

On the problem of parameter and model uncertainty in risk management

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Introduction

PhD at University of Lausanne, with supervisor Prof. Hansjörg Albrecher

- **Structured reinsurance deals with reference to relative market performance.** Published in *Insurance: Mathematics and Economics*.
L. Vincent, H. Albrecher et Y. Krvavych (2021)
- **An alternative path to extremes.** Preprint.
L. Vincent (2022)
- **Risk estimation under parameter and model uncertainty.**
Preprint.
L. Vincent (2022)
- **A note on risk assessment under parameter uncertainty.**
Preprint.
L. Vincent (2022)

Risk modelling 101

Risk estimation under parameter and model uncertainty.

A risk ... e.g. a potential financial loss, natural catastrophe, etc.

- Modelled as a random variable ... $X \sim F$
- $F(x) = \mathbb{P}(X \leq x)$... distribution of X
- $F \Rightarrow \text{VaR, ES, etc.}$... risk measures

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In practice ... F is unknown

- Data $\Rightarrow \hat{F}$... estimated/fitted distribution
- $\hat{F} \Rightarrow \widehat{\text{VaR}}, \widehat{\text{ES}}$, etc. ... estimated risk measures

Aleatoric vs. epistemic uncertainty

Two types of uncertainty

- $Y \sim F$... aleatoric uncertainty
- $F \neq \hat{F}$... epistemic uncertainty

Manage risk using a fitted distribution \hat{F} ?

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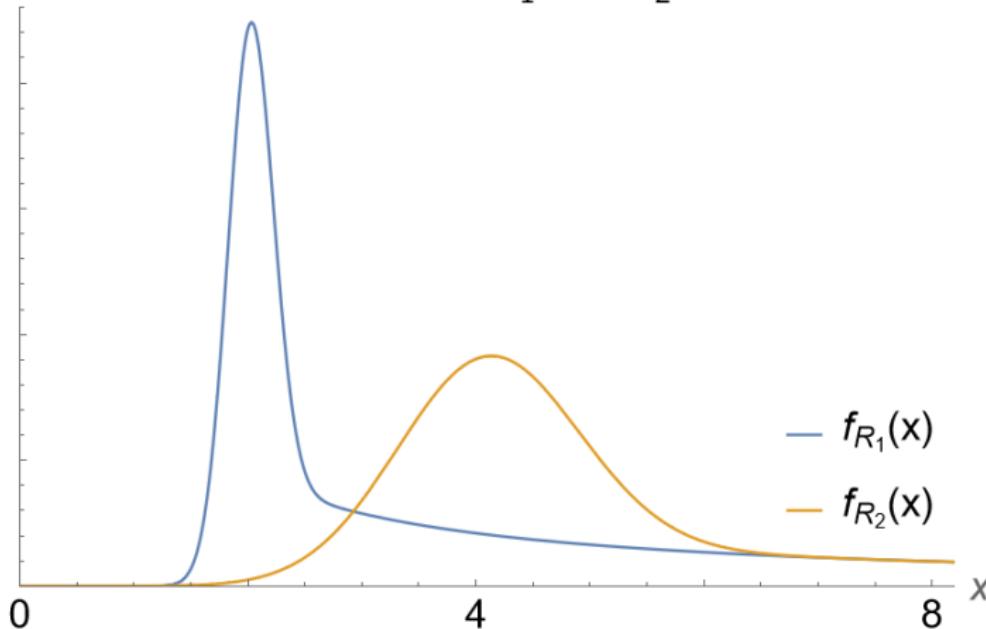
Manage risk using a fitted distribution \hat{F} ?

