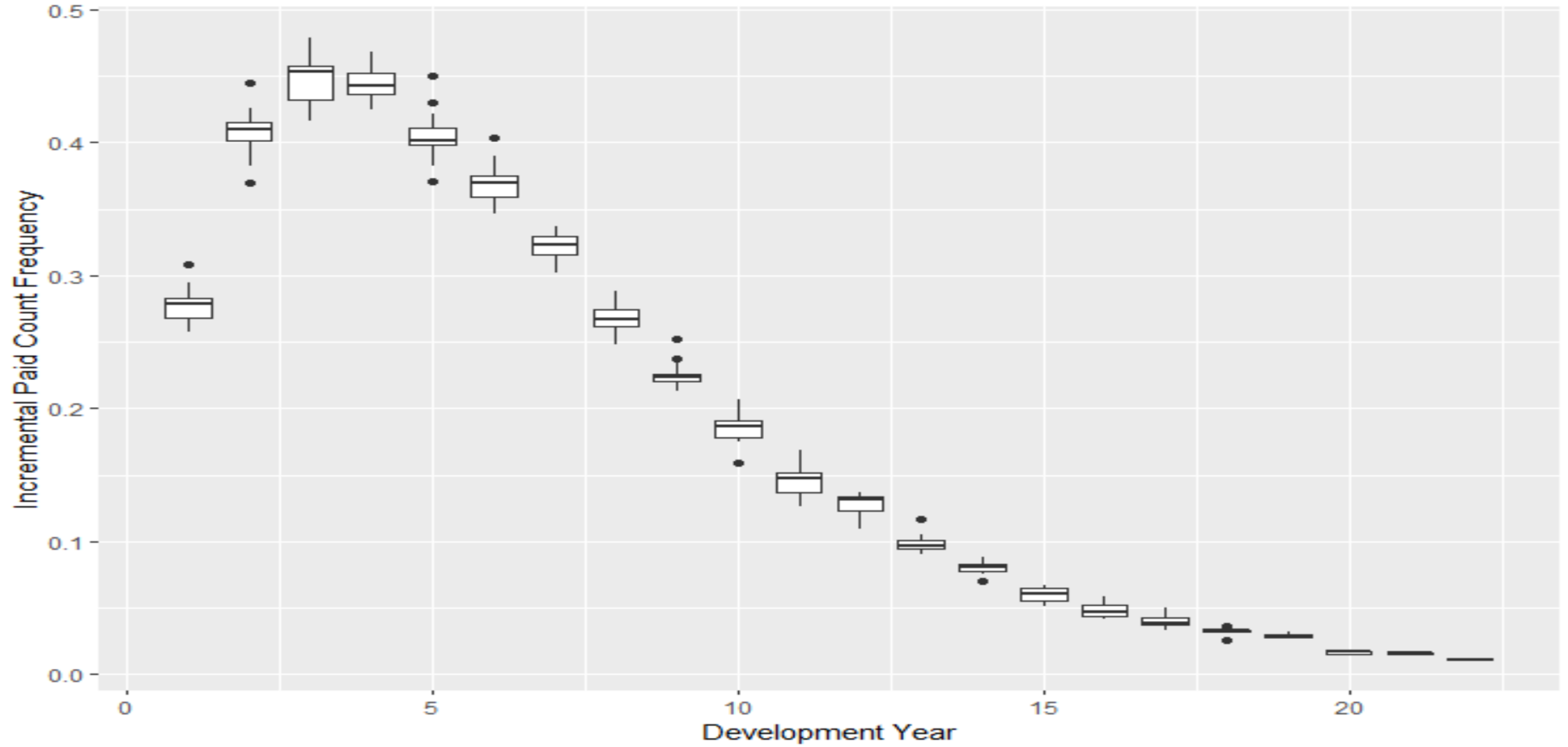
Mean Natural Log of Mean Payment Deflated Across Accident Years



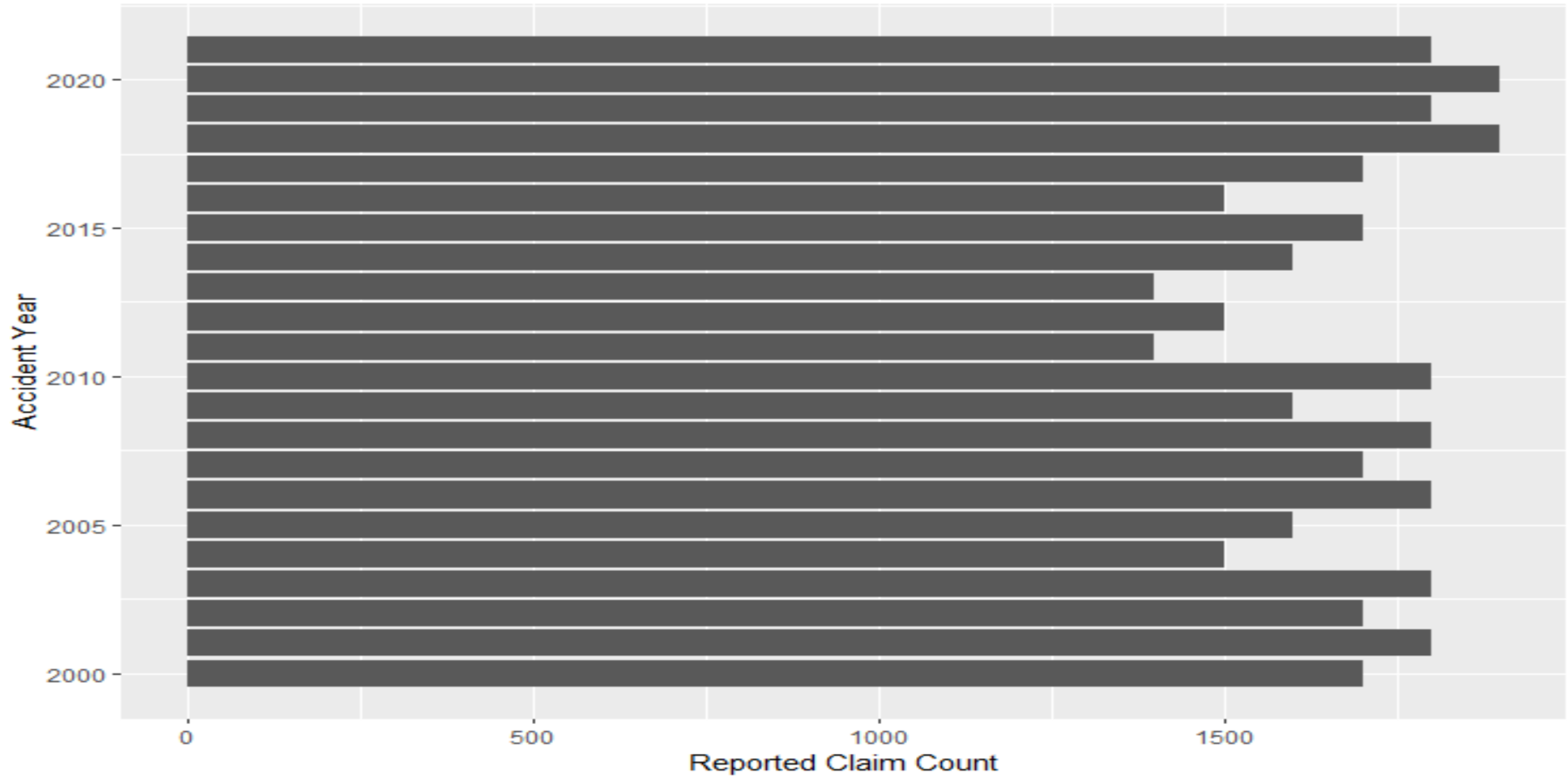
Standard Deviation Natural Log of Mean Payment Deflated Across Accident Years



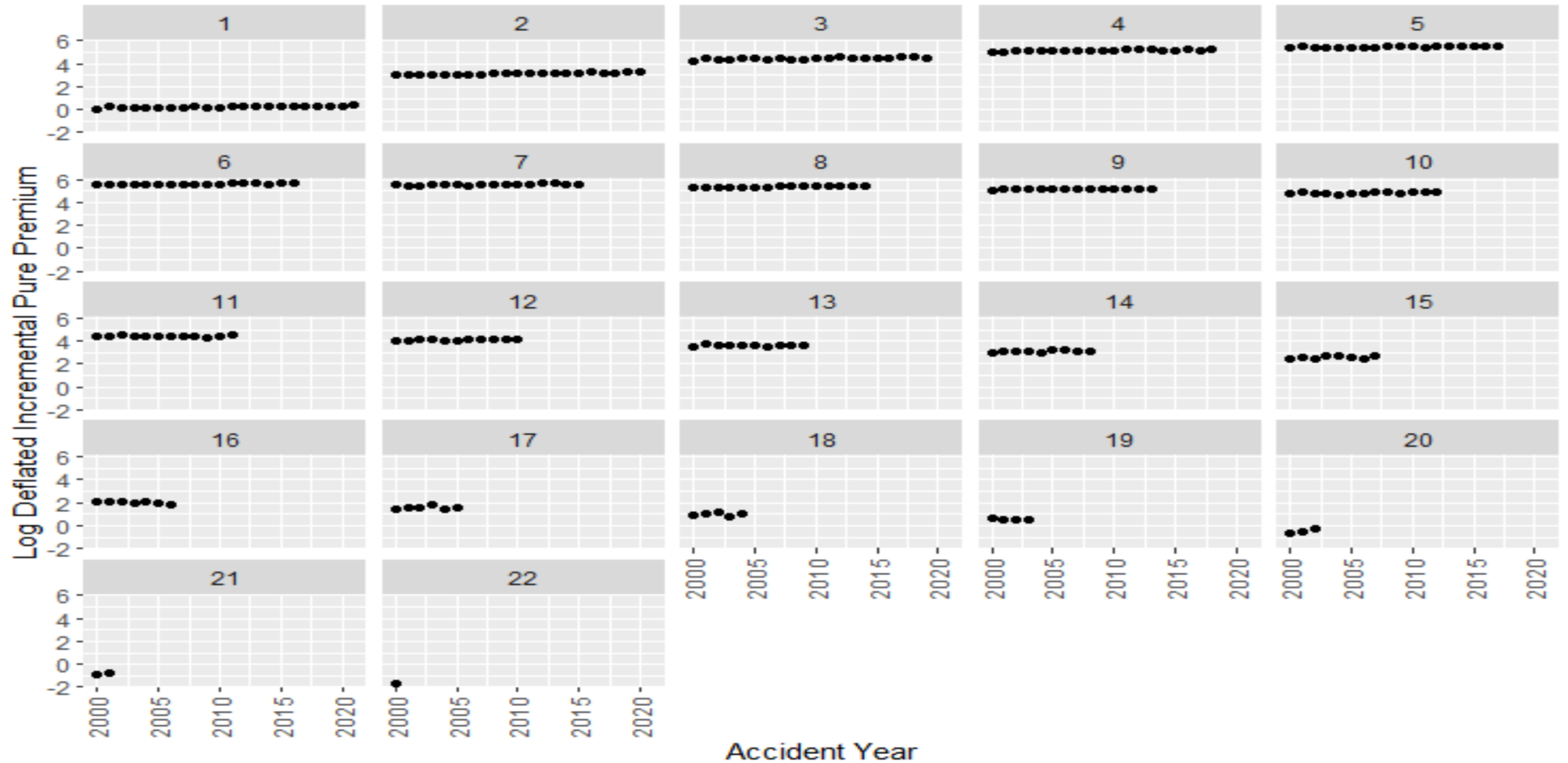
Box Plot Incremental Paid Count Freq
Across Accident Year



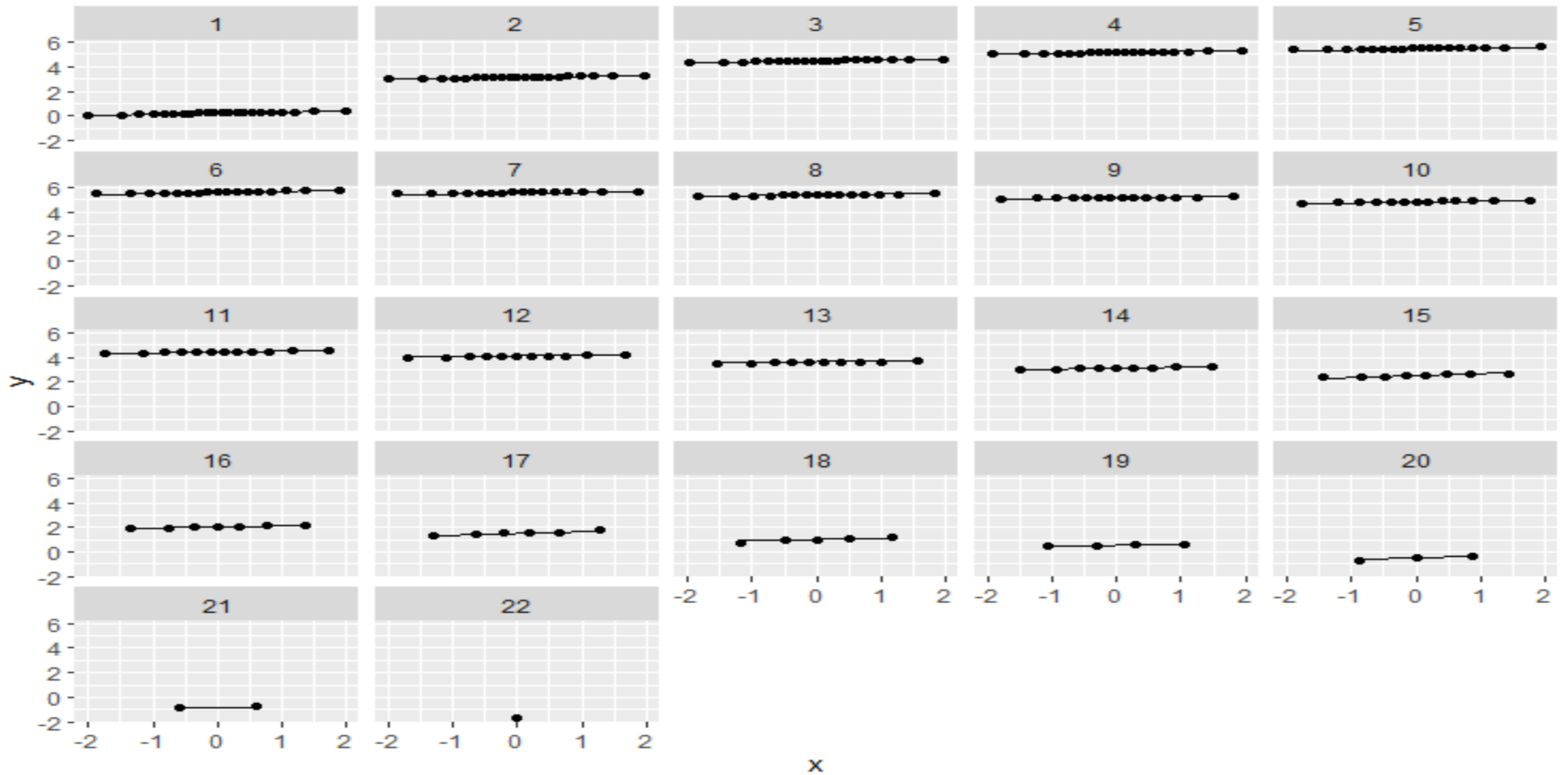
Reported Claim Count at One Year By Accident Year



Plot of Log of Deflated Incremental Pure Premium Across Accident Year Grouped by Development Year



QQ Plot for Log of Deflated Pure Premium Incremental Payments By Development Year



EDA Observations

- Lognormal distribution should work for dependent variable
- Mean shows some form of a parabolic development pattern
- Sigma generally decreases first 10 development years then starts to increase
- No sign of change in rate of loss cost increases
- No sign of change in claim handling or business mix

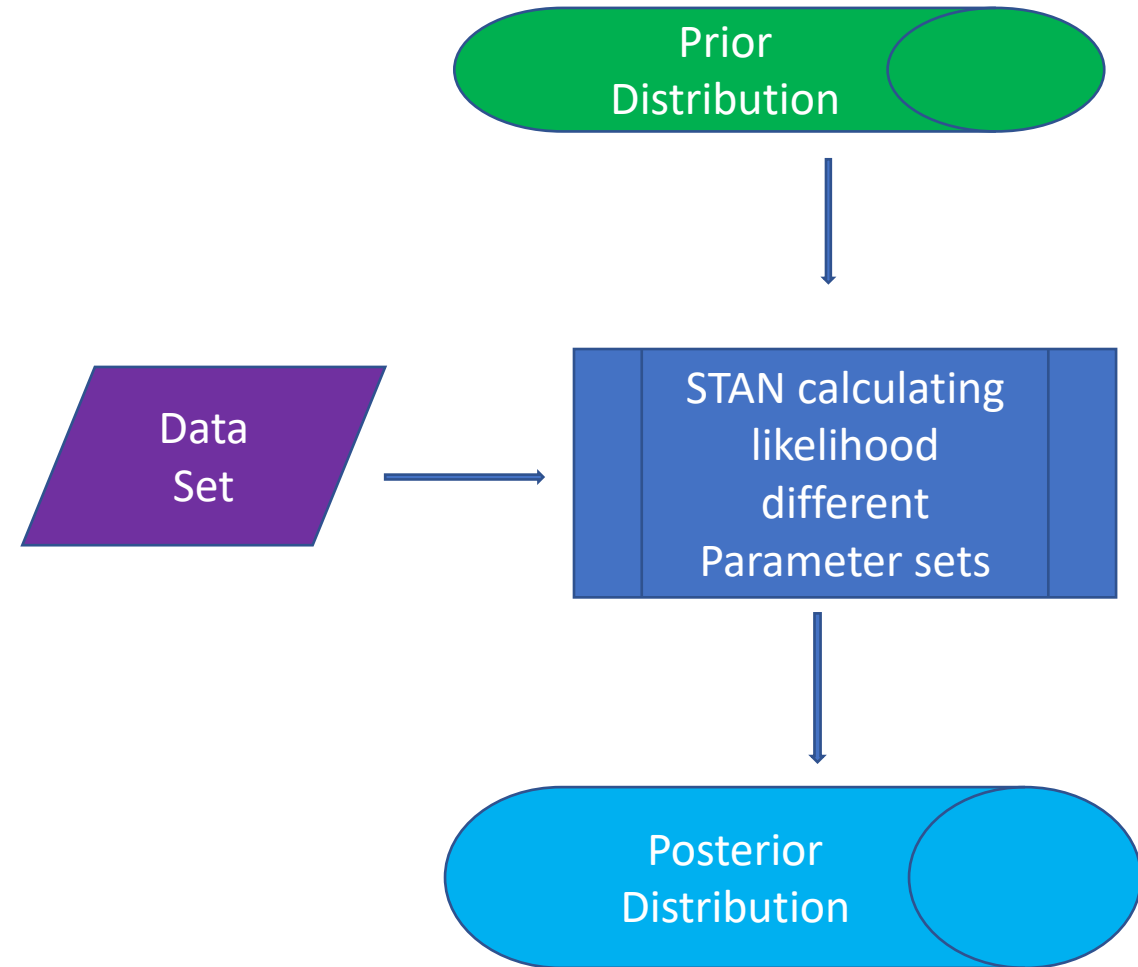
Alternative Models

Model forms

- Model incremental paid loss to isolate effect of inflation
- Normalize losses (incremental deflated pure premium)
 - Divide incremental paid losses by exposure (claim count at 12 months)
 - Divide incremental losses by accumulated CPI type index
- Incremental pure premium vs. incremental counts and average deflated severity
 - Modeling counts and amounts implies measuring correlation between counts & severity estimates which is a more complicated model
 - Built in log pointwise prediction comparison tools need a single, common predictive value
 - Appealing but impractical model form for this exercise
- Potential model forms:
 - Polynomial for development year & random effects (group) for accident year
 - Non-linear for development year & random effects (group) for accident year
 - Generalized Additive Model for development year & random effects (group) for accident year
 - Additional inflation effects as calendar year time elapsed

Effect of STAN on Modeling

- Execution time is much better
- Bayesian MCMC solves for parametric parameters iteratively for a wide range of model structures
- Start with a given set of parameters (prior distribution) and compare successive alternative sets effect on likelihood function substituting parameters that yield better likelihood results until convergence
- STAN uses an algorithm (Hamiltonian) that is much more efficient at selecting the next set of parameters to try than earlier Bayesian MCMC algorithms
- Diagnostics on algorithm execution success available



Polynomial 1 Prior Definitions

```
Polynomial_1_prior <- c(prior(normal(.6,.2),class=b, coef= Dev_Yr_6_Cap),  
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_6_Cap_Sqrd),  
  prior(normal(-.2,.1),class=b, coef= Dev_Yr_6_Spline),  
  prior(normal(-.1,.05),class=b, coef= Dev_Yr_6_Spline_Sqrd),  
  prior(normal(1,.25),class=b, coef=Ln_Dev_Yr),  
  prior(normal(-1,.25),class=b ,coef =Intercept),  
  prior(normal(.02,.01),class=b, coef=Cal_Yr_Time),  
  prior(normal(0.2,.1),class=b ,coef =Intercept, dpar=sigma),  
  prior(normal(-.1,.05),class=b, coef=Dev_Yr_10_Cap,dpar=sigma),  
  prior(student_t(3,.1,.05),class=b, coef=Dev_Yr_10_Spline,dpar=sigma))
```

Prior distributions reflect your knowledge of the subject in terms of plausible results.

