An Analysis of the Capital Structure of an Insurance Company

by

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Abstract

This paper attempts to analyze the capital structure of an insurance company in a way that: (1) views the insurance company as an ongoing enterprise; and (2) allows for the stochastic nature of insurance business. A model is developed. This model is used to analyze the effect of uncertainty in the loss reserves, the underwriting cycle and the cost of insurance regulation to the consumer. The paper considers both the investor's and the regulator's point of view.

The research for this paper was supported by a grant from the Actuarial Education and Research Fund

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1. Introduction

An insurance company is in the business of transferring risk. It does this by accepting premium from policyholders and paying claims. It can happen that the premium collected is less than the total amount paid for claims. If this is the case, the insurer is expected to pay for the claims from the capital¹ of the insurance company.

This paper addresses the following question.

How much capital will be invested in a given insurance company?

The owners of (or investors in) the insurance company are concerned with the yield and the safety of their investment. The money they invest in the insurance company must be competitive with respect to the yield and safety of alternative investments.

The insurance regulator has a vital interest in this question. The concern is that the insurance company have enough money to fulfill its obligations to the policyholders when loss experience is worse than expected.

A deterministic analysis of the capital structure of an insurance company might proceed as follows.

^{1.} We shall use the terms "capital" and "surplus" interchangably to represent the owner's equity in the insurance company.

Let: P = risk premium (or expected loss); L = security (or profit) loading; U = initial surplus; and i_u = interest rate earned on the surplus.

The expected rate of return on the owner's equity, i, satisfies the following equation:

$$U \cdot i = P \cdot L + U \cdot i_{11}, \qquad (1.1)$$

Fix P, L and i_u. It is easily seen that lowering U will increase the rate of return, i. There are two forces which limit how low U will go. First, the rate of return may get sufficiently high to attract more capital. For example let:

> P = \$20,000,000; L = .025; and i_u = .06.

Suppose the competitive rate of return is found to be i = 12%, we can solve Equation 1.1 for U = \$8,333,333. If the surplus were to go less than \$8,333,333 then new capital would be attracted to the insurance company. Conversely, if the surplus were to go above \$8,333,333 the owners could invest the excess surplus elsewhere and obtain a greater return on their investment.

A second limiting force is that of regulation. Regulators are interested in assuring that the insurance company can fulfill its obligation to the policyholders. Putting a lower bound on U will

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help accomplish this purpose. However, it should be pointed out that this action is not without cost to the policyholders. Suppose, in the above example, the regulator decides to require a surplus of \$9,333,333. If the competitive rate of return remains at 12%, the insurance company will be forced to raise its profit loading, L. Solving Equation 1.1 gives L = .028. Raising U by \$1,000,000 will cost the policyholders \$60,000.

While this analysis captures some essential points of insurance company operations, there are many other factors that should be considered. These factors include the following.

- 1. An insurance company is an ongoing operation.
- 2. The amount paid for claims varies from year to year.
- 3. The insurance industry is very competitive. The profit loading varies from year to year in a fashion described as the "underwriting cycle."
- 4. The ultimate claim cost is not determined at the end of the policy year. The result is uncertainty in the liabilities, and hence the surplus of the insurance company.

This paper analyzes the effect these factors will have on the capital structure of an insurance company. The analysis will consider the same questions as the deterministic analysis given above, namely: what surplus will give a competitive rate of return

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to the insurance company owners; and what is the cost to policyholders of minimum surplus regulation?

We begin with a model which describes how claim amounts vary.

2. The Collective Risk Model

We shall use the collective risk model to describe the incurred losses, X_t in year t. This model assumes separate claim severity distributions and claim count distributions for each line of insurance written by the insurer. We shall use the version of the model described by Heckman and Meyers [1983] and Meyers and Schenker [1983].

This version of the model can be described by the following algorithm.

- 1. Select β at random from an inverse gamma distribution with E[1/ β] = 1 and Var[1/ β] = b
- 2. For each line of insurance, k, do the following.
 - 2.1 Select χ at random from a gamma distribution with E[χ] = 1 and Var[χ] = c_k.
 - 2.2 Select a random number of claims, N, from a Poisson distribution with mean $\chi \cdot \lambda_{\nu}$.
 - 2.3 Select N claims at random from the claim severity distribution for line of insurance k.
- 3. Set X_t equal to the sum of all claims selected in step 2, multiplied by ρ .

The parameter c_k , called the contagion parameter, is a measure of

uncertainty in our estimate of the expected claim count, λ_k , for line k. The parameter b, called the mixing parameter, is a measure of uncertainty of the scale of the claim severity distributions. Note that the random scaling factor, β , acts on all claim severity distributions simultaneously.

For demonstration purposes we have selected a comparatively small insurance company writing a single line of insurance. The claim severity distribution is a Pareto distribution with c.d.f.

$$S(z) = 1 - (a/(a+z))^{\alpha}$$

where a = 10000 and α = 2. Each claim is subject to a \$500,000 limit.

The expected number of claims, λ , is set equal to 2039.544. The parameters b and c are set equal to 0 and .04 respectively. The resulting risk premium for this insured is \$20,000,000.

Exhibit I shows the resulting aggregate loss distribution as calculated by the Heckman/Meyers algorithm [1983].

We will refer to this example as the ABC Insurance Company in what follows.

3. The Distribution of Surplus

We will view the insurance company as an ongoing operation. It collects premiums, pays claims, pays dividends to the owners (or stockholders). Occasionally, the owners will be required to contribute additional capital in order to maintain the surplus at a level specified by the regulator.