For instance

- X and Y ... two independent risks
- $\hat{F}_X = \hat{F}_Y$... with identical fitted distributions
- $n_X > n_Y$... but unequal sample sizes

Aleatoric vs. epistemic uncertainty – Another example

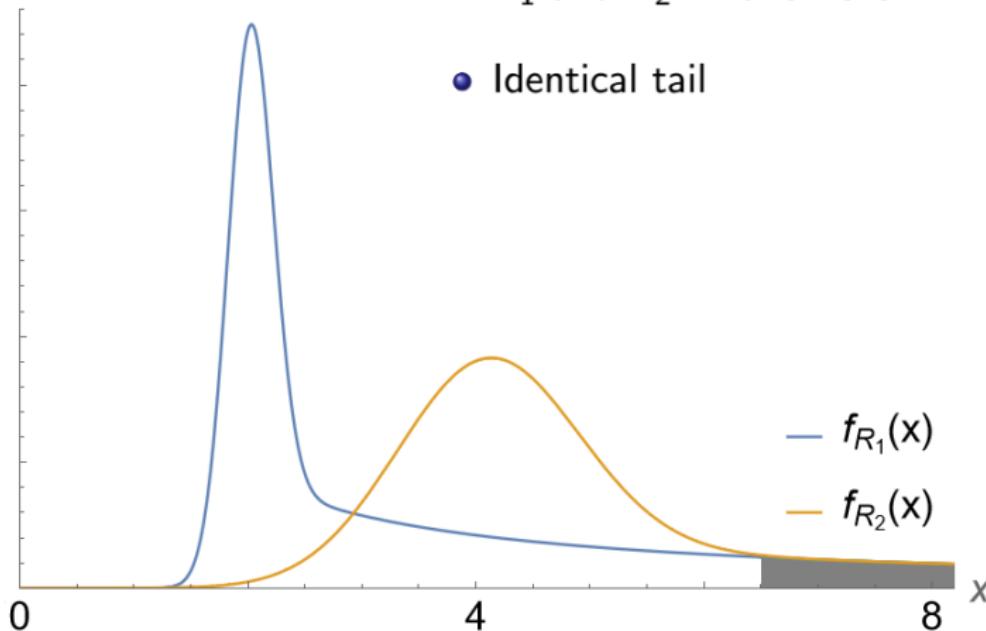
- R_1 and R_2 ... two risks



Aleatoric vs. epistemic uncertainty – Another example