The priors with a “dpar=” are used to give instructions for model components besides the mean or mu for the lognormal distribution

A listing of variable definitions is given in the appendix.

Grammar for brms models

- Brm: calls the brms routine
- Components bolted on as needed via “+” signs
- bf(): defines the formulas used to estimate the mean and other parameters
- Iter: tells the routine how many times to simulate the parameters to solve in MCMC routine
- Prior: identifies the prior distributions to be used in a model
- Seed: sets the simulation seed to ensure results can be replicated
- Control: instructions to STAN when default settings don't work
- Data: name of data set

Polynomial Model 1 brms Instructions

```
Model_Polynomial_1 <- brm(bf(Trended_Incr_PP_Def ~ 0 + Intercept + Ln_Dev_Yr +  
  Dev_Yr_6_Cap + Dev_Yr_6_Cap_Sqrd +  
  Dev_Yr_6_Spline + Dev_Yr_6_Spline_Sqrd + (1 || Acc_Yr) +  
  Cal_Yr_Time ,  
  sigma ~ 0 + Intercept + Dev_Yr_10_Cap + Dev_Yr_10_Spline ),  
  iter = 4000,  
  prior = Polynomial_1_prior,  
  seed = 8603529,  
  control = list(max_treedepth = 15 ),  
  data = Train_Triangle_All_Operation, family = lognormal())
```

Mu or the mean's model is defined after the first “~”.

Sigma's model is defined after the second “~”.

Other modeling instructions are bolted on as needed.

Polynomial Model 1 Results Summary

```
summary(Model_Polynomial_1)
```

```
Family: lognormal
```

```
Links: mu = identity; sigma = log
```

```
Formula: Trended_Incr_PP_Def ~ 0 + Intercept + Ln_Dev_Yr + Dev_Yr_6_Cap + Dev_Yr_6_Cap_Sqrd  
+ Dev_Yr_6_Spline + Dev_Yr_6_Spline_Sqrd + (1 || Acc_Yr) + Cal_Yr_Time
```

```
sigma ~ 0 + Intercept + Dev_Yr_10_Cap + Dev_Yr_10_Spline
```

```
Data: Train_Triangle_All_Operation (Number of observations: 253)
```

```
Draws: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;  
total post-warmup draws = 8000
```

```
Group-Level Effects:
```

```
~Acc_Yr (Number of levels: 22)
```

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sd(Intercept)	0.03	0.01	0.01	0.05	1.00	2529	3532

```
Population-Level Effects:
```

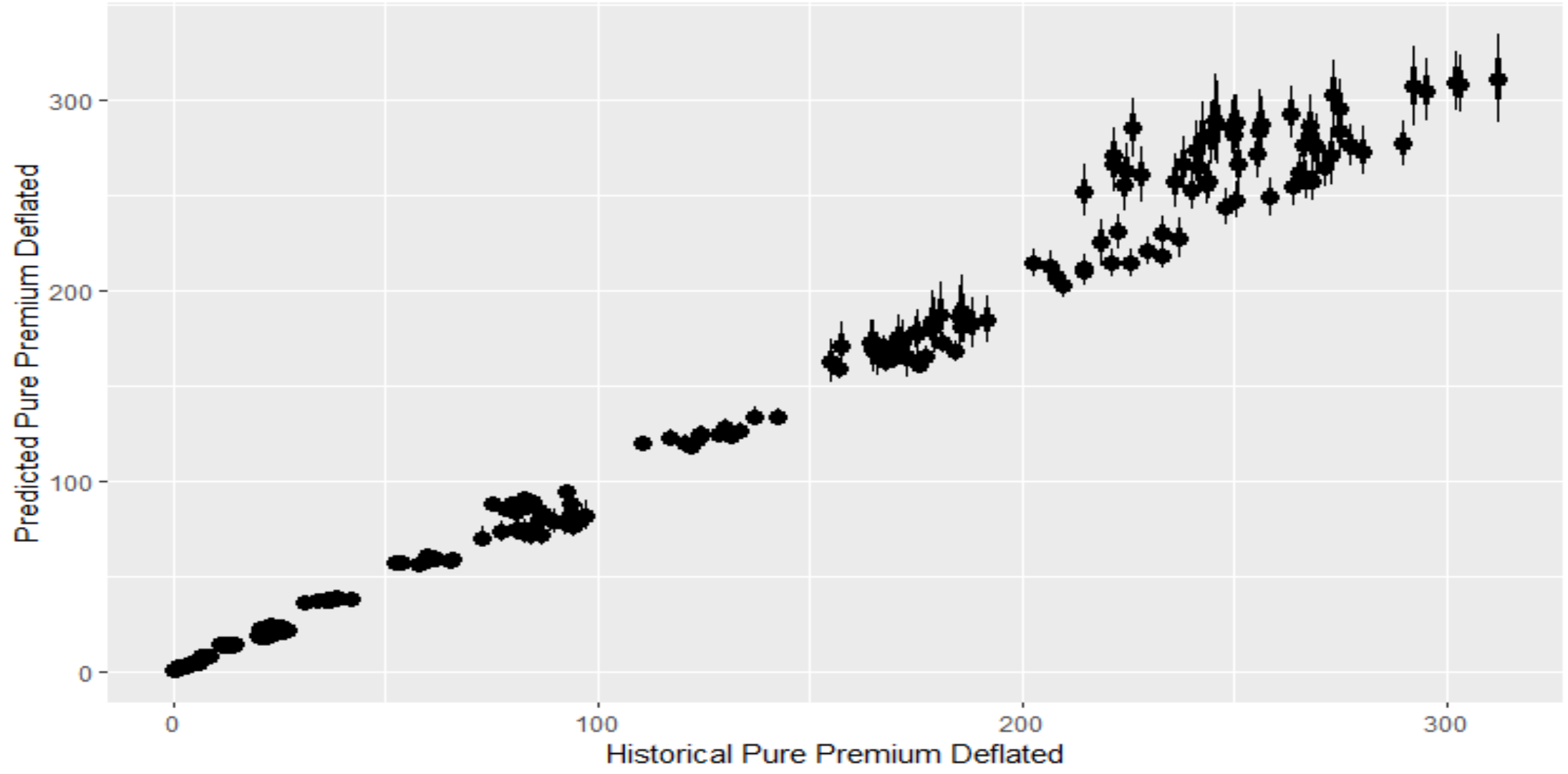
	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
Intercept	-0.51	0.13	-0.77	-0.25	1.00	3371	4709
Ln_Dev_Yr	2.24	0.18	1.89	2.60	1.00	2466	4264
Dev_Yr_6_Cap	1.16	0.11	0.93	1.38	1.00	2153	3348
Dev_Yr_6_Cap_Sqrd	-0.14	0.01	-0.15	-0.12	1.00	2600	4169
Dev_Yr_6_Spline	-0.45	0.03	-0.50	-0.40	1.00	2296	3765
Dev_Yr_6_Spline_Sqrd	-0.01	0.00	-0.02	-0.01	1.00	3051	5357
Cal_Yr_Time	0.01	0.00	0.00	0.01	1.00	4143	5340
sigma_Intercept	-0.37	0.09	-0.53	-0.20	1.00	4437	5434
sigma_Dev_Yr_10_Cap	-0.30	0.01	-0.32	-0.27	1.00	3832	4614
sigma_Dev_Yr_10_Spline	0.22	0.02	0.17	0.27	1.00	5625	6240

Draws were sampled using `sampling(NUTS)`. For each parameter, `Bulk_ESS` and `Tail_ESS` are effective sample size measures, and `Rhat` is the potential scale reduction factor on split chains (at convergence, `Rhat = 1`).

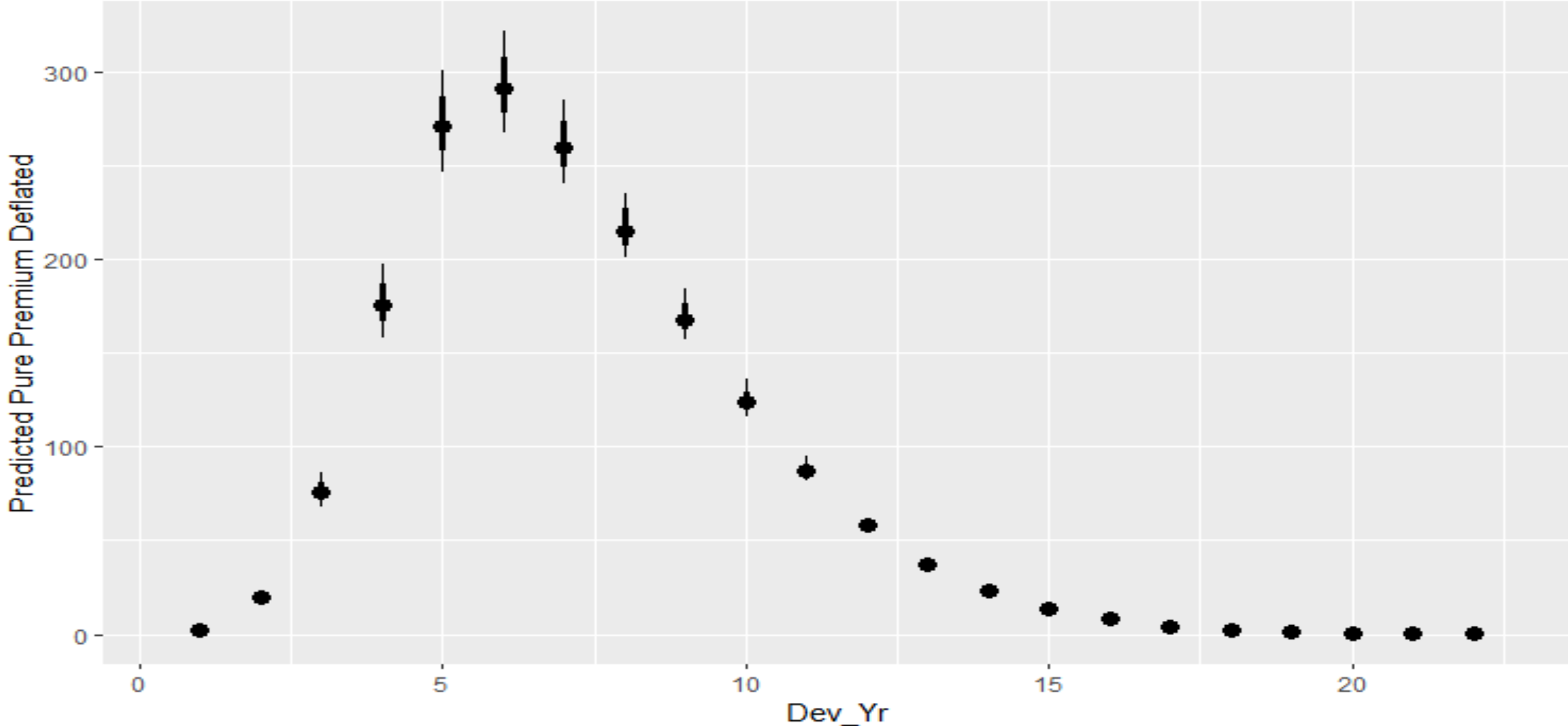
Bayesian MCMC performance summary terms

- Rhat: used to summarize parameter comparison between chains with 1.00 signifying the comparison is good.
- Bulk_ESS: Number of non-correlated iterations used to get measure of effective overall sample size for an estimate
- Tail_ESS: Number of non-correlated iterations used to get measure of effective sample size for an estimate in the tail of the distribution
- Group Level effects: displays standard deviation across groups used in least squares credibility weighting of mean result for a given group
- Population level effects: equivalent to GLM independent variables

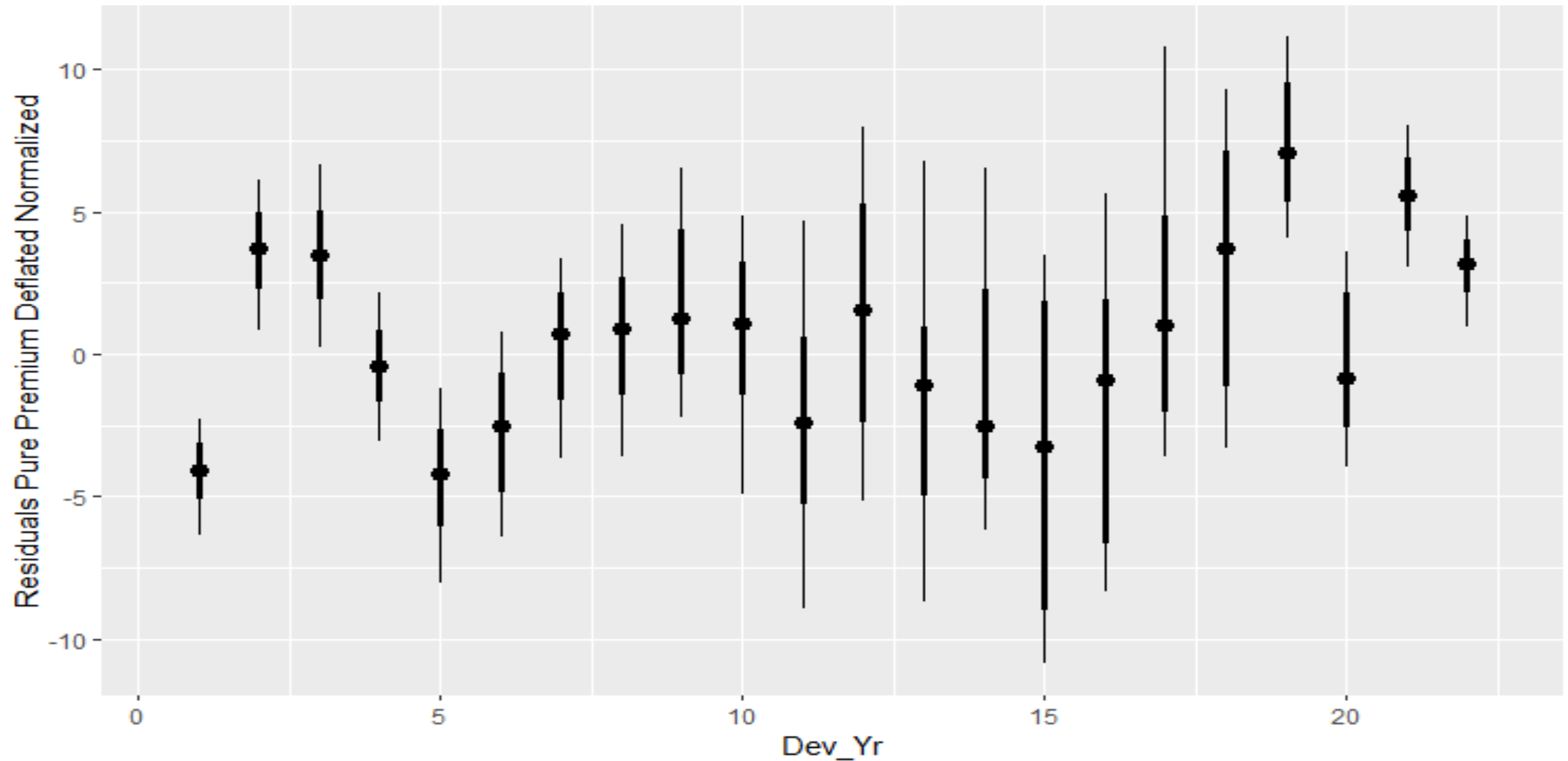
Historical Pure Premium vs. Distribution of Predicted Model Polynomial 1



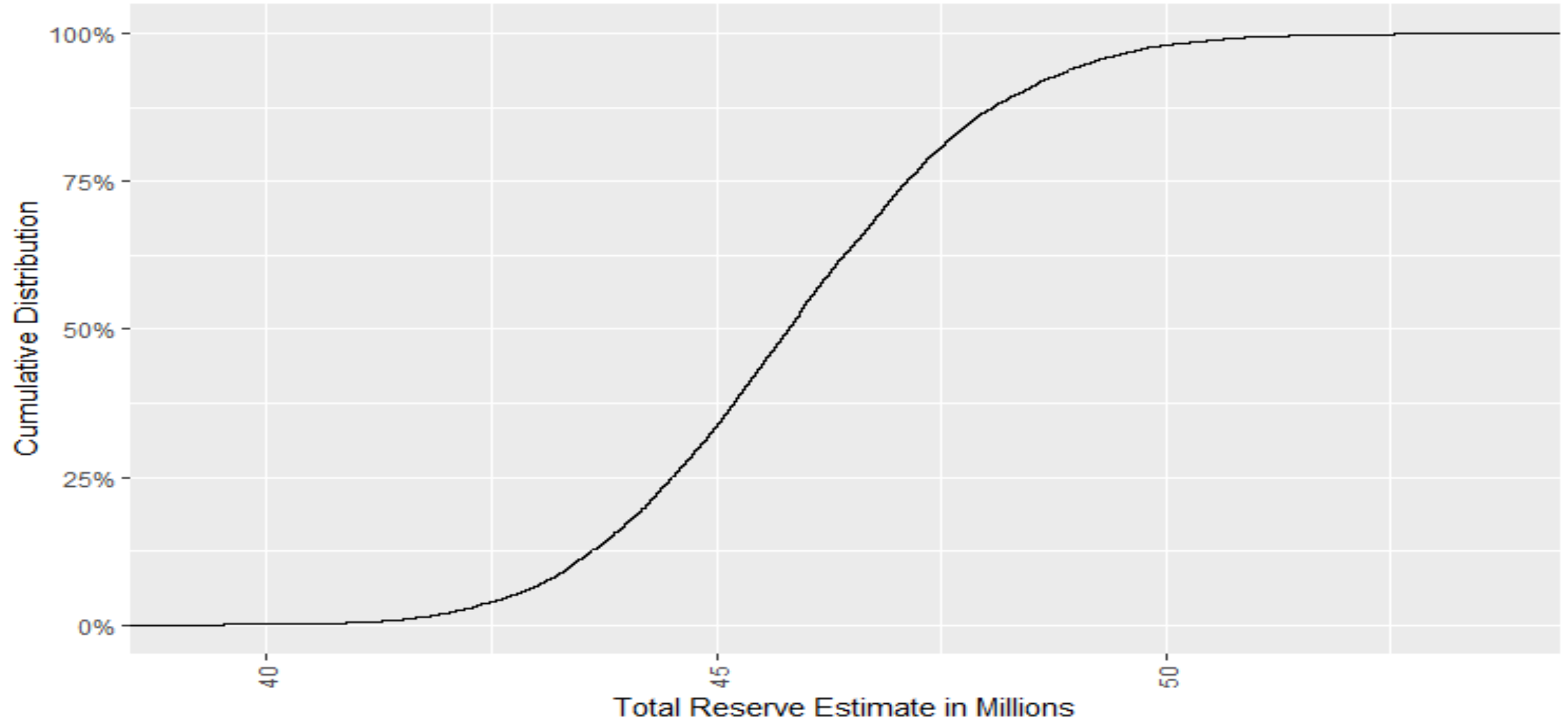
Historical Pure Premium vs. Distribution of Predicted Model Polynomial 1



Development Year vs. Normalized Residuals Polynomial Model 1



**Total Reserve Estimate
Model Polynomial 1
Simulated Future Inflation**



Polynomial 2 Prior Definitions

```
Polynomial_2_prior <- c(prior(normal(.6,.2),class=b, coef= Dev_Yr_6_Cap),  
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_6_Cap_Sqrd),  
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_6_Spline),  
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_6_Spline_Sqrd),  
  prior(normal(1,.25),class=b, coef=Dev_Yr_2_Factor2),  
  prior(normal(2,.25),class=b, coef=Dev_Yr_2_Factor3),  
  prior(normal(-1,.25),class=b ,coef =Intercept),  
  prior(normal(.02,.01),class=b, coef=Cal_Yr_Time),  
  prior(normal(-.01,.005),class=b, coef=Dev_Yr_10_Cap,dpar=sigma),  
  prior(student_t(2,.02,.01),class=b, coef=Dev_Yr_10_Spline,dpar=sigma))
```


Polynomial Model 2 brms Instructions

```
Model_Polynomial_2 <- brm(bf(Trended_Incr_PP_Def ~ 0 + Intercept + Dev_Yr_2_Factor +  
    Dev_Yr_6_Cap + Dev_Yr_6_Cap_Sqrd + Dev_Yr_6_Spline  
    + Dev_Yr_6_Spline_Sqrd + Cal_Yr_Time + (1||Acc_Yr),  
    sigma ~ Dev_Yr_10_Cap + Dev_Yr_10_Spline ),  
    iter = 4000,  
    prior= Polynomial_2_prior,  
    seed= 8603529,  
    data = Train_Triangle_All_Operation, family = lognormal())
```

Polynomial Model 2 Results Summary

summary(Model_Polynomial_2)

```
Family: lognormal
Links: mu = identity; sigma = log
Formula: Trended_Incr_PP_Def ~ 0 + Intercept + Dev_Yr_2_Factor + Dev_Yr_6_Cap
        + Dev_Yr_6_Cap_Sqrd + Dev_Yr_6_Spline + Dev_Yr_6_Spline_Sqrd + Cal_Yr_Time + (1 || Acc_Yr)
        sigma ~ Dev_Yr_10_Cap + Dev_Yr_10_Spline
Data: Train_Triangle_All_Operation (Number of observations: 253)
Draws: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
       total post-warmup draws = 8000
```

Group-Level Effects:

~Acc_Yr (Number of levels: 22)

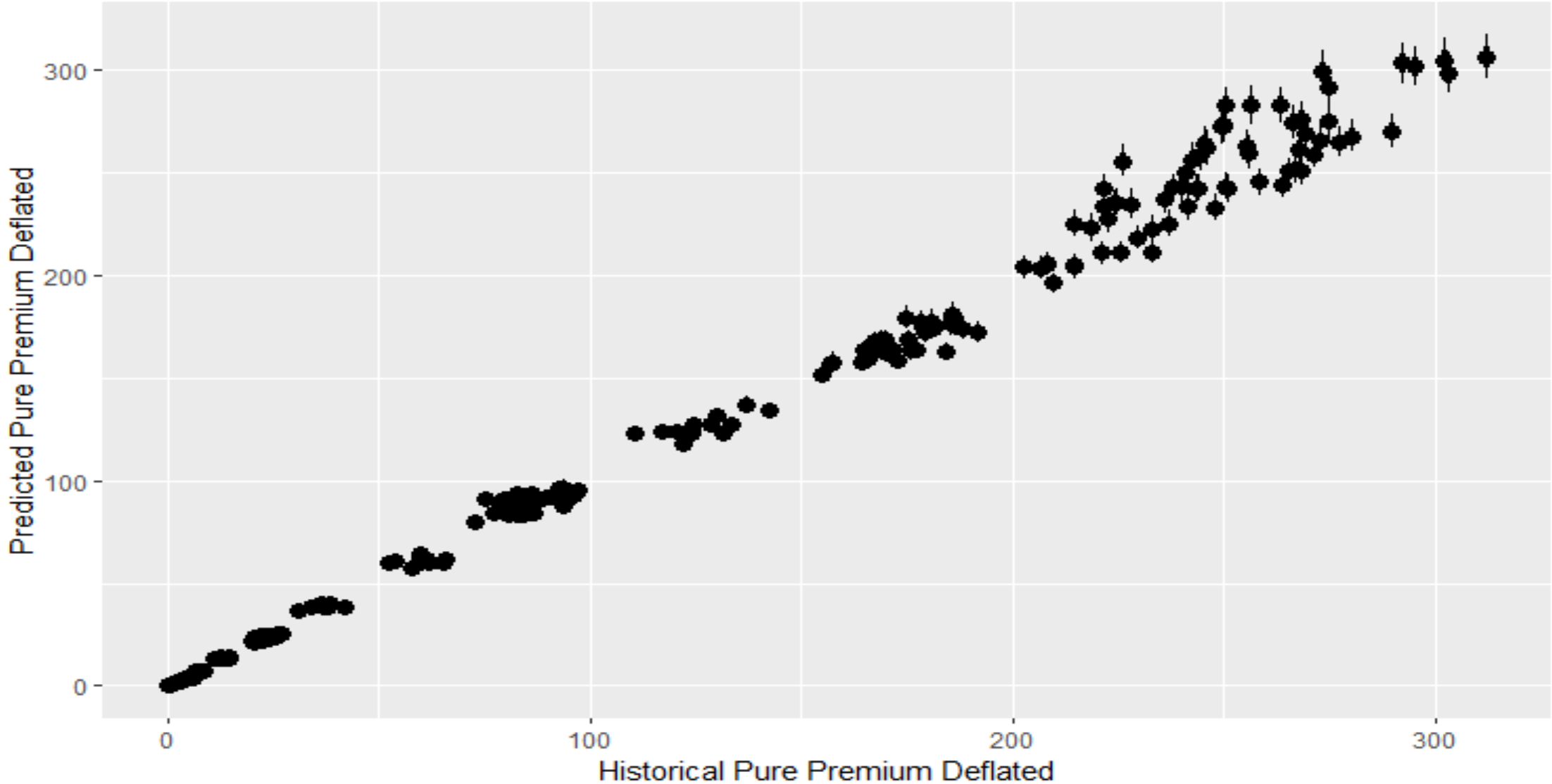
	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sd(Intercept)	0.02	0.01	0.01	0.03	1.00	2671	2719

Population-Level Effects:

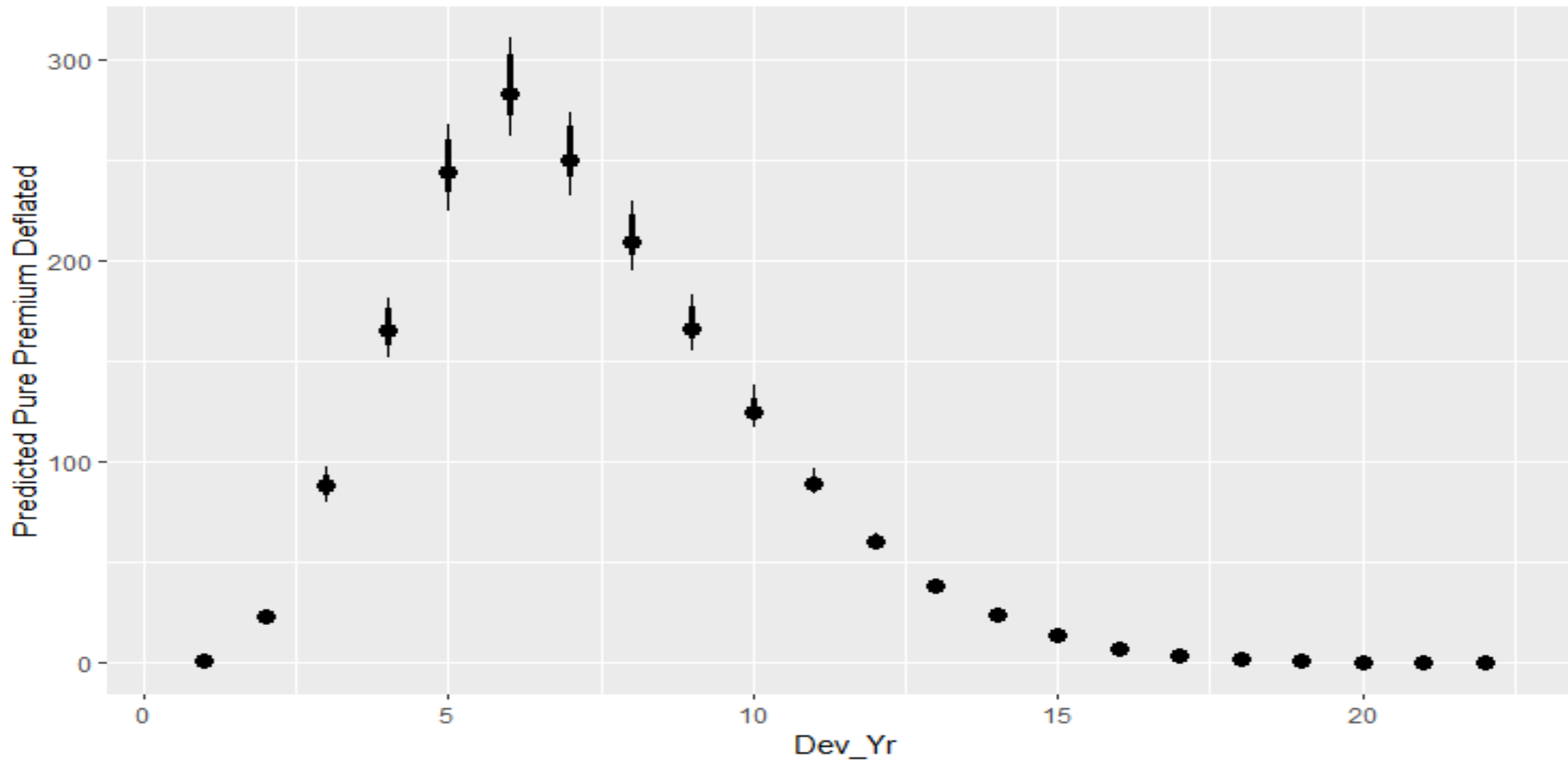
	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sigma_Intercept	-2.88	0.07	-3.00	-2.74	1.00	5753	5598
Intercept	-1.22	0.05	-1.31	-1.12	1.00	2148	3641
Dev_Yr_2_Factor2	1.79	0.04	1.72	1.86	1.00	2288	3638
Dev_Yr_2_Factor3	2.26	0.06	2.14	2.39	1.00	2093	3217
Dev_Yr_6_Cap	1.47	0.05	1.37	1.57	1.00	1973	2856
Dev_Yr_6_Cap_Sqrd	-0.12	0.01	-0.13	-0.11	1.00	1974	2861
Dev_Yr_6_Spline	-0.10	0.01	-0.12	-0.09	1.00	4320	5297
Dev_Yr_6_Spline_Sqrd	-0.03	0.00	-0.03	-0.02	1.00	4383	5083
Cal_Yr_Time	0.01	0.00	0.01	0.01	1.00	5867	5854
sigma_Dev_Yr_10_Cap	-0.02	0.00	-0.03	-0.01	1.00	10244	6103
sigma_Dev_Yr_10_Spline	0.25	0.02	0.20	0.29	1.00	4987	5873

Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

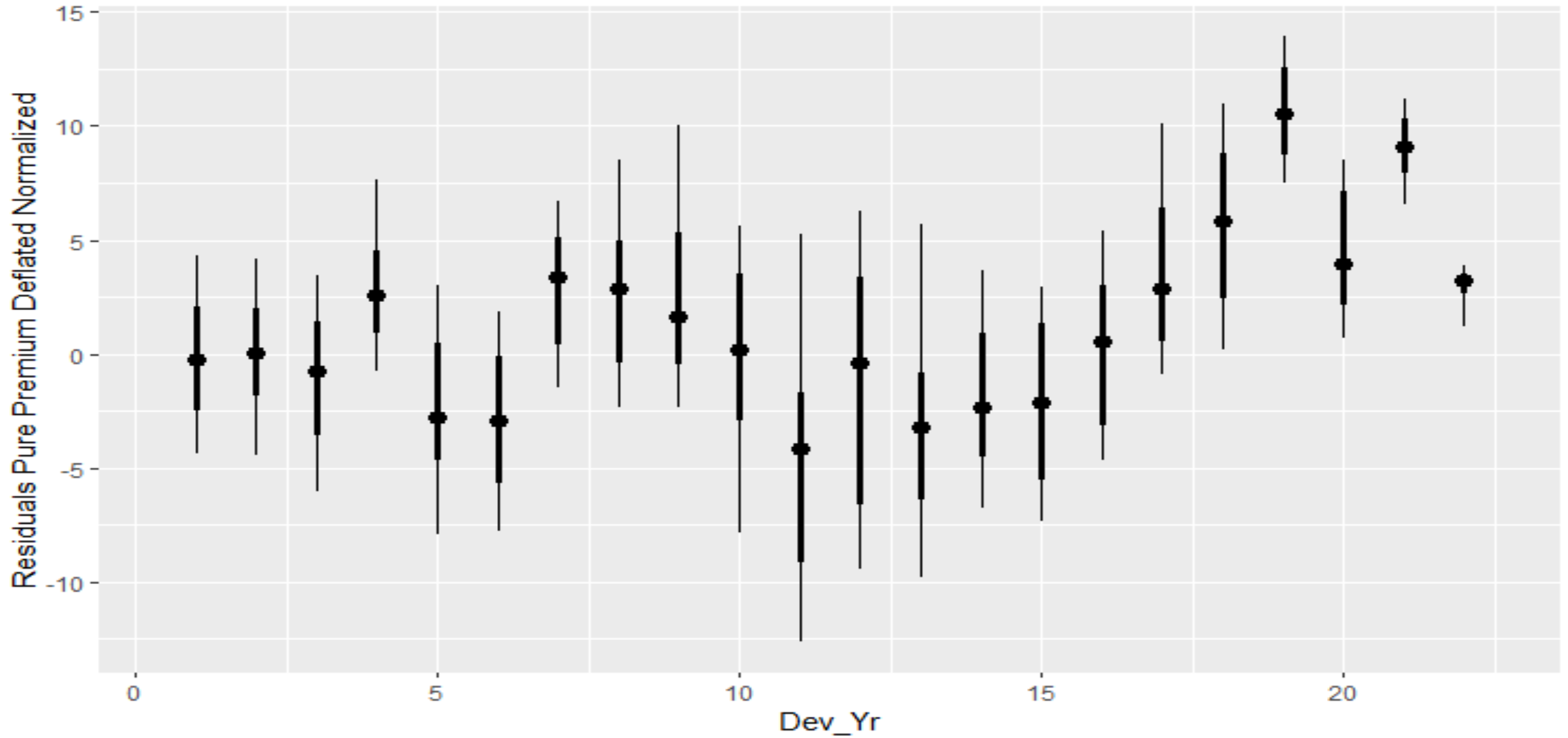
Historical Pure Premium Deflated vs. Distribution of Predicted Polynomial Model 2



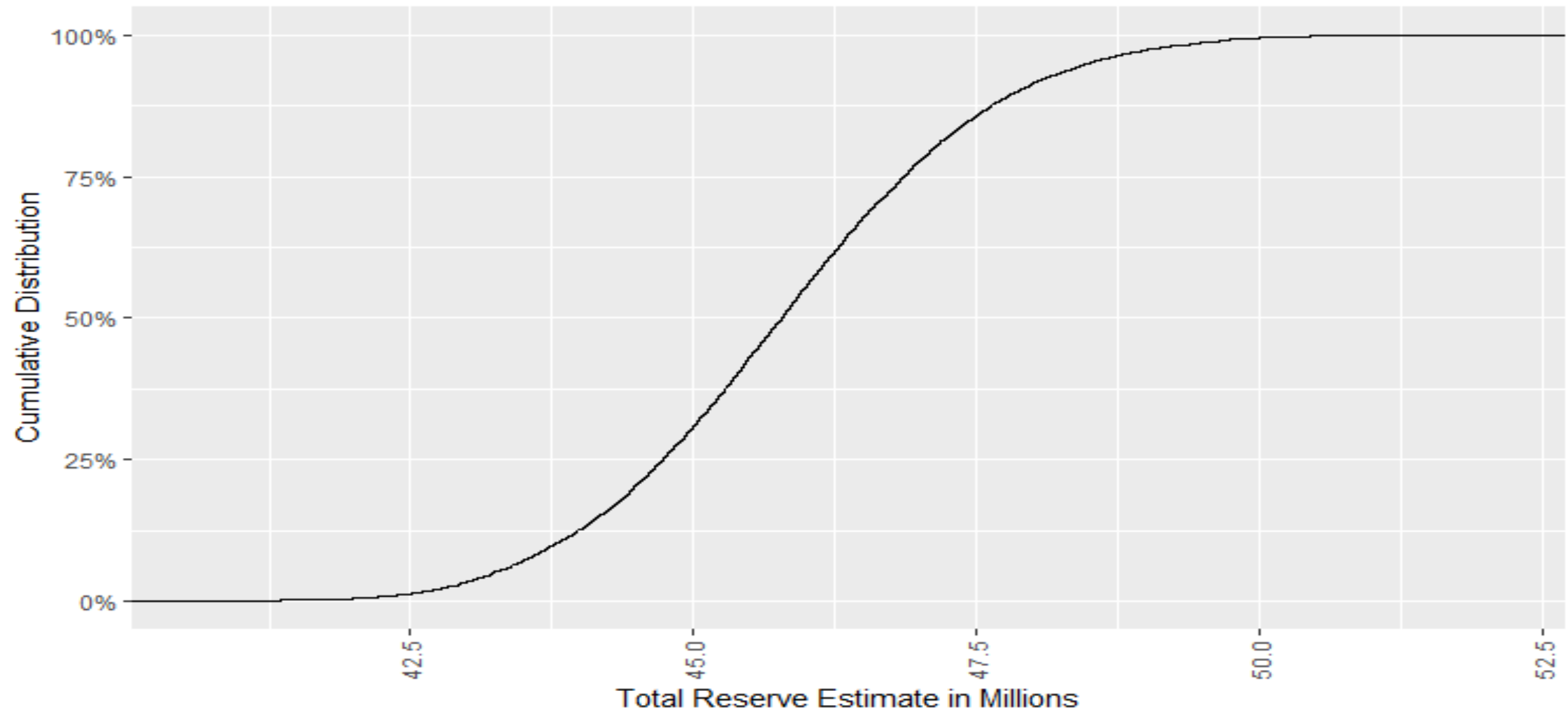
Historical Pure Premium vs. Distribution of Predicted Polynomial Model 2



Development Year vs. Normalized Residuals Polynomial Model 2



**Total Reserve Estimate
Model Polynomial 2
Simulated Future Inflation**



Comments on Multivariate Model Slides

- Excluded from model comparison
- Slides included to show some features of building a reserve model if you have to split it into counts and severity to get a decent fit
- Prior definitions
 - Split Poisson vs. Lognormal using “resp=”
 - Other components work as before
- Split the model formula using separate objects for counts and severity
 - Keeps the length of code for a given task readable
 - Bolt the model objects together using “+”

Multivariate Model Prior Definitions

```
multivariate_prior <- c(prior(normal(.6,.2),class=b, coef= Dev_Yr_8_Cap, resp=TrendedMeanPaymentDef),
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_8_Cap_Sqrd,resp=TrendedMeanPaymentDef),
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_8_Spline,resp=TrendedMeanPaymentDef),
  prior(normal(-.2,.2),class=b, coef= Dev_Yr_8_Spline_Sqrd,resp=TrendedMeanPaymentDef),
  prior(normal(1,.25),class=b, coef=Dev_Yr_1_Factor2,resp=TrendedMeanPaymentDef),
  prior(normal(1,.25),class=b ,coef =Intercept,resp=TrendedMeanPaymentDef),
  prior(normal(.02,.01),class=b, coef=Cal_Yr_Time,resp=TrendedMeanPaymentDef),
  prior(normal(-2,.5),class=b ,coef =Intercept,dpar=sigma,resp=TrendedMeanPaymentDef),
  prior(normal(-.01,.005),class=b, coef=Dev_Yr_15_Cap,dpar=sigma,resp=TrendedMeanPaymentDef),
  prior(normal(-1,.5),class=b, coef=Intercept,resp=PaidCnt),
  prior(normal(.2,.1),class=b, coef=Dev_Yr_4_Cap,resp=PaidCnt),
  prior(normal(1,.1),class=b, coef=Ln_Dev_Yr,resp=PaidCnt),
  prior(normal(-.2,.1),class=b, coef=Dev_Yr_4_Spline,resp=PaidCnt))
```


Multivariate Model brms Instructions

```
form_count <- bf(Paid_Cnt | rate(Rptd_Cnt) ~ 0 + Intercept
  + Dev_Yr_4_Cap
  + Ln_Dev_Yr
  + Dev_Yr_4_Spline

  + (1||Acc_Yr), family = poisson())

form_sev <-bf(Trended_Mean_Payment_Def~ 0 + Intercept
  + Dev_Yr_8_Cap
  + Dev_Yr_8_Cap_Sqrd
  + Dev_Yr_8_Spline
  + Dev_Yr_8_Spline_Sqrd
  + Dev_Yr_1_Factor
  + Cal_Yr_Time
  + (1||Acc_Yr),
  sigma ~ 0 + Intercept +Dev_Yr_15_Cap ,
  family =lognormal() )

Model_Multivariate <- brm(form_sev + form_count ,
  prior=multivariate_prior,
  data = Train_Triangle_All_Operati
```