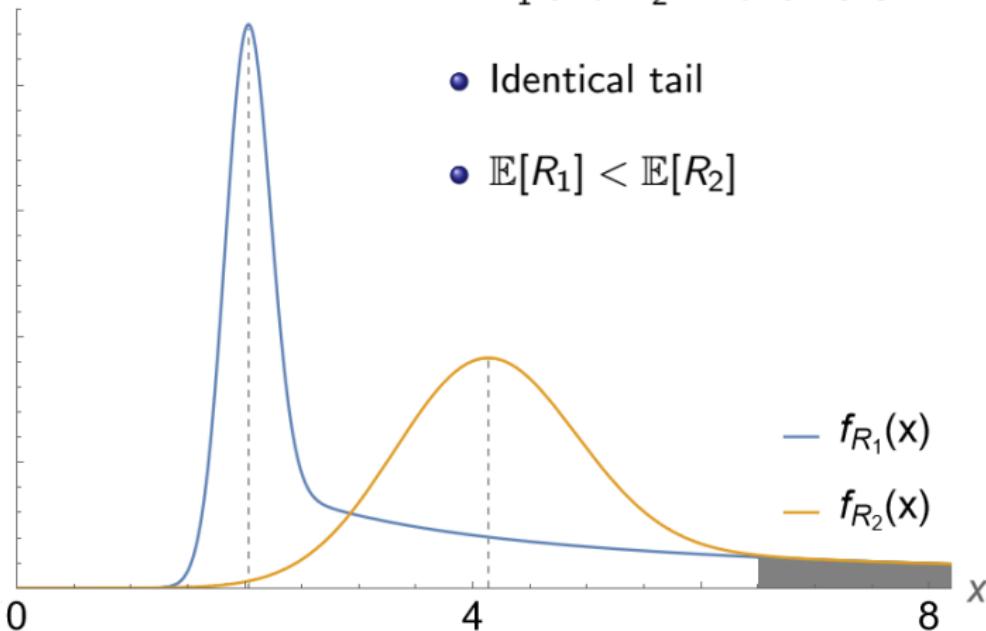
- R_1 and R_2 ... two risks

- Identical tail



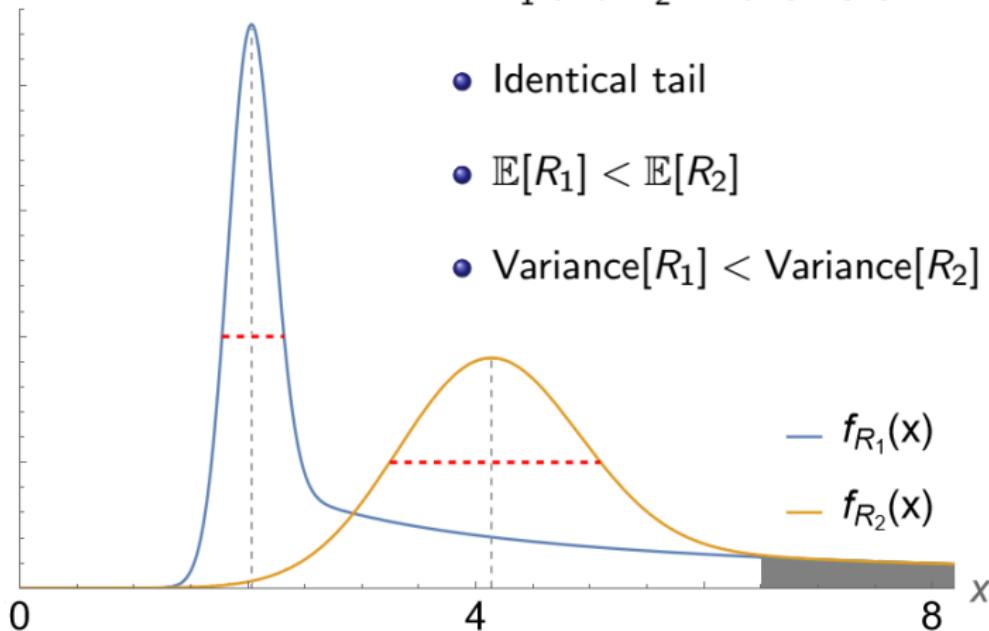
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- $\mathbb{E}[R_1] < \mathbb{E}[R_2]$



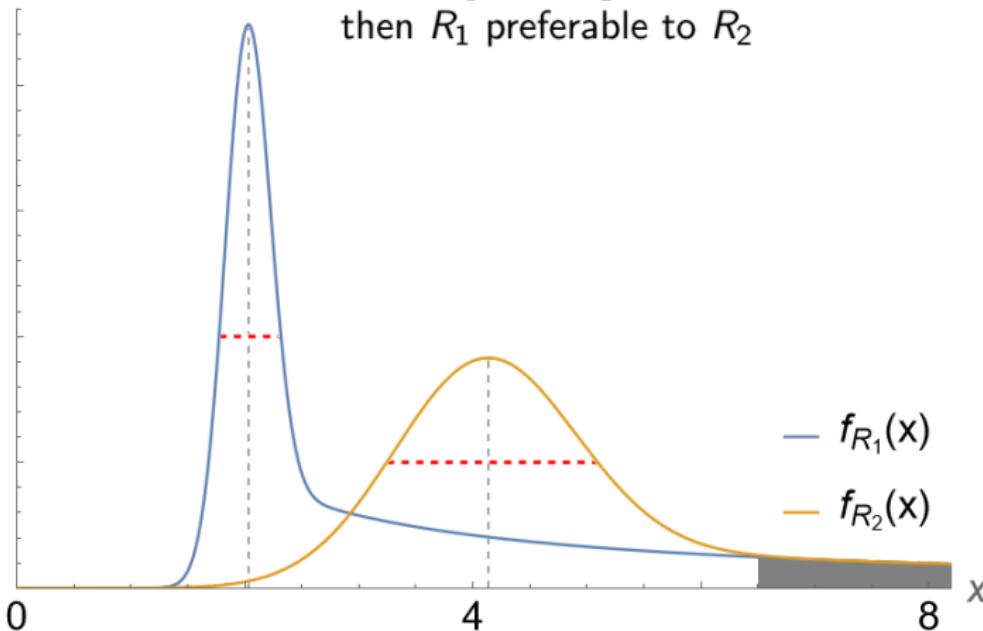
Aleatoric vs. epistemic uncertainty – Another example