Multivariate Model Results Summary

```

Family: MV(lognormal, poisson)
Links: mu = identity; sigma = log
       mu = log
Formula: Trended_Mean_Payment_Def ~ 0 + Intercept + Dev_Yr_8_Cap + Dev_Yr_8_Cap_Sqrd + Dev_Yr_8_Spline
        + Dev_Yr_8_Spline_Sqrd + Dev_Yr_1_Factor + Cal_Yr_Time + (1 || Acc_Yr)
        sigma ~ 0 + Intercept + Dev_Yr_15_Cap
        Paid_Cnt | rate(Rptd_Cnt) ~ 0 + Intercept + Dev_Yr_4_Cap + Ln_Dev_Yr + Dev_Yr_4_Spline + (1 || Acc_Y
Data: Train_Triangle_All_Operation (Number of observations: 253)
Draws: 4 chains, each with iter = 2000; warmup = 1000; thin = 1;
       total post-warmup draws = 4000
    
```

Group-Level Effects:

~Acc_Yr (Number of levels: 22)

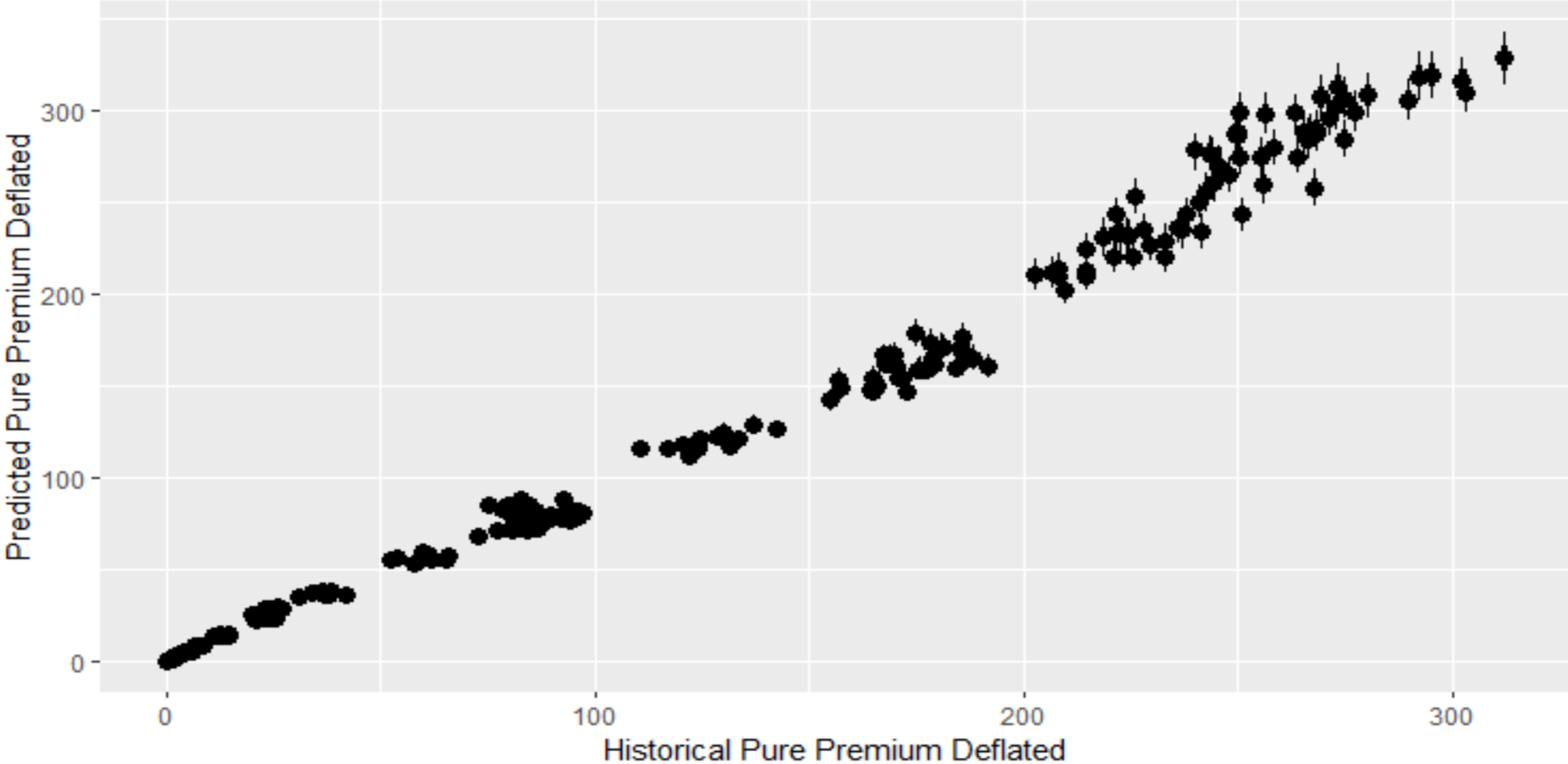
	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sd(TrendedMeanPaymentDef_Intercept)	0.01	0.01	0.00	0.03	1.00	1533	1789
sd(PaidCnt_Intercept)	0.02	0.01	0.01	0.03	1.00	1556	2187

Population-Level Effects:

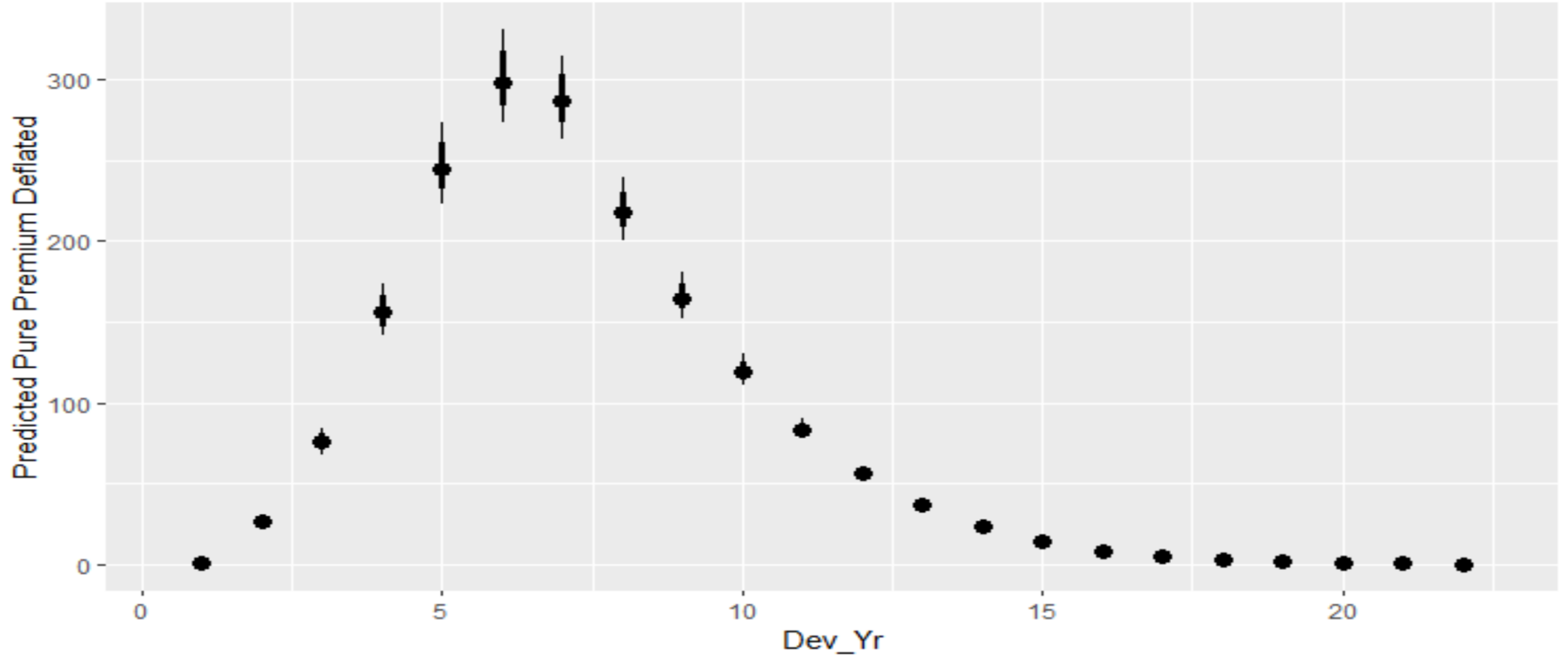
	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
TrendedMeanPaymentDef_Intercept	0.04	0.03	-0.01	0.10	1.00	4628	3231
TrendedMeanPaymentDef_Dev_Yr_8_Cap	1.48	0.02	1.44	1.52	1.00	3322	3061
TrendedMeanPaymentDef_Dev_Yr_8_Cap_Sqrd	-0.11	0.00	-0.11	-0.10	1.00	3347	3171
TrendedMeanPaymentDef_Dev_Yr_8_Spline	-0.09	0.01	-0.11	-0.08	1.00	6696	3244
TrendedMeanPaymentDef_Dev_Yr_8_Spline_Sqrd	-0.01	0.00	-0.01	-0.01	1.00	6412	3328
TrendedMeanPaymentDef_Dev_Yr_1_Factor2	1.48	0.04	1.41	1.55	1.00	4355	3272
TrendedMeanPaymentDef_Cal_Yr_Time	0.01	0.00	0.01	0.01	1.00	5502	3292
sigma_TrendedMeanPaymentDef_Intercept	-2.37	0.05	-2.47	-2.26	1.00	5246	2927
sigma_TrendedMeanPaymentDef_Dev_Yr_15_Cap	-0.02	0.00	-0.03	-0.01	1.00	5743	3437
PaidCnt_Intercept	-0.99	0.02	-1.02	-0.96	1.00	2647	2966
PaidCnt_Dev_Yr_4_Cap	-0.29	0.02	-0.32	-0.26	1.00	2438	2788
PaidCnt_Ln_Dev_Yr	0.97	0.03	0.91	1.03	1.00	2509	2677
PaidCnt_Dev_Yr_4_Spline	-0.29	0.00	-0.30	-0.29	1.00	2649	2834

Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

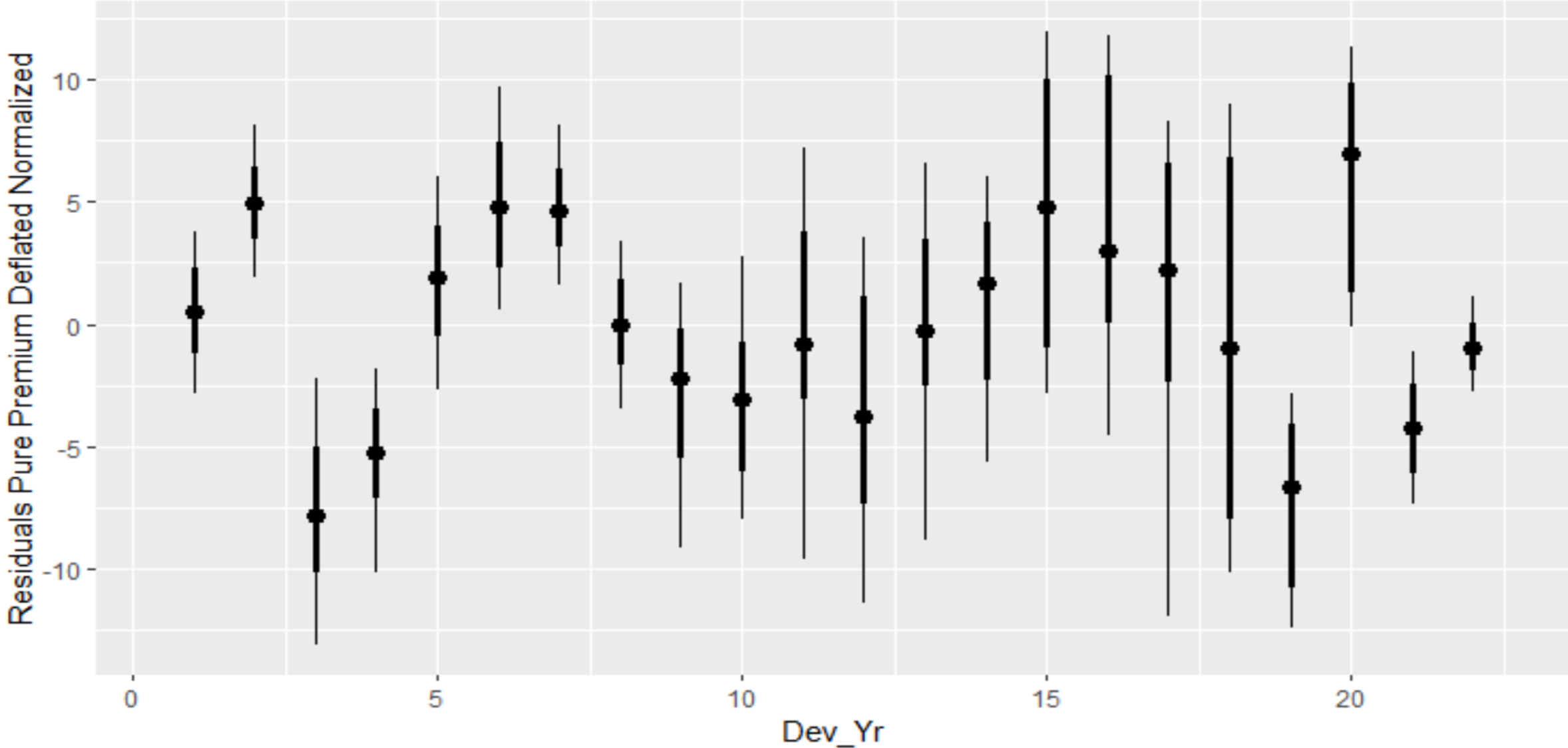
Historical Pure Premium vs. Distribution of Predicted Multivariate Model



Dev_Yr vs. Distribution of Predicted
Multivariate MModel



Development Year vs. Normalized Residuals Multivariate Model



Nonlinear Model Prior Definitions

```
Non_linear_prior_4 <- c(prior(normal(-.1,.05),class=b, nlpar=b1,ub=0),  
  prior(normal(1.5,.25),class=b, nlpar=b2,ub=0),  
  prior(normal(-.1,.05),class=b, nlpar=b3,ub=0),  
  prior(normal(-.1,.05),class=b, nlpar=b4,ub=0),  
  prior(normal(1,.5),class=b, nlpar=b5,ub=2),  
  prior(normal(.02,.02),class=b, nlpar=infl,lb=0),  
  prior(normal(-1,.1),class=b, nlpar=dev1,ub =-.4),  
  prior(normal(-2,.1),class=b ,coef =Intercept, dpar=sigma),  
  prior(normal(-.01,.005),class=b, coef=Dev_Yr_10_Cap,dpar=sigma),  
  prior(student_t(3,.05,.01),class=b, coef=Dev_Yr_10_Spline_Ln,dpar=sigma))
```

Nonlinear Model brms Instructions

```
Model_NL_4 <- brm(
  bf(Trended_Incr_PP_Def ~
    ( exp(dev1 + b1 *Dev_Yr_5_Cap
      + b2*Ln_Dev_Yr
      + b3 * Dev_Yr_5_Spline
      + b4 * Dev_Yr_5_Spline_Sqrd
      + b5 * Dev_Yr_GT_1
    )
    +infl*Cal_Yr_Time),

  b1 +b2 +b3 + +b4+ b5 +infl ~ 1,dev1 ~ 1 +(1||Acc_Yr),
  sigma ~ 0 + Intercept +Dev_Yr_10_Spline_Ln +Dev_Yr_10_Cap,
  nl = TRUE),
  data = Train_Triangle_All_Operation, family = lognormal(),
  prior= Non_linear_prior_4,
  seed= 8603529,
  iter= 4000,
  control = list(adapt_delta = 0.99,
                 max_treedepth =12))
```

Nonlinear Model Results Summary

```
Family: lognormal
Links: mu = identity; sigma = log
Formula: Trended_Incr_PP_Def ~ (exp(dev1 + b1 * Dev_Yr_5_Cap + b2 * Ln_Dev_Yr + b3 * Dev_Yr_5_Spline
+ b4 * Dev_Yr_5_Spline_Sqrd + b5 * Dev_Yr_GT_1) + infl * Cal_Yr_Time)
b1 ~ 1
b2 ~ 1
b3 ~ 1
b4 ~ 1
b5 ~ 1
infl ~ 1
dev1 ~ 1 + (1 || Acc_Yr)
sigma ~ 0 + Intercept + Dev_Yr_10_Spline_Ln + Dev_Yr_10_Cap
Data: Train_Triangle_All_Operation (Number of observations: 253)
Draws: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
total post-warmup draws = 8000
```

Group-Level Effects:

~Acc_Yr (Number of levels: 22)

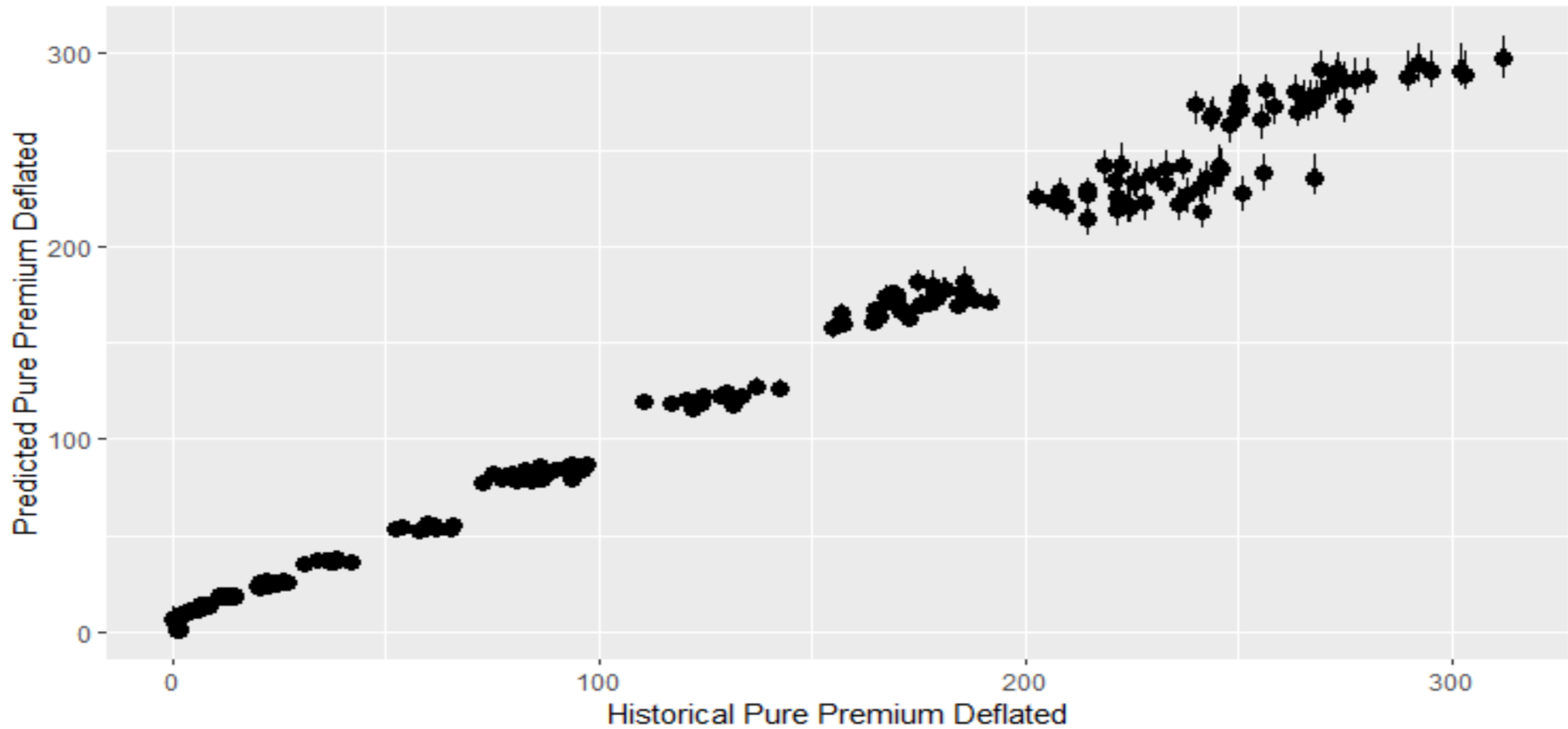
	Estimate	Est.Error	1-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sd(dev1_Intercept)	0.00	0.00	0.00	0.01	1.00	2556	3119

Population-Level Effects:

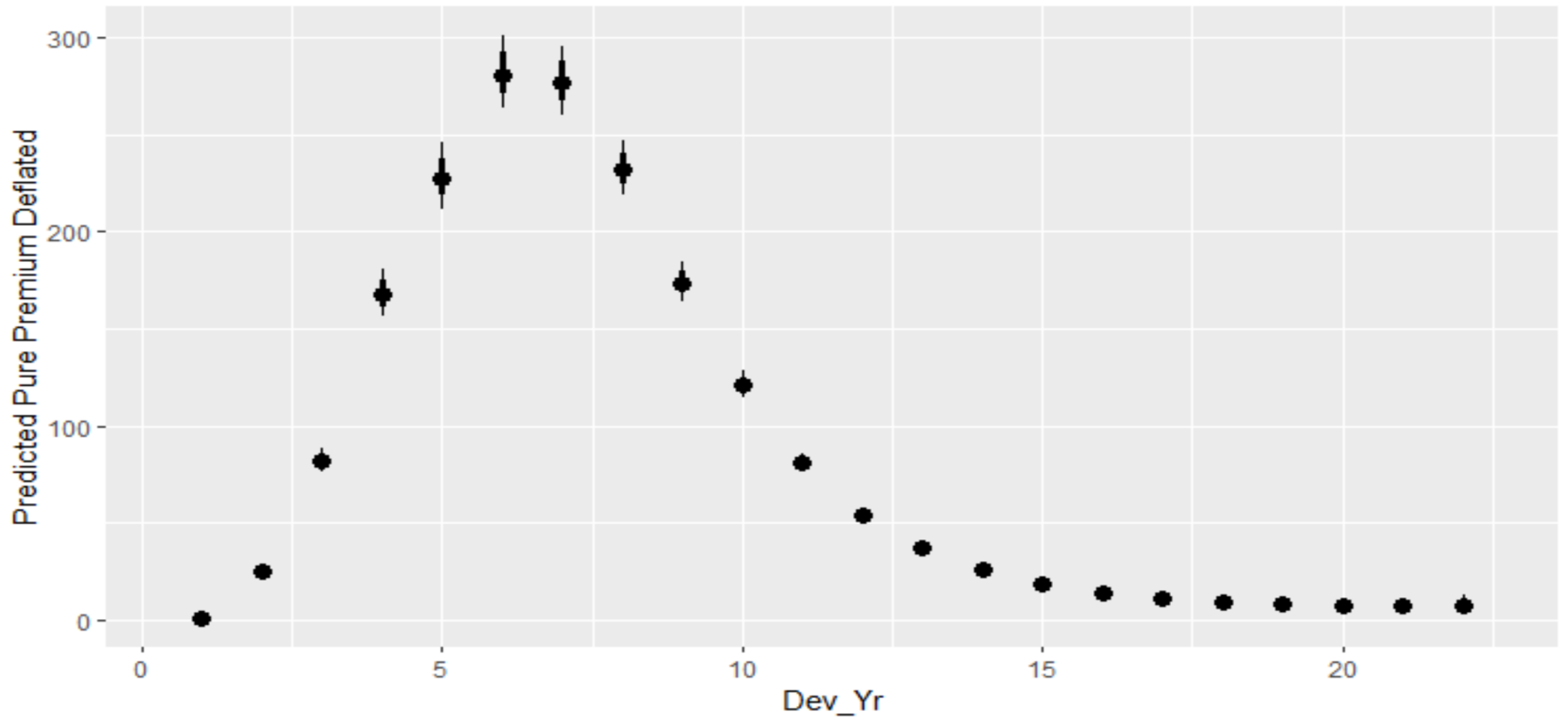
	Estimate	Est.Error	1-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
b1_Intercept	-0.27	0.01	-0.29	-0.24	1.00	3385	4533
b2_Intercept	1.45	0.03	1.38	1.52	1.00	3280	4305
b3_Intercept	-0.23	0.00	-0.24	-0.22	1.00	3224	4263
b4_Intercept	-0.00	0.00	-0.00	-0.00	1.00	6008	3723
b5_Intercept	1.93	0.05	1.83	2.00	1.00	4674	3548
infl_Intercept	0.01	0.00	0.00	0.01	1.00	4563	3458
dev1_Intercept	-1.26	0.05	-1.33	-1.15	1.00	4566	3762
sigma_Intercept	-2.50	0.06	-2.60	-2.38	1.00	7136	5925
sigma_Dev_Yr_10_Spline_Ln	1.19	0.06	1.08	1.31	1.00	7106	5937
sigma_Dev_Yr_10_Cap	-0.02	0.00	-0.03	-0.01	1.00	9890	6757

Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

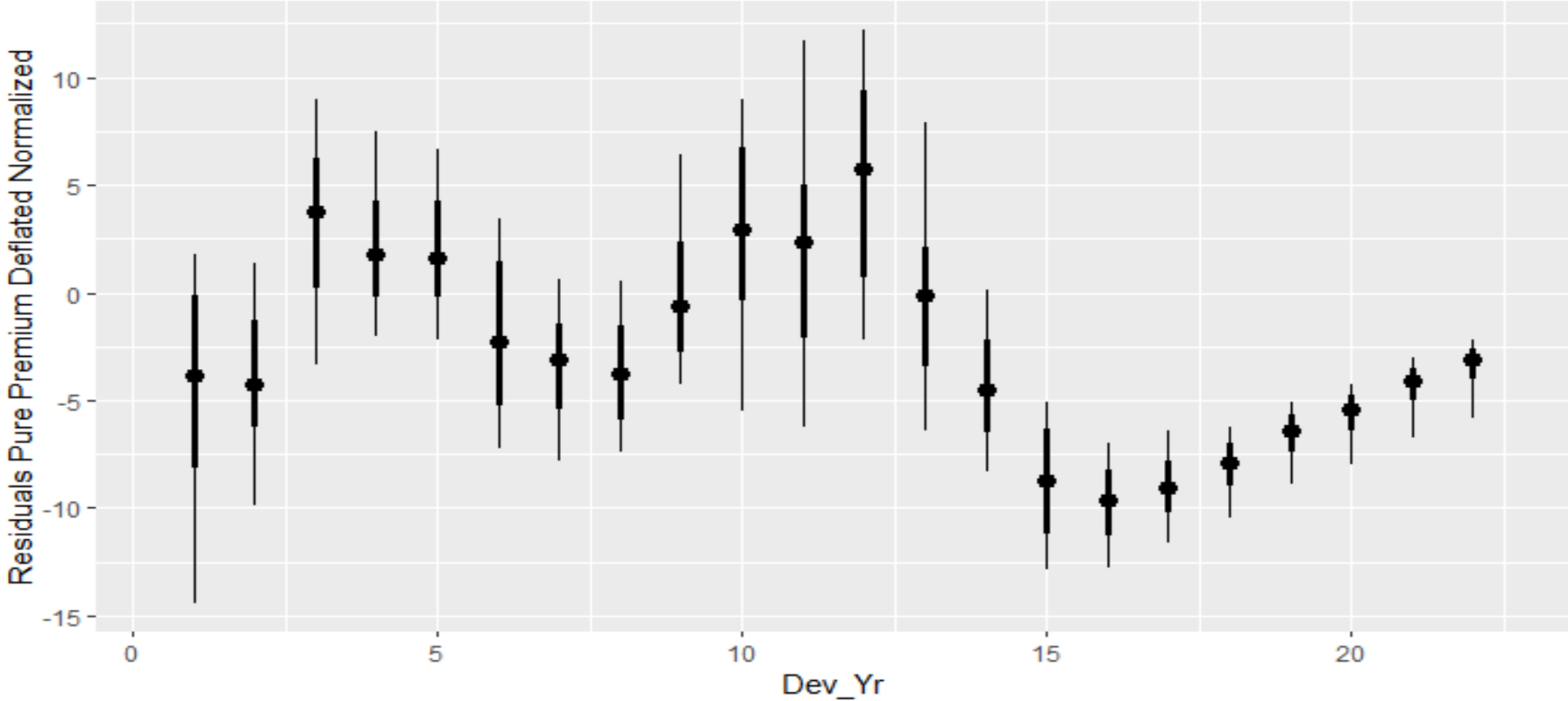
Historical Pure Premium vs. Distribution of Predicted Non-Linear Model



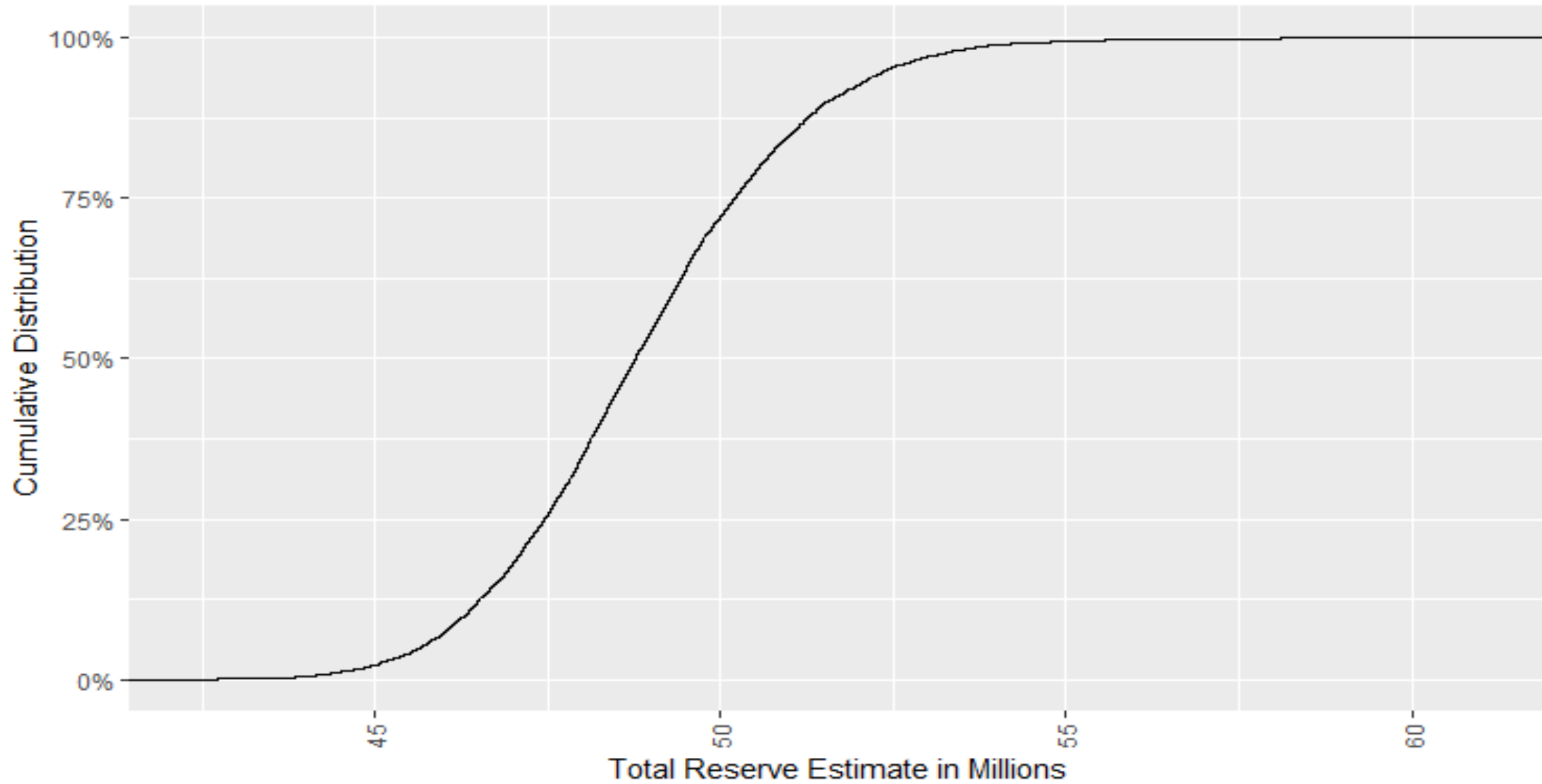
Dev_Yr vs. Distribution of Predicted Non-Linear Model



Development Year vs. Normalized Residuals Non-Linear Model



Total Reserve Estimate
Model Nonlinear
Simulated Future Inflation



Checking Prior Distribution for Plausibility

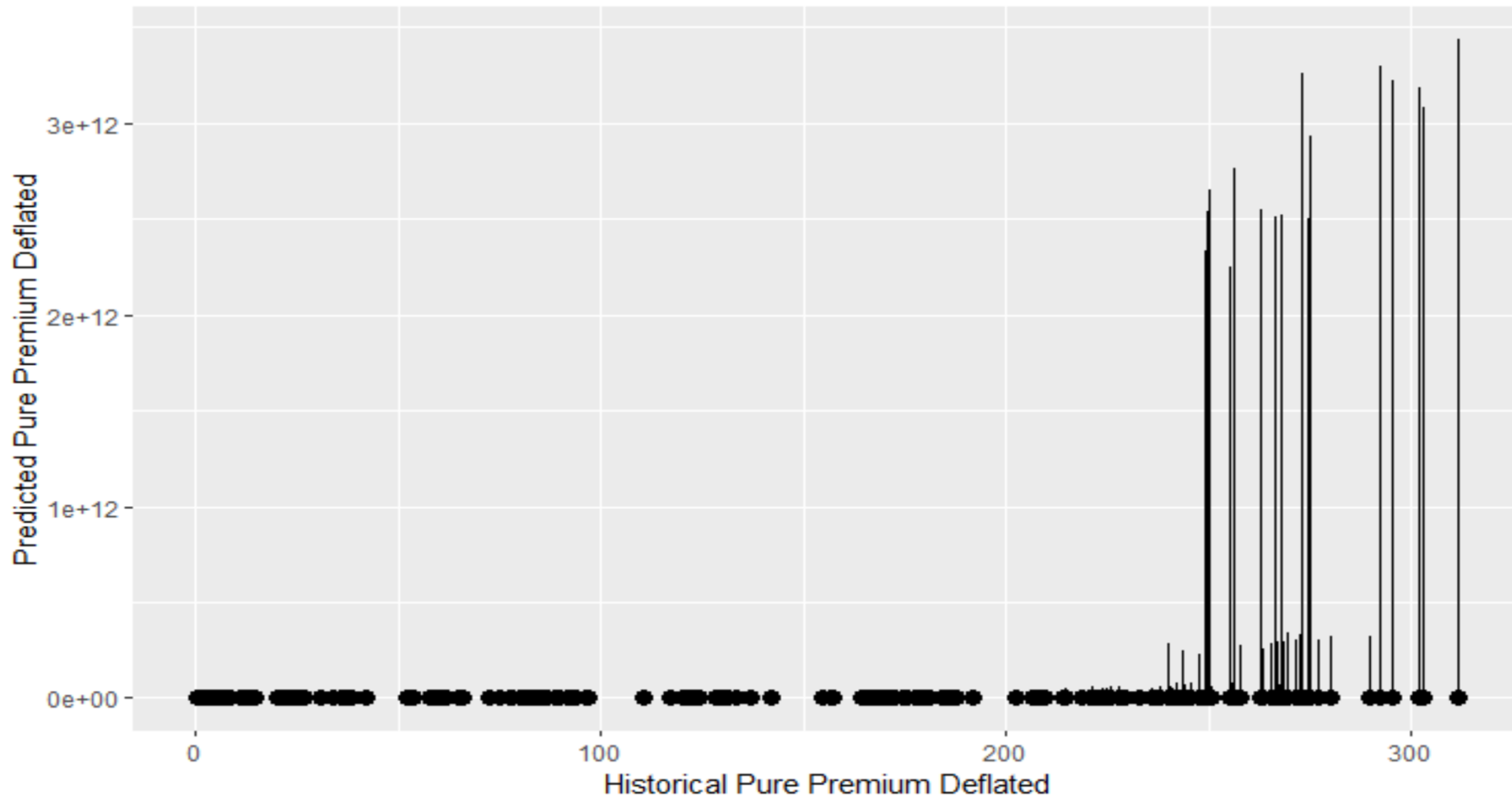
- Generally, a good practice to simulate prior distributions
 - Essential for some model forms and small data sets
 - Sometimes, just checking mean parameters in EXCEL is enough
 - In larger data sets, prior distribution effects may not matter
- Brms offers option to simulate population level prior distributions
- Example follows to show how software today reduces time burden to create Bayesian MCMC models

Nonlinear Model brms Instructions (Use Only Prior Distributions)

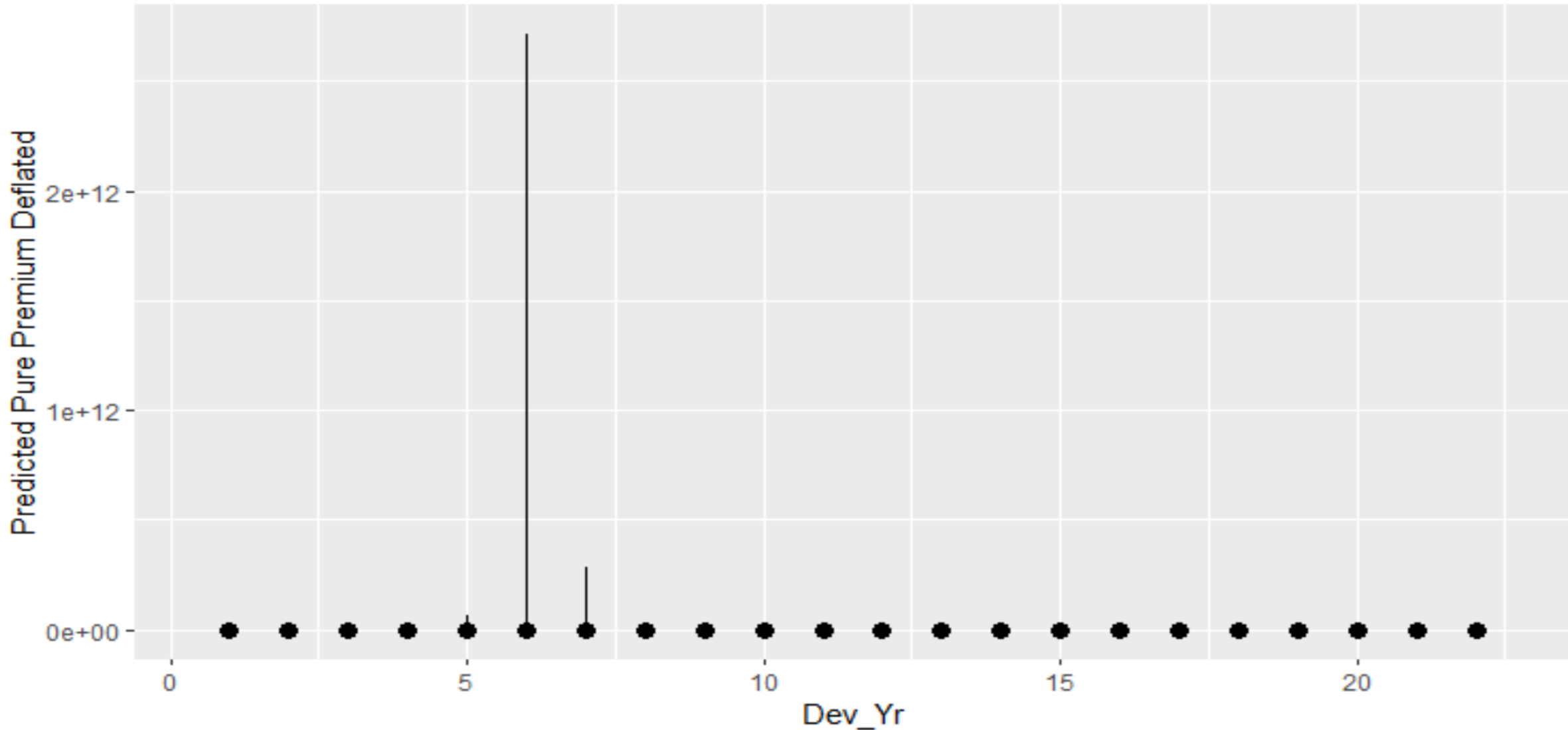
```
Model_NL_4_prior <- brm(  
  bf(Trended_Incr_PP_Def ~  
    ( exp(dev1 + b1 *Dev_Yr_5_Cap  
      + b2*Ln_Dev_Yr  
      + b3 * Dev_Yr_5_Spline  
      + b4 * Dev_Yr_5_Spline_Sqrd  
      + b5 * Dev_Yr_GT_1  
    )  
    +infl*Cal_Yr_Time),  
  
  b1 +b2 +b3 + +b4+ b5 +infl ~ 1,dev1 ~ 1 ,  
  sigma ~ 0 + Intercept +Dev_Yr_10_Spline_Ln +Dev_Yr_10_Cap,  
  nl = TRUE),  
  sample_prior = "only",  
  data = Train_Triangle_All_Operation, family = lognormal(),  
  save_pars = save_pars(all=TRUE),  
  prior= Non_linear_prior_4,  
  seed= 8603529,  
  iter= 4000,  
  control = list(adapt_delta = 0.99,  
    max_treedepth =12))
```

Include statement
"sample_prior = "only"
Drop random effects
(group) component of
variable dev1