- R_1 and R_2 ... two risks
- Identical tail
- $\mathbb{E}[R_1] < \mathbb{E}[R_2]$
- $\text{Variance}[R_1] < \text{Variance}[R_2]$



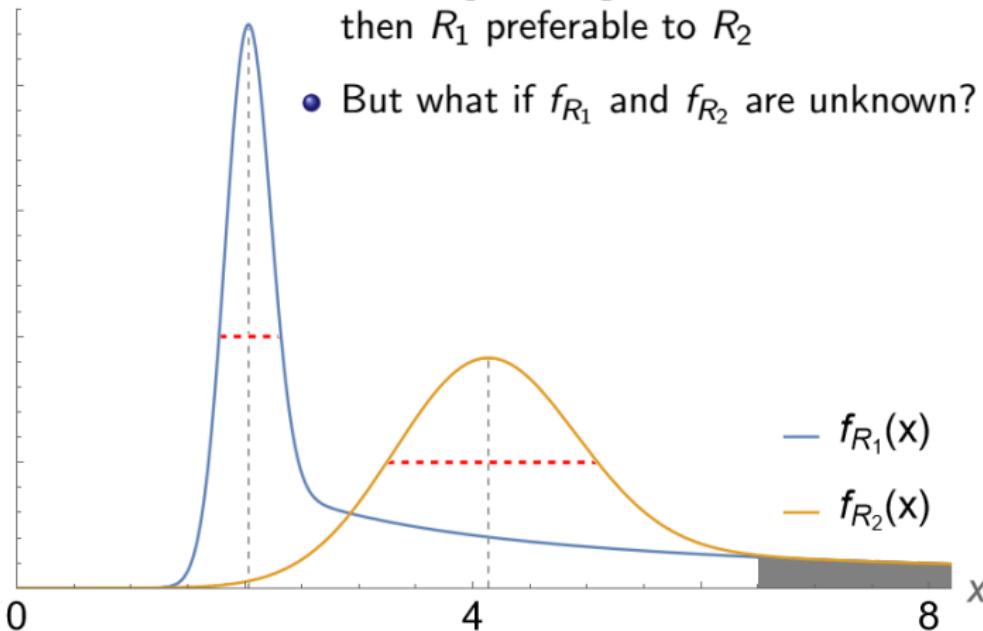
Aleatoric vs. epistemic uncertainty – Another example