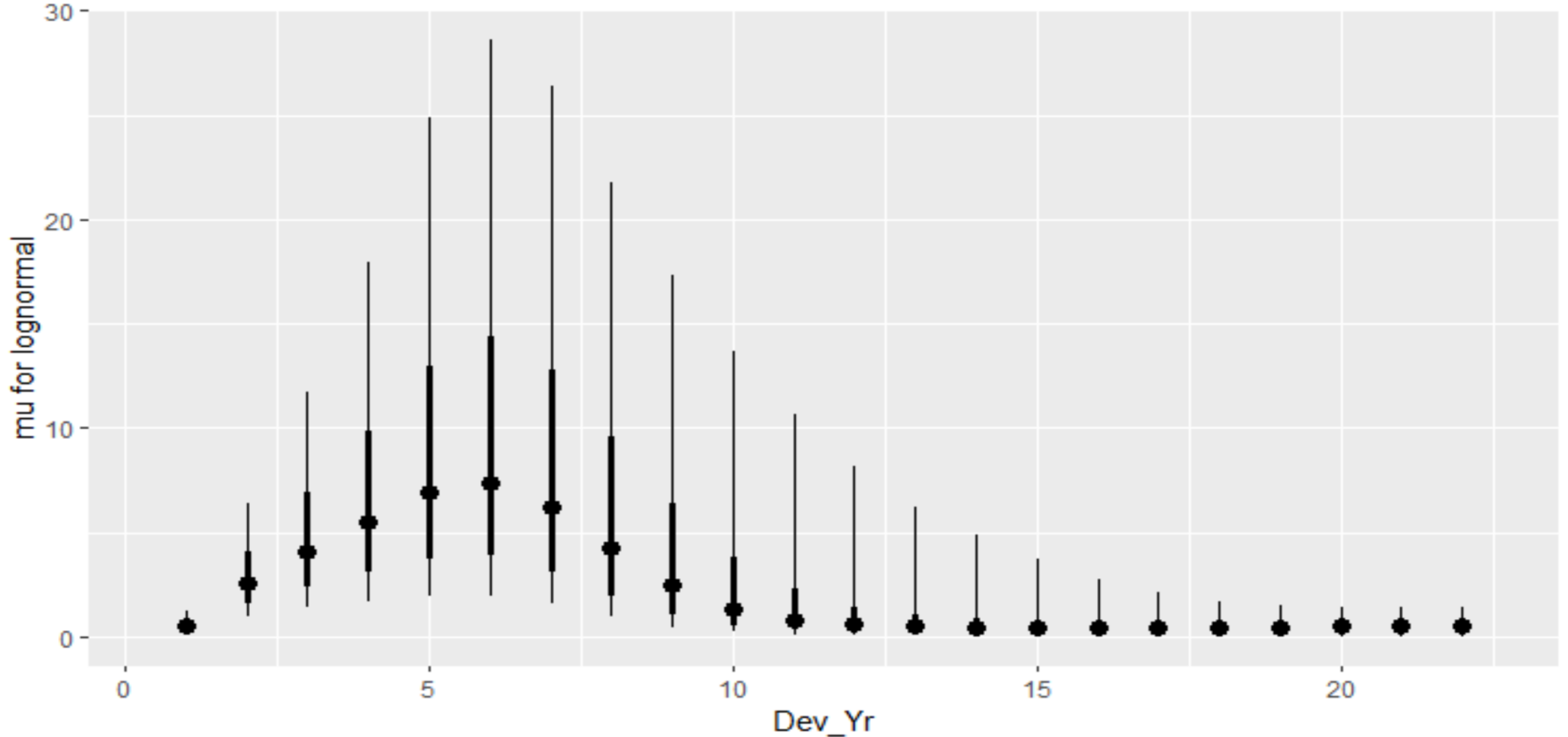
Historical Pure Premium vs. Distribution of Predicted Using Only Prior Non-Linear Model



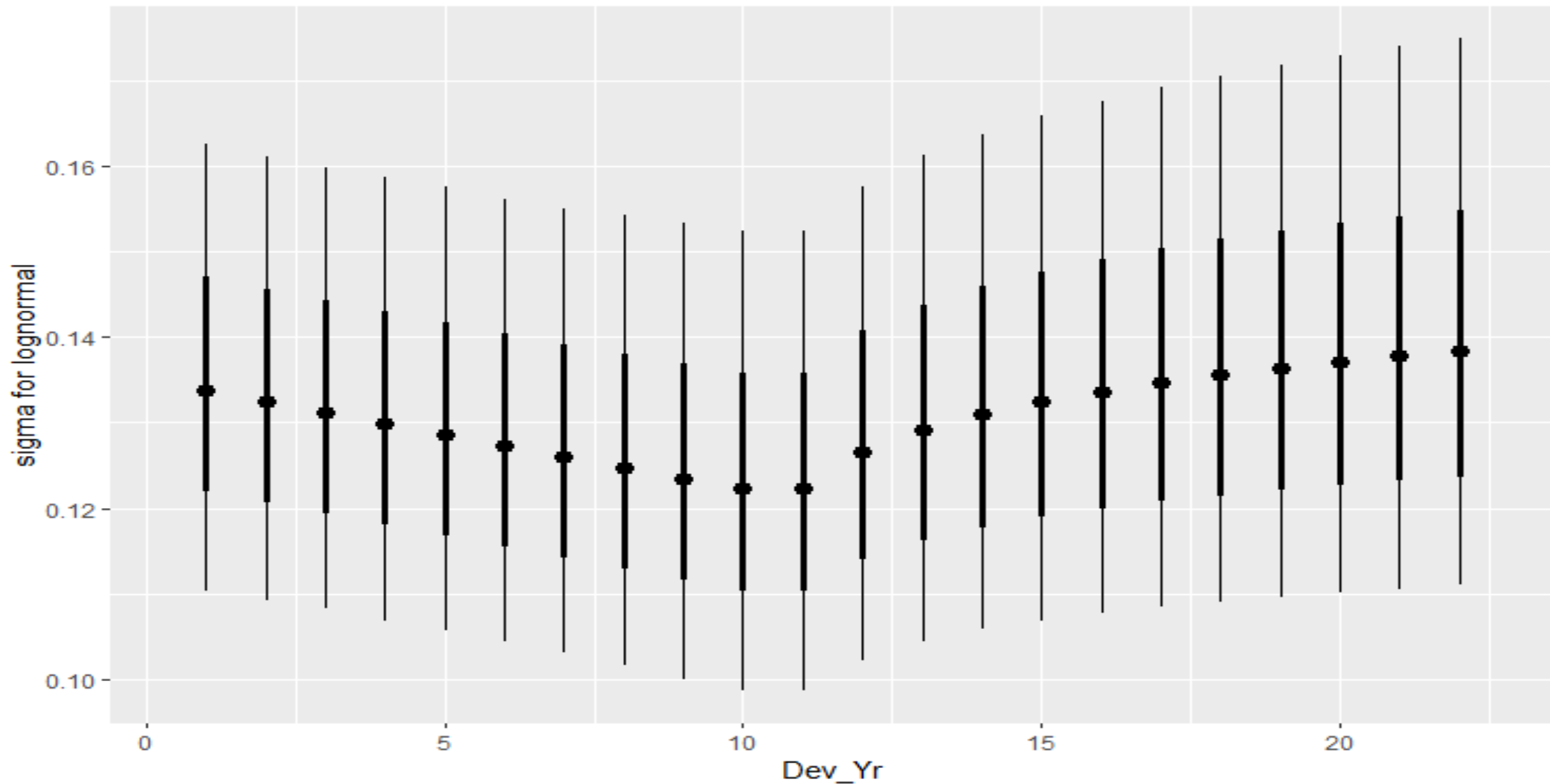
Dev_Yr vs. Distribution of Predicted Using Only Prior Non-Linear Model



Historical Pure Premium vs. Distribution of Predicted mu using only prior Non-Linear Model



Historical Pure Premium vs. Distribution of Predicted sigma using only prior Non-Linear Model



Comments on Generalized Additive Model Example

- Excluded from list of models to be compared given unresolved divergent error message
- Included in presentation:
 - Illustrates STAN/brms does give useful feedback on MCMC integrity
 - Provides illustration of bayesplot and ShinyStan diagnostics
 - Demonstrates why one has to pay attention to warnings
- Bayesplot is a package one can invoke within Rstudio with a range of diagnostics
- ShinyStan is a package one can launch which produces a set of prepackaged diagnostics

Generalized Additive Model Prior Definitions

```
GAM_prior <- c(prior(normal(.02,.01),class=b, coef=Cal_Yr_Time),
               prior(normal(0,.25),class=b, coef=Intercept),
               prior(normal(-2,.5),class=b ,coef =Intercept, dpar=sigma),
               prior(normal(-.02,.01),class=b, coef=Dev_Yr_10_Cap,dpar=sigma),
               prior(student_t(3,.02,.01),class=b, coef=Dev_Yr_10_Spline,dpar=sigma))
```

```
prior_summary(Model_GAM)
```

	prior	class	coef	group	resp	dpar	nlpar	lb	ub	source
	(flat)	b								default
	normal(0.02, 0.01)	b	Cal_Yr_Time							user
	normal(0, 0.25)	b	Intercept							user
	(flat)	b	sDev_Yr_1							(vectorized)
	(flat)	b				sigma				default
	normal(-0.02, 0.01)	b	Dev_Yr_10_Cap			sigma				user
	student_t(3, 0.02, 0.01)	b	Dev_Yr_10_Spline			sigma				user
	normal(-2, 0.5)	b	Intercept			sigma				user
	student_t(3, 0, 2.5)	sd						0		default
	student_t(3, 0, 2.5)	sd		Acc_Yr				0		(vectorized)
	student_t(3, 0, 2.5)	sd	Intercept	Acc_Yr				0		(vectorized)
	student_t(3, 0, 2.5)	sds						0		default
	student_t(3, 0, 2.5)	sds	s(Dev_Yr, k = 3, m = 2)					0		(vectorized)

Complete set
of prior
specifications
including
defaults

Generalized Additive Model brms Instructions

```
Model_GAM <- brm(bf(Trended_Incr_PP_Def ~ 0 + Intercept + s(Dev_Yr, k=3, m=2)
  + Cal_Yr_Time + (1 || Acc_Yr),
  sigma ~ 0 + Intercept + Dev_Yr_10_Cap + Dev_Yr_10_Spline),
  iter = 4000,
  prior = GAM_prior,
  seed = 8603529,
  control = list(adapt_delta = .99,
    max_treedepth = 15),
  data = Train_Triangle_All_Operation, family = lognormal())
```

Generalized Additive Model formula for Dev_Yr invoked in “s(Dev_Yr, k=3, m=2)”

Generalized Additive Model Results Summary

```
Family: lognormal
Links: mu = identity; sigma = log
Formula: Trended_Incr_PP_Def ~ 0 + Intercept + s(Dev_Yr, k = 3, m = 2)
        + Cal_Yr_Time + (1 || Acc_Yr)
        sigma ~ 0 + Intercept + Dev_Yr_10_Cap + Dev_Yr_10_Spline
Data: Train_Triangle_All_Operation (Number of observations: 253)
Draws: 4 chains, each with iter = 4000; warmup = 2000; thin = 1;
       total post-warmup draws = 8000
```

Smooth Terms:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sds(sDev_Yr_1)	11.92	6.23	5.59	27.30	1.00	4251	4354

Group-Level Effects:

~Acc_Yr (Number of levels: 22)

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sd(Intercept)	2.83	0.54	1.93	4.05	1.01	993	1948

Population-Level Effects:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
Intercept	0.47	0.27	-0.06	1.01	1.00	4380	4602
Cal_Yr_Time	0.03	0.01	0.01	0.05	1.00	3852	4740
sigma_Intercept	0.41	0.06	0.28	0.53	1.00	4571	3990
sigma_Dev_Yr_10_Cap	-0.07	0.01	-0.09	-0.05	1.00	5081	4702
sigma_Dev_Yr_10_Spline	-0.14	0.03	-0.19	-0.07	1.00	2905	1705
sDev_Yr_1	1.53	0.08	1.38	1.68	1.00	3360	3652

Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

warning message:

There were 3 divergent transitions after warmup. Increasing adapt_delta above 0.99 may help. See <http://mc-stan.org/misc/warnings.html#divergent-transitions-after-warmup>

Checking Syntax for GAM Model

The screenshot displays the RStudio interface with the following components:

- Source Editor:** Contains R code for defining priors and fitting a Bayesian GAM model using `brm` from the `brms` package. The code includes priors for `Cal_Yr_Time`, `Intercept`, `Dev_Yr_10_Cap`, and `Dev_Yr_10_Spline`. The model formula is `bf(Trended_Incr_PP_Def ~ 0 + Intercept + s(Dev_Yr, k=3, m=2) + Cal_Yr_Time + (1|Acc_Yr))`.
- Console:** Shows the execution of the code, including the compilation of the Stan program and the start of MCMC sampling. It displays a warning message about divergent transitions after warmup.
- Environment Pane:** Shows the output of the `summary(fit_smooth1)` command, providing a detailed summary of the model fit, including the family (gaussian), formula, data, and smooth terms.

```
GAM_prior <- c(prior(normal(.02,.01),class=b, coef=Cal_Yr_Time),
               prior(normal(0,.25),class=b, coef=Intercept),
               prior(normal(-2,.5),class=b,coef =Intercept, dpar=sigma),
               prior(normal(-.02,.01),class=b, coef=Dev_Yr_10_Cap,dpar=sigma),
               prior(student_t(3,.02,.01),class=b, coef=Dev_Yr_10_Spline,dpar=sigma))

Model_GAM <- brm(bf(Trended_Incr_PP_Def ~ 0 +Intercept + s(Dev_Yr, k=3, m=2)
                  +Cal_Yr_Time +(1|Acc_Yr),
                  sigma ~ 0 + Intercept + Dev_Yr_10_Cap + Dev_Yr_10_Spline),
                iter = 4000,
                prior= GAM_prior,
                seed= 8603529,
                control = list(adapt_delta = .99,
                              max_treedepth=15),
                data = Train_Triangle_All_Operation, family = lognormal())

launch_shinystan(Model_GAM)

prior_summary(Model_GAM)

summary(Model_GAM)
```

```
R 4.0.5 · C:/Users/mikep/Spring_2023_Meeting/
+ iter = 4000,
+ prior= GAM_prior,
+ seed= 8603529,
+ control = list(adapt_delta = .99,
+               max_treedepth=15),
+ data = Train_Triangle_All_Operation, family = lognormal()
Compiling Stan program...
Start sampling
Warning messages:
1: There were 3 divergent transitions after warmup. See
https://mc-stan.org/misc/warnings.html#divergent-transitions-after-warmup
to find out why this is a problem and how to eliminate them.
2: Examine the pairs() plot to diagnose sampling problems
```

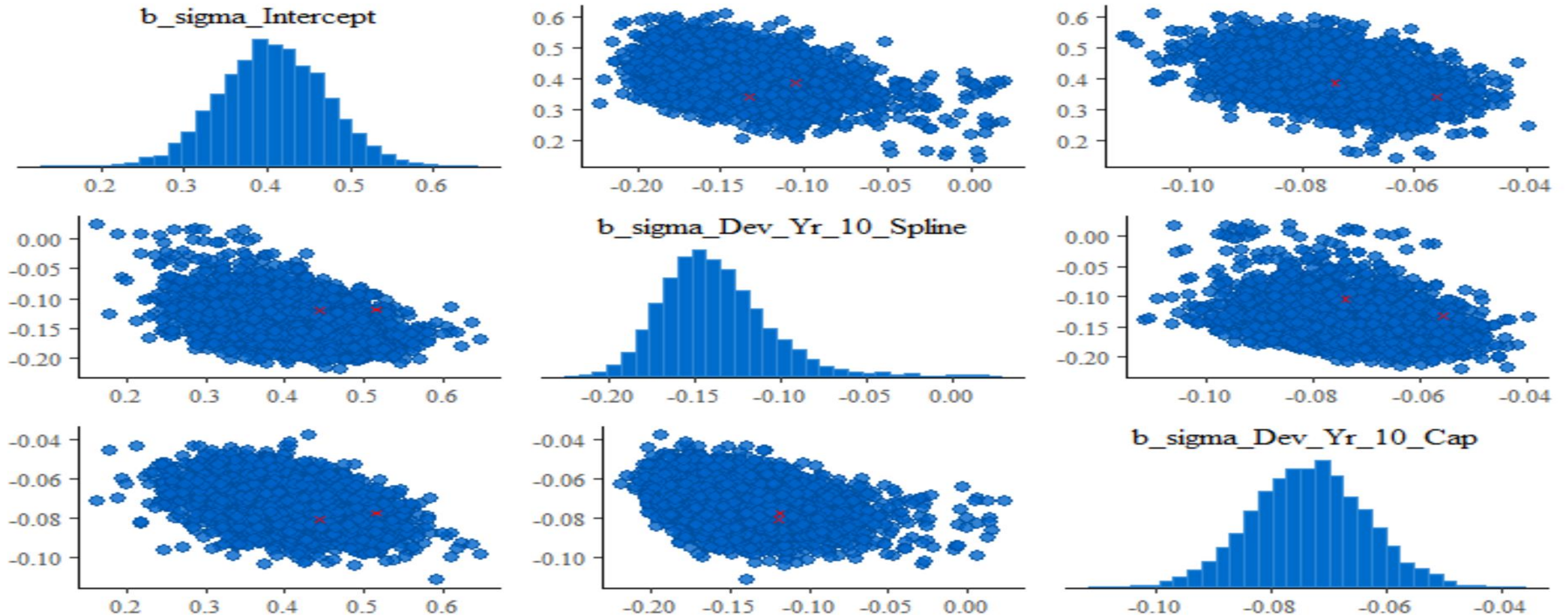
```
Estimating Distributional Models with brms
1 8.515469 0.1905577 0.08812703 0.74807568 0.42217250 7.423304
2 5.802847 0.6493315 0.09816140 0.79910490 0.30320760 10.116929
3 15.045523 0.9683143 0.79207055 0.99040837 0.75895849 14.073861
4 20.458492 0.2172666 0.78221959 0.08559773 0.01160683 20.604627
5 11.436937 0.1583259 0.62951180 0.39425608 0.45021855 11.905628
6 15.195921 0.8157348 0.35058777 0.38213741 0.64069148 13.917235
```

```
summary(fit_smooth1)

Family: gaussian
Links: mu = identity; sigma = log
Formula: y ~ s(x1) + s(x2) + (1 | fac)
sigma ~ s(x0) + (1 | fac)
Data: dat_smooth (Number of observations: 200)
Draws: 2 chains, each with iter = 2000; warmup = 1000; thin = 1;
total post-warmup draws = 2000

Smooth Terms:
Tail_ESS Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS
```

Generalized Additive Model bayesplot example for pairs() graph



Save & Close

SHINYSTAN DIAGNOSE ESTIMATE EXPLORE MORE ▾



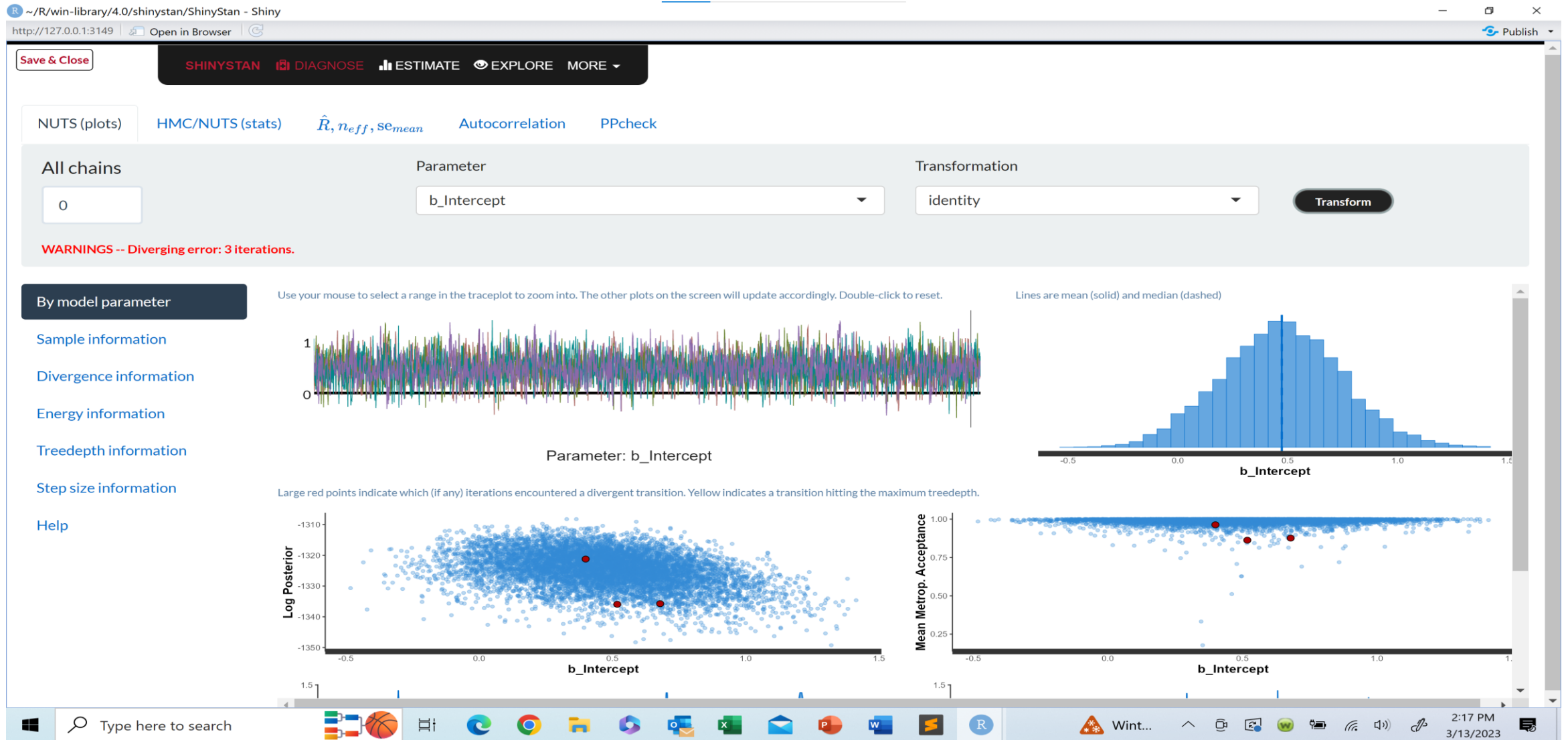
DIAGNOSE

ESTIMATE

EXPLORE

MORE

Generalized Additive Model ShinyStan



Generalized Additive Model ShinyStan

~/R/win-library/4.0/shinystan/ShinyStan - Shiny
http://127.0.0.1:3149 Open in Browser Publish

Save & Close SHINYSTAN DIAGNOSE ESTIMATE EXPLORE MORE

NUTS (plots) HMC/NUTS (stats) \hat{R} , n_{eff} , se_{mean} Autocorrelation PPcheck

All chains 0 Parameter b_Intercept Transformation identity Transform

WARNINGS -- Diverging error: 3 iterations.

By model parameter

Sample information

Divergence information

Energy information

Treedepth information

Step size information

Help

Use your mouse to select a range in the traceplot to zoom into. The other plots on the screen will update accordingly. Double-click to reset.

Lines are mean (solid) and median (dashed)

Large red points indicate which (if any) iterations encountered a divergent transition. Yellow indicates a transition hitting the maximum treedepth.

Log Posterior

Log Posterior

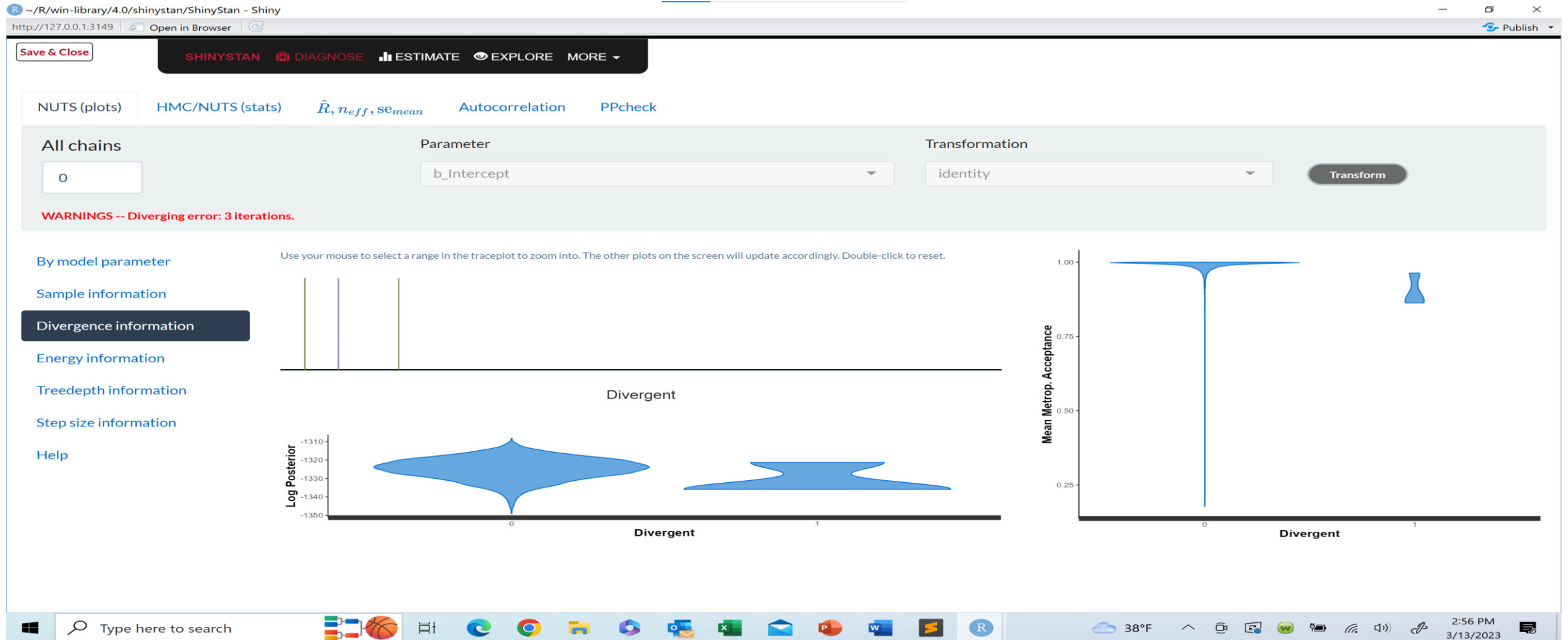
Log Posterior

Mean Metrop. Acceptance

Mean Metrop. Acceptance

Mean Metrop. Acceptance

Generalized Additive Model ShinyStan



Generalized Additive Model ShinyStan



Generalized Additive Model ShinyStan

The screenshot shows a web browser window displaying the ShinyStan application. The browser's address bar shows the URL `http://127.0.0.1:3149`. The application's navigation bar includes buttons for "SHINYSTAN", "DIAGNOSE", "ESTIMATE", "EXPLORE", and "MORE". The main content area is titled "Glossary" and features a sidebar on the left with a list of terms: "Effective sample size", "Monte Carlo uncertainty", "Rhat", "No-U-Turn Sampler (NUTS)", "accept_stat", "divergent" (highlighted), "energy", "stepsize", "n_leapfrog", and "treedepth". The main content area for "divergent" includes a "Quick definition" section stating that it represents the number of leapfrog transitions with diverging error, and a "More details" section explaining that divergent transitions are caused by numerical instability and can lead to biased results. A citation for Betancourt (2017) is provided at the bottom of the main content area.

Save & Close

SHINYSTAN DIAGNOSE ESTIMATE EXPLORE MORE

Glossary

Effective sample size

Monte Carlo uncertainty

Rhat

No-U-Turn Sampler (NUTS)

accept_stat

divergent

energy

stepsize

n_leapfrog

treedepth

divergent

Quick definition

The number of leapfrog transitions with diverging error. Because NUTS terminates at the first divergence this will be either 0 or 1 for each iteration. The average value of `divergent` over all iterations is therefore the proportion of iterations with diverging error.

More details

When numerical issues arise during the evaluation of the parameter Jacobians or the model log density, an exception is raised in the underlying code and the current expansion of the Hamiltonian forward and backward in time is halted. This is marked as a divergent transition.

The primary cause of divergent transitions in Euclidean HMC (other than bugs in the model code) is numerical instability in the leapfrog integrator used to simulate the Hamiltonian evaluation. The fundamental problem is that a fixed step size is being multiplied by the gradient at a particular point, to determine the next simulated point. If the stepsize is too large, this can overshoot into ill-defined portions of the posterior.

If there are (post-warmup) divergences then the results may be biased and should not be used.

In some cases, simply lowering the initial step size and increasing the target acceptance rate will keep the step size small enough that sampling can proceed.

The exact cause of each divergent transition is printed as a warning message in the output console. This can be useful in cases where managing the step size is insufficient. In such cases, a reparameterization is often required so that the posterior curvature is more manageable; see the section about Neal's Funnel in the Stan manual for an example.

For more details see [Betancourt, M. \(2017\). A conceptual introduction to Hamiltonian Monte Carlo.](#)

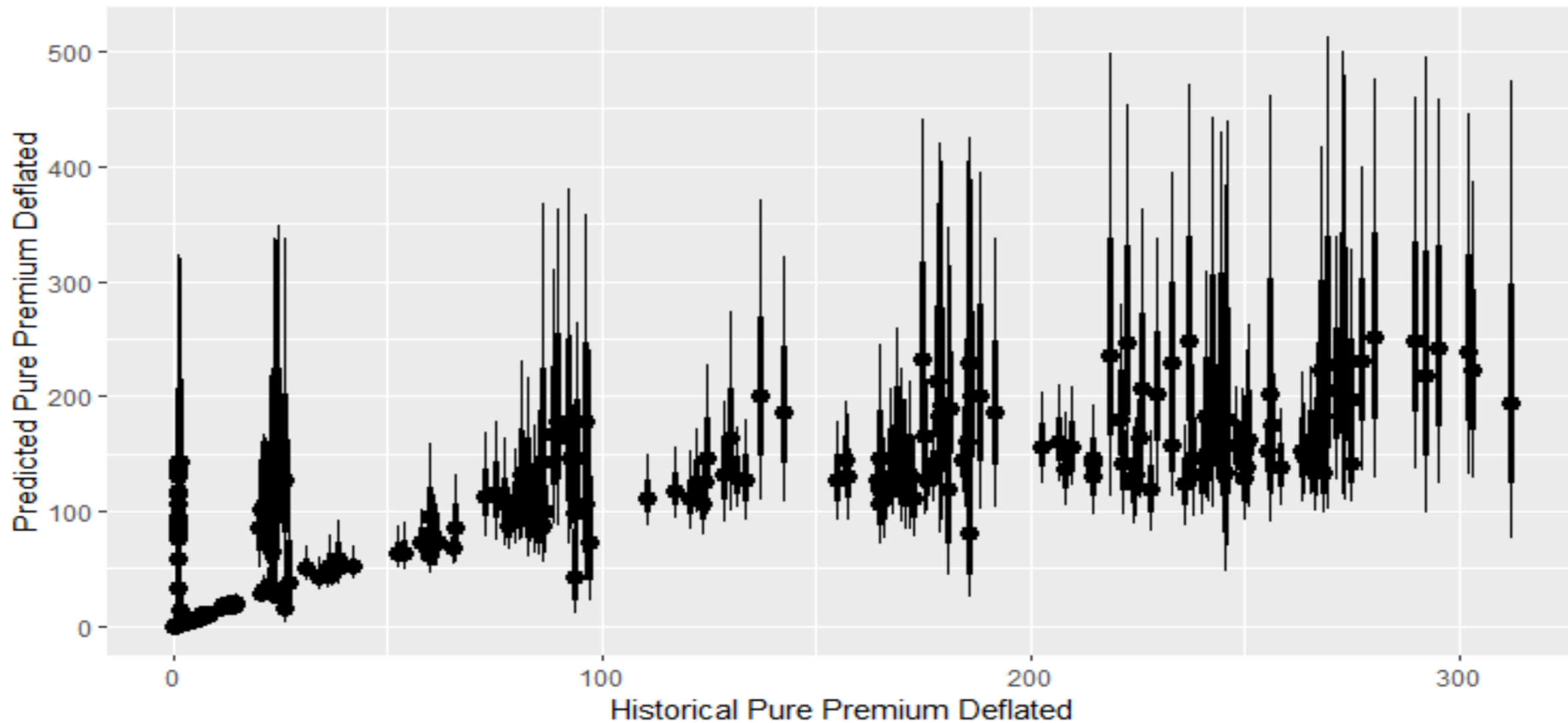
Glossary entries are compiled (with minor edits) from various excerpts of the Stan Modeling Language User's Guide and Reference Manual (CC BY (v3))

Type here to search

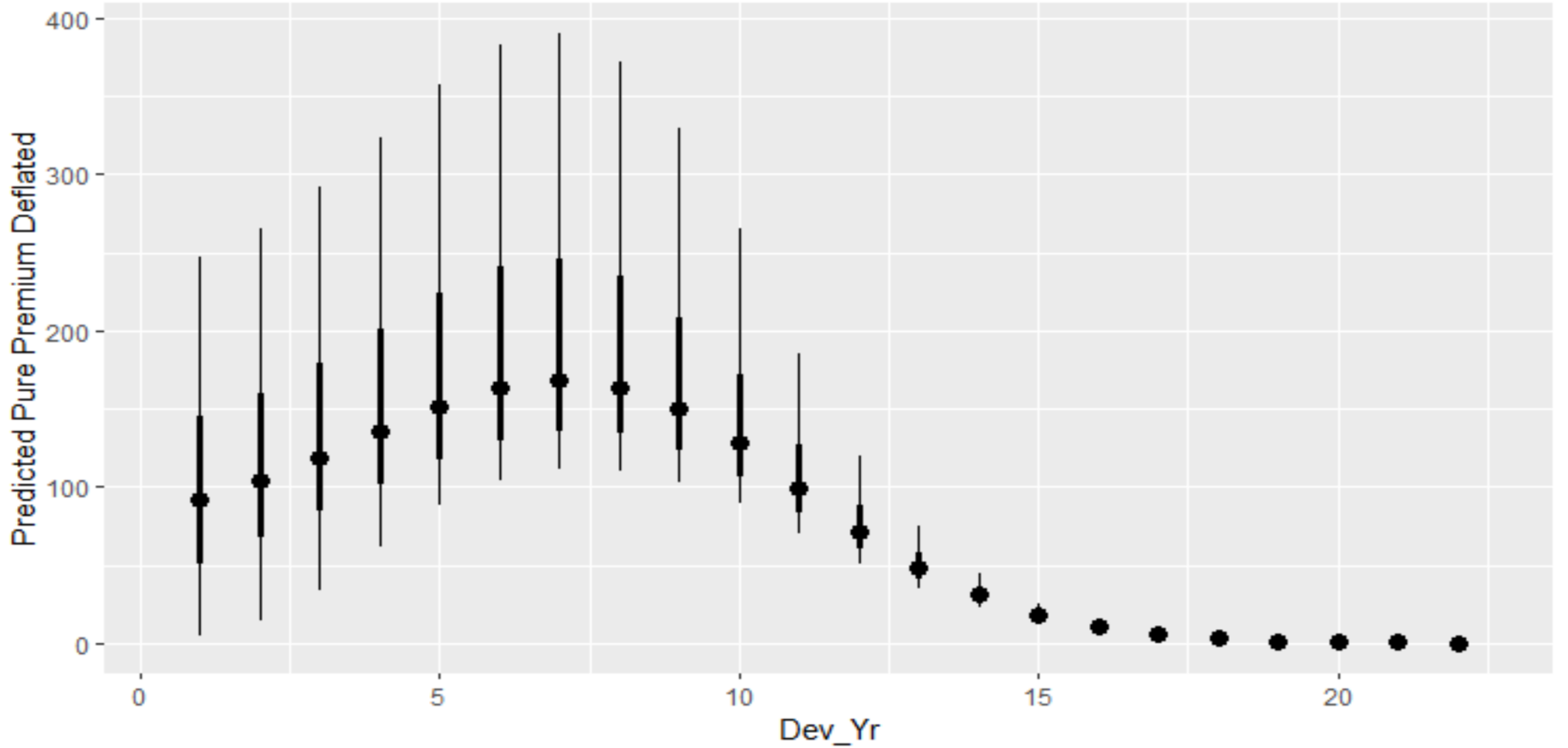
38°F

3:12 PM
3/13/2023

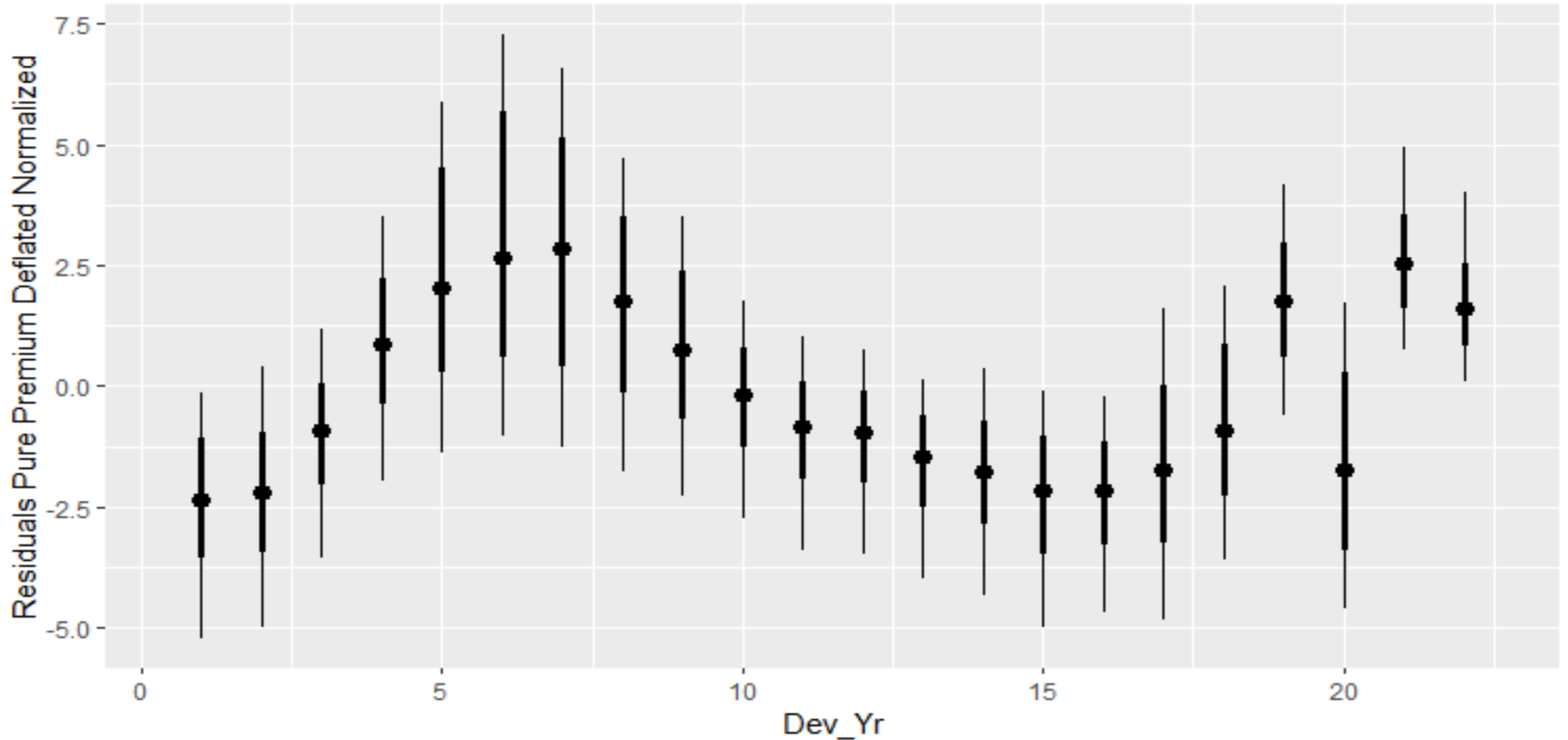
Historical Pure Premium vs. Distribution of Predicted Model GAM