- So if f_{R_1} and f_{R_2} are known, then R_1 preferable to R_2



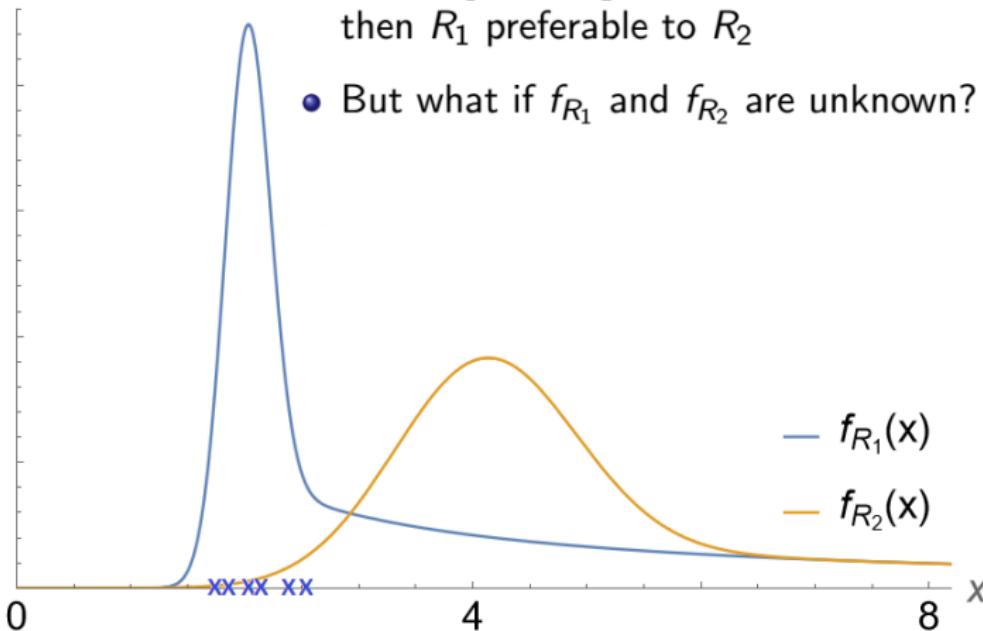
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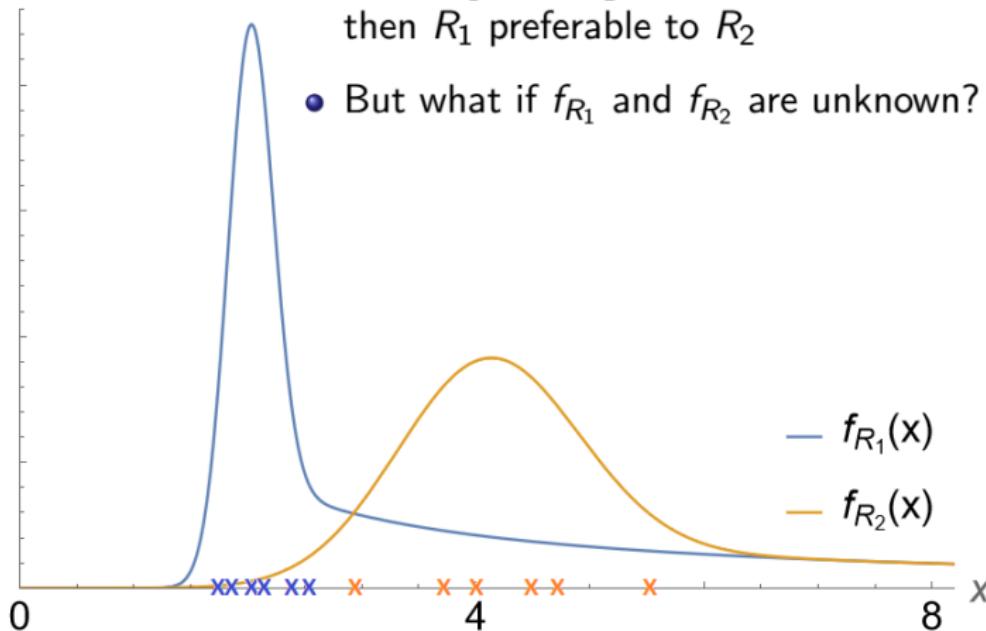
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Model and parameter uncertainty

Estimating F through a parametric approach

- 1) A parametric model is assumed ... $F = F_\theta$
- 2) Data $\Rightarrow \hat{\theta}$... estimated parameter (e.g. maximum likelihood)
- 3) $\hat{F} = F_{\hat{\theta}}$... estimated distribution

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\Rightarrow Two sources for $F \neq F_{\hat{\theta}}$ (epistemic uncertainty)

- $F \neq F_\theta$... wrong model (model uncertainty)
- $\theta \neq \hat{\theta}$... wrong parameter (parameter uncertainty)

Capital calculation

Setting

- Potential future loss ... $X \sim F$
- Backed with some capital ... c
- “Solvency” ... if $X \leq c$

Target solvency probability $\beta \in (0, 1]$

- Target capital ... $c_\beta = F^{-1}(\beta) = \text{VaR}_\beta[X]$
- Solvency probability ... $\mathbb{P}(X \leq c_\beta) = \beta$

Capital calculation under parameter uncertainty

With parameter uncertainty

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- θ ... unknown parameter

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- $F_{\hat{\theta}} \Rightarrow \hat{c}_\beta$... calculated capital

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Resulting solvency probability?

- In general ... $\mathbb{P}(X \leq \hat{c}_\beta) \neq \beta$

... see e.g. Gerrard and Tsanakas (2011), Fröhlich and Weng (2015)

Expected solvency probability

Performance measure? Before/after data realized

- After ... Ex-post capital $\hat{c}_\beta \Rightarrow \mathbb{P}(X \leq \hat{c}_\beta)$
- Before ... Ex-ante capital $\hat{C}_\beta \Rightarrow \mathbb{P}(X \leq \hat{C}_\beta)$

“Expected solvency probability” $\mathbb{P}(X \leq \hat{C}_\beta) \dots \begin{cases} \neq \beta & \text{in general} \\ < \beta & \text{in relevant cases} \end{cases}$

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Why?

- Remember ... Data $\Rightarrow \hat{\theta} \Rightarrow F_{\hat{\theta}} \Rightarrow \hat{c}_\beta$
- \Rightarrow Ignores parameter uncertainty

Capital calculation under parameter uncertainty

Bayesian approach ... Unknown parameter θ treated as random

- Prior $p(\theta)$ \Rightarrow Posterior $p(\theta | \text{Data})$
- $\tilde{F} = \int_{\theta} F_{\theta} \cdot p(\theta | \text{Data}) d\theta$... Predictive distribution for X
- $\tilde{F} \Rightarrow \tilde{c}_{\beta}$ and \tilde{C}_{β} ... Ex-post and ex-ante capital

Expected solvency probability $\mathbb{P}(X \leq \tilde{C}_{\beta})$?

With parameter uncertainty only

Gerrard and Tsanakas (2011)

- “Transformed location-scale families” ... e.g. $X = g(\theta_1 + \theta_2 Z)$
 - “Objective” prior ... $p(\theta_1, \theta_2) \propto \theta_2^{-1}$
- ⇒ $\mathbb{P}(X \leq \tilde{C}_\beta) = \beta$... independently of the unknown parameters

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In my paper ... Extension to a more general framework, for instance

- Sample space of data and X ... can depend on θ_1, θ_2
- Data ... not necessarily mutually independent

The Pareto and Pareto-type cases

- Loss $X \sim \text{Pareto}(\alpha, t) \dots 1 - F(x) = (\frac{t}{x})^\alpha$
 - Data ... n largest values of a Pareto random sample of size $k \geq n$
- $\Rightarrow \tilde{c}_\beta = \exp \{u_1(\text{data}) + u_2(\text{data}) \cdot g(k, n, \beta)\}$

So if the loss X is exactly Pareto but with unknown parameters

$\Rightarrow \mathbb{P}(X \leq \tilde{C}_\beta) = \beta \dots$ independently of the unknown parameters

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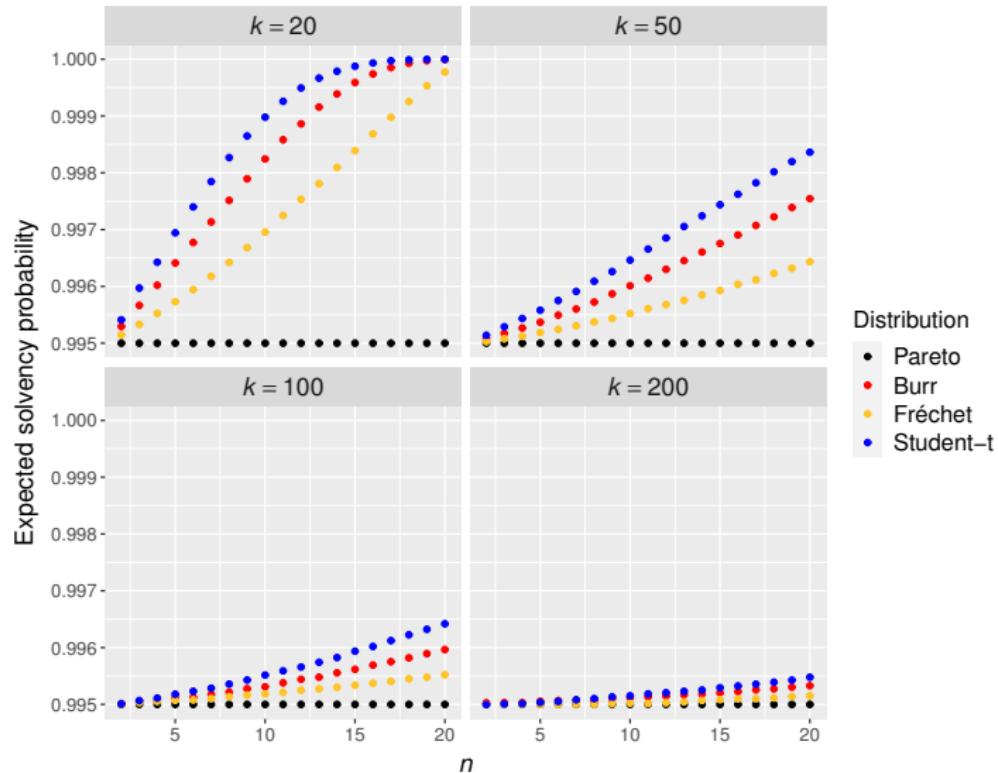
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$$\Rightarrow \mathbb{P}(X \leq \tilde{C}_\beta) = \beta \dots \text{independently of the unknown parameters}$$