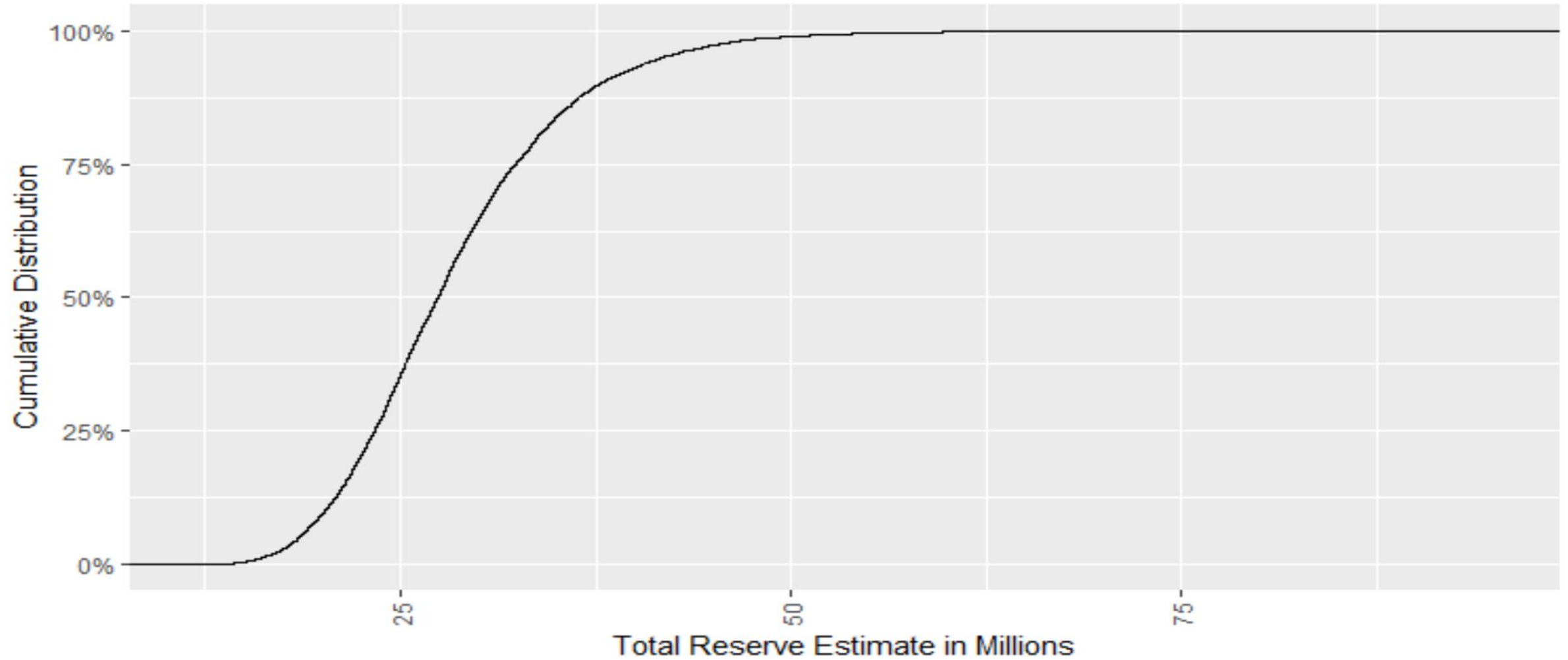
Historical Pure Premium vs. Distribution of Predicted Model_GAM



Development Year vs. Normalized Residuals GAM Model



Total Reserve Estimate
Model GAM
Simulated Future Inflation



Model Selection

Bayesian MCMC Model Comparison Metrics

- PSIS

- Pareto Smoothed Importance Sampling Cross Validation
- Approximation of leave one out cross validation
- Uses individual observation log pointwise predictive distribution details
- Uses relative effect of a given observation as input to weighting
- Avoids computational cost of literal leave one out cross validation
- Diagnostics on safety of using approximation furnished

- WAIC:

- Widely Applicable Information Criteria
- Uses log pointwise predictive distribution details
- Similar to AIC in that there is a penalty term
- Informed guess on out-of-sample KL divergence
- Useful for comparison between models
- Diagnostics on safety of using furnished

Example of Criteria for Comparison Safeguards

```
Poly_1_loo <- add_criterion(Model_Polynomial_1,c("loo", "waic"),
+                          moment_match=TRUE)
Warning messages:
1: Some Pareto k diagnostic values are slightly high. See help('pareto-k-diagnostic') for details.

   2:00
11 (4.3%) p_waic estimates greater than 0.4. We recommend trying loo instead.
>
> Poly_1_loo_prep <- loo(Model_Polynomial_1)
>
> Poly_1_loo_mm <- loo_moment_match(Model_Polynomial_1, loo=Poly_1_loo_prep)
Warning message:
Some Pareto k diagnostic values are slightly high. See help('pareto-k-diagnostic') for details.

>
> Poly_1_loo_mm

Computed from 8000 by 253 log-likelihood matrix

      Estimate      SE
elpd_loo   -801.1  25.0
p_loo       18.7   3.3
looic       1602.2 49.9
-----
Monte Carlo SE of elpd_loo is 0.1.

Pareto k diagnostic values:

```

		Count	Pct.	Min. n_eff
(-Inf, 0.5]	(good)	250	98.8%	2273
(0.5, 0.7]	(ok)	3	1.2%	137
(0.7, 1]	(bad)	0	0.0%	<NA>
(1, Inf)	(very bad)	0	0.0%	<NA>

All Pareto k estimates are ok ($k < 0.7$).

Model Comparison Using loo (approximate leave one out)

- Comparison of log pointwise predictive density
- Need both the difference and the `se_diff` (standard error of difference)
- Comparison is to best model Polynomial 2 in this case (0.0 is best)
- Sometimes standard error relative to difference is so large that differences may not be large enough to say one model is better

```
> model_compare <- loo_compare(Poly_1_loo, Poly_2_loo,  
+                               Nonlinear_loo,  
+                               criterion = "loo")
```

```
> model_compare
```

	elpd_diff	se_diff
Poly_2_loo	0.0	0.0
Nonlinear_loo	-94.3	13.6
Poly_1_loo	-132.2	16.0

Model Averaging Results

- Weight shows the sampling of forecast losses to obtain the best results in terms of estimated expected log pointwise predictive distribution results.
- The weights are selected to obtain the optimal approximate leave one out (loo) result from the weighted forecast.
- Note that weights for the predicted forecast not the parameter values.

```
loo_model_weights(Model_NL_4, Model_Polynomial_1, Model_Polynomial_2)
```

```
Method: stacking
```

```
-----  
                weight  
Model_NL_4      0.053  
Model_Polynomial_1 0.000  
Model_Polynomial_2 0.947
```

```
Warning message:
```

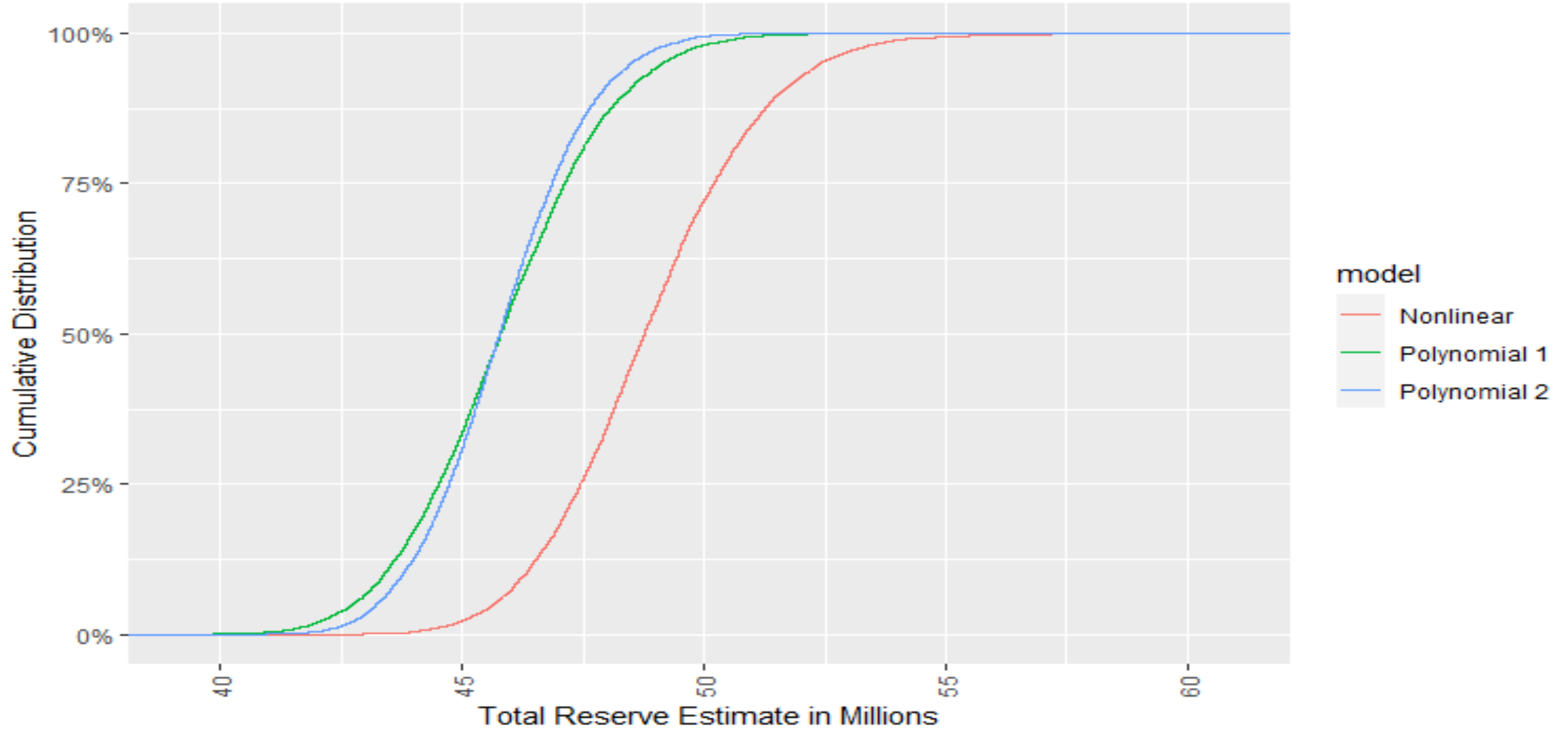
```
Some Pareto k diagnostic values are slightly high.  
See help('pareto-k-diagnostic') for details.
```

Model Forecast Comparison (millions)

	Polynomial 1	Polynomial 2	Non_Linear	GAM
Min	39.21	40.09	42.38	11.85
1st Quartile	44.52	44.67	47.46	23.33
Median	45.81	45.76	48.77	27.39
Mean	45.85	45.77	48.88	28.4
3rd Quartile	47.12	46.8	50.21	32.36
Max	53.67	51.61	61.05	95.17

Note: Forecasts assume base inflation continues with lognormal distribution $\mu = .03$ and $\sigma = .02$ and loss cost multiplier continues with lognormal distribution $\mu = .01$ and $\sigma = .01$.

Total Reserve Estimate By Model



Conclusion

- Bayesian probability models are not new
 - Reverend Thomas Bayes in 1754
- Combination of software, algorithm development and computing power makes application practical
 - Metropolis Hastings: First described in 1953 publication
 - Tidyverse programming concepts: ggplot2 first released in 2005
 - STAN (Hamiltonian): Released in 2015
 - Brms released in 2017
- Use requires picking up new vocabulary and concepts
- Value to actuaries:
 - Links credibility weighting to variety of parametric models
 - Document one's prior knowledge of the problem at hand

Appendix

Additional Resources

- **Brms**
 - brms: An R Package for Bayesian Multilevel Models using Stan, Paul-Christian Burkner, Journal of Statistical Software 2017
 - Advanced Bayesian Multilevel Modeling with the R Package brms by Paul-Christian Bürkner, The R Journal Vol. 10/1, July 2018
- **STAN**
 - <https://mc-stan.org/>
- **Rstudio**
 - <https://rstudio.com/>
- **Tidyverse**
 - Rstudio help
 - R for Data Science: Import, Tidy, Transform, Visualize, and Model Data 1st Edition, Hadley Wickham, Garret Golemund
- **Bayesian MCMC textbooks**
 - Statistical Rethinking: A Bayesian Course with Examples in R and STAN (Chapman & Hall/CRC Texts in Statistical Science) 2nd Edition, by Richard McElreath
 - Bayesian Data Analysis (Chapman & Hall/CRC Texts in Statistical Science) 3rd Edition, Gelman et.al.
- **Tidybayes**
 - <https://cran.r-project.org/web/packages/tidybayes>

Create Count and Severity Distributions (Page 1)

```
Change_Operation <- Triangle_Dates_Operation %>%
  group_by(Acc_Yr, Dev_Yr) %>%
  mutate(Ln_Dev_Yr = log(Dev_Yr),
         p_rate = case_when(
           Acc_Yr == 2000 ~ exp(-.98 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2001 ~ exp(-.97 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2002 ~ exp(-.98 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2003 ~ exp(-.99 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2004 ~ exp(-1.01 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2005 ~ exp(-1.02 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2006 ~ exp(-1 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2007 ~ exp(-.98 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2008 ~ exp(-.99 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2009 ~ exp(-1.02 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2010 ~ exp(-1.01 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2011 ~ exp(-.98 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2012 ~ exp(-.99 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2013 ~ exp(-1 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2014 ~ exp(-1.01 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2015 ~ exp(-1.02 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2016 ~ exp(-.97 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2017 ~ exp(-.98 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2018 ~ exp(-1.02 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2019 ~ exp(-1.03 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2020 ~ exp(-.99 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
           Acc_Yr == 2021 ~ exp(-.98 + Dev_Yr * (-.30) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
         ),
         # p_rate = exp(-1 + Dev_Yr * (-.35) + Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (1.0)),
         p_rate_adj = Rptd_Cnt * p_rate,
```

Create Count and Severity Distributions (Page 2)

```
mu_In = case_when(
  Acc_Yr == 2000 ~ (1.97+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2001 ~ (1.98+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2002 ~ (1.99+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2003 ~ (2.02+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2004 ~ (2.03+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2005 ~ (2.01+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2006 ~ (1.97+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2007 ~ (1.98+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2008 ~ (1.99+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2009 ~ (2+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2010 ~ (2.02+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2011 ~ (2.03+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2012 ~ (2.04+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2013 ~ (1.99+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2014 ~ (2.01+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2015 ~ (2.02+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2016 ~ (1.98+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2017 ~ (1.97+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2018 ~ (2.02+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2019 ~ (2.04+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2020 ~ (1.98+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),
  Acc_Yr == 2021 ~ (2.02+ Dev_Yr * (-.6 )+ Dev_Yr_Sqrd * (0) + Ln_Dev_Yr * (4.5)),),
  sigma_In = exp(-2.0 +Dev_Yr *.015+ Dev_Yr_Sqrd * 0 + (.1)*Ln_Dev_Yr),
  Paid_Cnt = rpois(1,lambda =p_rate_adj) ,
  Expected_Payment= exp(mu_In + .5 * (sigma_In**2)),
  Mean_Payment = if_else (Paid_Cnt ==0, 0,
    mean(rlnorm(Paid_Cnt,meanlog = mu_In,sdlog=sigma_In))),
  Incr_Payment = Paid_Cnt * Mean_Payment,
  Pure_Prem =Incr_Payment/Rptd_Cnt)
```

Create Transformed Variables (Page 1)

```
Complete_Triangle_Operation <- left_join(Change_Operation,Cal_Yr_Trend_Acc, by=c('Cal_Yr')) %>%
  mutate(Trended_Incr_Payment = Incr_Payment * Accum_Loss_Cost_Trend_Mid,
         Trended_Incr_Payment_Def =Trended_Incr_Payment/Accum_Infl_Index_Mid,
         Trended_Incr_PP =Trended_Incr_Payment/Rptd_Cnt,
         Trended_Mean_Payment = Mean_Payment * Accum_Loss_Cost_Trend_Mid,
         Trended_Mean_Payment_Def = Trended_Mean_Payment/Accum_Infl_Index_Mid,
         Trended_Incr_PP_Def =Trended_Incr_Payment_Def/Rptd_Cnt,
         Paid_Cnt_Freq =Paid_Cnt/Rptd_Cnt,
         Ln_Paid_Cnt_Freq = log(Paid_Cnt_Freq),
         Inv_Rptd_Cnt =1/Rptd_Cnt,
         Ln_Trnd_Incr_PP =log(Trended_Incr_PP),
         Ln_Trnd_Incr_PP_Def =log(Trended_Incr_PP_Def),
         Ln_Mean_Payment_No_Trend =log(Mean_Payment),
         Ln_Mean_Payment_Def =log(Trended_Mean_Payment_Def),
         Cal_Yr_Time = Cal_Yr - 2000,
         Dev_Yr_4_Cap = if_else(Dev_Yr < 4, Dev_Yr,
                               as.integer(4) ),
         Dev_Yr_4_Spline=if_else(Dev_Yr < 4,0,
                               Dev_Yr-4),
         Dev_Yr_4_Cap_Sqrd = Dev_Yr_4_Cap * Dev_Yr_4_Cap ,
         Dev_Yr_4_Spline_Sqrd =Dev_Yr_4_Spline *Dev_Yr_4_Spline,
         Dev_Yr_5_Cap = if_else(Dev_Yr < 5, Dev_Yr,
                               as.integer(5) ),
         Dev_Yr_5_Spline=if_else(Dev_Yr < 5,0,
                               Dev_Yr-5),
         Dev_Yr_5_Spline_Sqrd = Dev_Yr_5_Spline* Dev_Yr_5_Spline,
         Dev_Yr_6_Cap = if_else(Dev_Yr < 6, Dev_Yr,
                               as.integer(6) ),
```

Create Transformed Variables (Page 2)

```
Dev_Yr_6_Cap_Sqrd = Dev_Yr_6_Cap * Dev_Yr_6_Cap ,
Dev_Yr_6_Spline=if_else(Dev_Yr < 6,0,
  Dev_Yr-6),
Dev_Yr_6_Spline_Sqrd = Dev_Yr_6_Spline* Dev_Yr_6_Spline,
Dev_Yr_6_Cap_Sqrd = Dev_Yr_6_Cap * Dev_Yr_6_Cap ,
Dev_Yr_GT_1 = if_else(Dev_Yr< 2,0,1),
Dev_Yr_1_Factor = as.factor(case_when(
  Dev_Yr == 1 ~ 1,
  Dev_Yr >1 ~2
)),
Dev_Yr_8_Cap = if_else(Dev_Yr < 8, Dev_Yr,
  as.integer(8) ),
Dev_Yr_8_Spline=if_else(Dev_Yr < 8,0,
  Dev_Yr-8),
Dev_Yr_8_Cap_Sqrd = Dev_Yr_8_Cap * Dev_Yr_8_Cap ,
Dev_Yr_8_Spline_Sqrd =Dev_Yr_8_Spline * Dev_Yr_8_Spline,
Dev_Yr_1_Factor = as.factor(case_when(
  Dev_Yr == 1 ~ 1,
  Dev_Yr >1 ~2
)),
Dev_Yr_3_Factor = as.factor(case_when(
  Dev_Yr == 1 ~ 1,
  Dev_Yr == 2 ~2,
  Dev_Yr == 3 ~3,
  Dev_Yr > 3 ~4
)),
Dev_Yr_2_Factor = as.factor(case_when(
  Dev_Yr == 1 ~ 1,
  Dev_Yr == 2 ~2,
  Dev_Yr > 2 ~3
)),
```


Create Transformed Variables (Page 3)

```
Dev_Yr_3_14_Factor = as.factor(case_when(  
  Dev_Yr == 1 ~ 1,  
  Dev_Yr == 2 ~ 2,  
  Dev_Yr == 3 ~ 3,  
  Dev_Yr == 4 ~ 4,  
  Dev_Yr == 5 ~ 4,  
  Dev_Yr == 6 ~ 4,  
  Dev_Yr == 7 ~ 4,  
  Dev_Yr == 8 ~ 4,  
  Dev_Yr == 9 ~ 4,  
  Dev_Yr == 10 ~ 4,  
  Dev_Yr == 11 ~ 4,  
  Dev_Yr == 12 ~ 4,  
  Dev_Yr == 13 ~ 4,  
  Dev_Yr > 13 ~ 5,  
)),  
Dev_Yr_2_14_Factor = as.factor(case_when(  
  Dev_Yr == 1 ~ 1,  
  Dev_Yr == 2 ~ 2,  
  Dev_Yr == 3 ~ 3,  
  Dev_Yr == 4 ~ 3,  
  Dev_Yr == 5 ~ 3,  
  Dev_Yr == 6 ~ 3,  
  Dev_Yr == 7 ~ 3,  
  Dev_Yr == 8 ~ 3,  
  Dev_Yr == 9 ~ 3,  
  Dev_Yr == 10 ~ 3,  
  Dev_Yr == 11 ~ 3,  
  Dev_Yr == 12 ~ 3,  
  Dev_Yr == 13 ~ 3,  
  Dev_Yr > 13 ~ 4,  
)),  
Dev_Yr_10_Spline = if_else(Dev_Yr < 10, 0,  
  Dev_Yr - 10),  
Dev_Yr_10_Cap = if_else(Dev_Yr < 10, Dev_Yr,  
  as.integer(10)),  
Dev_Yr_12_Cap = if_else(Dev_Yr < 10, Dev_Yr,  
  as.integer(12)),  
Dev_Yr_15_Cap = if_else(Dev_Yr < 15, Dev_Yr,  
  as.integer(15)),  
Dev_Yr_12_Spline = if_else(Dev_Yr < 12, 0,  
  Dev_Yr - 12),  
Dev_Yr_10_Spline_Cap_15 = case_when(  
  Dev_Yr < 10 ~ 0,  
  Dev_Yr < 16 ~ Dev_Yr - 10,  
  Dev_Yr >= 16 ~ 5),  
Dev_Yr_10_Spline_Ln = if_else(Dev_Yr < 11, 0,  
  log(Dev_Yr_10_Spline)),  
Dev_Yr_15_Spline = if_else(Dev_Yr < 15, 0,  
  Dev_Yr - 15),  
Dev_Yr_15_Spline_Ln = if_else(Dev_Yr < 16, 0,  
  log(Dev_Yr_15_Spline)))
```