What if not exactly Pareto?

- “Pareto-type” tail ... $1 - F(x) = \ell(x) \cdot (\frac{t}{x})^\alpha$
- $\lim_{x \rightarrow \infty} \ell(x) = 1$

Illustration



Introduction

A note on risk assessment under parameter uncertainty.

Predictive distributions under “objective” prior \Rightarrow very conservative ...
For instance

- $X \sim \text{Pareto}(\alpha, 1)$
- Data $\Rightarrow \hat{\alpha}$... maximum likelihood

Then

- Estimated distribution ... $\hat{F}(x) = 1 - x^{-\hat{\alpha}}$ (Pareto)
- Predictive distribution ... $\tilde{F}(x) = 1 - \left(1 + \frac{\ln(x^{\hat{\alpha}})}{n}\right)^{-n}$ (Log-Pareto)

Too conservative?

Joint solvency probability

Setting ... Risks $X_1 \perp X_2$

- $X_i \sim \text{Pareto}(\alpha_i, 1)$ for both i
- Allocation of a fixed total capital ... $c = c_1 + c_2$
- Joint solvency probability ... $\psi(c_1) = F_1(c_1) \cdot F_2(c - c_1)$

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Possible criterion ... Maximize $\psi(c_1)$

$$\text{Best allocation} \Rightarrow c_1^* = \arg \max_{c_1 \in [0, c]} \psi(c_1)$$

Parameter uncertainty ... What if α_1, α_2 are unknown?

Estimated vs. predictive distributions

Two approaches

- With estimated distributions ... Data $\Rightarrow \hat{F}_1, \hat{F}_2 \Rightarrow \hat{c}_1^*$
- With predictive distributions ... Data $\Rightarrow \tilde{F}_1, \tilde{F}_2 \Rightarrow \tilde{c}_1^*$

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For one particular dataset

- Resulting joint survival probabilities ... $\psi(\hat{c}_1^*) \leq \psi(\tilde{c}_1^*)$

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- Resulting joint survival probabilities ... $\psi(\hat{c}_1^*) \leq \psi(\tilde{c}_1^*)$

Performance criterion? Before data realized

- Ex-ante joint solvency probabilities ... $\psi(\hat{C}_1^*)$ and $\psi(\tilde{C}_1^*)$

Estimated vs. predictive distributions

In short

- $\mathbb{E}[\psi(\tilde{C}_1^*)] > \mathbb{E}[\psi(\hat{C}_1^*)]$
- $\text{Variance}[\psi(\tilde{C}_1^*)] < \text{Variance}[\psi(\hat{C}_1^*)]$
- Less data \Rightarrow Larger improvement

So by accounting for parameter uncertainty, **predictive distributions** here in general lead to better resources (capital) allocations than estimated distributions

Conclusion

Key points

- Epistemic uncertainty shall be taken into account when managing risk
- Parameter and model uncertainty are two types of epistemic uncertainty
- Bayesian approach can help for both capital calculation and resource (capital) allocation problems

... find details in my papers!

Thanks for your attention.

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References I

- Fröhlich, A. and Weng, A. (2015). Modelling parameter uncertainty for risk capital calculation. *European Actuarial Journal*, 5(1):79–112.
- Gerrard, R. and Tsanakas, A. (2011). Failure probability under parameter uncertainty. *Risk Analysis: An International Journal*, 31(5):727–744.