The financial status of an insurance company is usually measured at year end. Accordingly, a discrete treatment of financial results is assumed, i.e. the state of a company's finances will be calculated at time t = 0, 1, 2, ... where t is in years.

Let: P = risk premium (assumed constant for all years); $L_t = security loading for year t;$ $X_t = incurred loss during year t;$ $D_t = stockholder dividends paid at the end of year t;$ $R_t = additional capital contributed at the end of year t;$ $U_t = surplus at the end of year t; and$ $i_n = interest rate (assumed constant) earned on the surplus.$

Our model of insurance company operations can be described as follows. Given the surplus, U_{t-1} , define the random variable V_t by

$$V_{t} = U_{t-1} \cdot (1+i_{u}) + P \cdot (1+L_{t}) - X_{t}$$

Let Umax be the maximum surplus and Umin be the minimum surplus determined by the insurance company management and/or regulators. Then we define: $D_{t} = \max (V_{t} - Umax, 0);$ $R_{t} = \max (Umin - V_{t}, 0); \text{ and thus}$ $U_{t} = V_{t} - D_{t} + R_{t}.$

This model is similar to that described by Beard, Pentikäinen and Pesonen [1984, p. 215].

Let $F_t(v)$ be the c.d.f. for V_t . Let $M = Umax \cdot (1+i_u) + P \cdot (1+L_t)$. M represents the maximum value of V_t . $F_t(v) = 1$ for $v \ge M$.

Let u_t , d_t , and r_t represent the expected values of the surplus, U_t , the dividend, D_t and the additional contributed capital, R_t at time t respectively. We have:

 $u_{t} = \bigcup_{\substack{M \\ \text{Umin}}}^{\text{Umax}} v \cdot dF_{t}(v) + \text{Umax} \cdot (1 - F_{t}(\text{Umax})) + \text{Umin} \cdot F_{t}(\text{Umin});$ $d_{t} = \bigcup_{\substack{M \\ \text{Umax}}}^{\text{M}} (v - \text{Umax}) \cdot dF_{t}(v); \text{ and}$ $r_{t} = \bigcup_{\substack{-\infty \\ -\infty}}^{\text{Umin}} (\text{Umin} - v) \cdot dF_{t}(v).$

Note that $u_{t} + d_{t} - r_{t} = E[V_{t}]$.

The requirement that additional capital be contributed applies even when the surplus is negative. It is possible for reinsurance companies or guaranty funds to contribute money to raise the surplus to 0. Cummins [1986] and Meyers [1987b] discuss some ways to price this reinsurance. Also of interest are $p_t(D)$, the probability of receiving a dividend, and $p_t(R)$, the probability that additional capital will be required. These probabilities are calculated for each year t.

Some notes on the history of this operating strategy is in order. This dividend paying strategy originated in the risk theory literature in a paper by de Finetti [1957]. It has been discussed by Bühlmann [1970, p. 164] and Borch [1974, p. 225]. A more general version of this strategy has been discussed by Tapiero, Zuckerman and Kahane [1983]. They insert an additional level, Ulong, between Umin and Umax. When V_t goes above Ulong then the amount, V_t - Ulong, is put into long term investments.

Meyers [1986] addresses the same questions addressed by this paper with an operating strategy that does not require the contribution of additional capital.

4. Yield Rates

The yield rate of an investment is defined to be the interest rate at which the present value of the investments is equal to the present value of the returns.

Let T be the investor's time horizon. The investments consist of the initial surplus at time zero and the additional contributions to surplus at each time t. The returns consist of dividends payable at each time t, and the average surplus at time T. Any yield rate calculation must, of course, reflect the probability that the payments are actually made.

Let i be the yield rate. The yield rate must satisfy the following equation:

$$u_0 + \sum_{t=1}^{T} r_t \cdot (1+i)^{-t} = \sum_{t=1}^{T} d_t \cdot (1+i)^{-t} + u_T \cdot (1+i)^{-T}.$$

This equation can be solved for i by the Newton-Raphson method.

The methodology described above has been incorporated into a computer program called the "Insurer Surplus Model." Meyers [1987a] describes the mathematical techniques used in this program. This program makes repeated use of the Heckman/Meyers algorithm.

Let us now consider the case of ABC Insurance Company. We make the following (debatable) assumptions.

1. The investors in ABC Insurance Company are risk neutral.

 The investors in ABC Insurance Company can easily shift their capital investments to seek the highest rate of return.

Suppose that the regulators require a minimum surplus of \$6,000,000, and that the market/regulators allow a security loading of .025. Suppose further that $i_u = .06$ and the investors select a time horizon of T=25 years. The company management calculates the following yields for varying levels of initial surplus (= maximum surplus).

Table	1	
Surplus		Yield
\$12,000,000		10.80%
10,000,000		11.66
8,000,000		12.79

To continue our example, let us suppose that the yield on alternative investments is 12% for T=25. It is a consequence of the above assumptions that the investors in ABC Insurance Company will adjust the surplus until a 12% yield is obtained. Thoughtful trial and error quickly gives an initial (= maximum) surplus of \$9,330,000. Note the the yield does vary with the time horizon, T, selected. The output the Insurer Surplus Model for this initial surplus is given in the following table.

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Table 2 Insurer Surplus Model Standard Assumptions							
t	Pt(R)	rt	^u t	dt	pt(D)	۴t	Yield
1	0.14518	371690	8501385	2260106	0.62482	0.02500	11.36
2	0.20393	580225	8256856	1834837	C.54171	0.02500	11.57
3	0.22181	644846	8183506	1713608	Q, 51713	0.02500	11.72
4	0.22719	664276	8161486	1677306	0.50976	0,02500	11.80
5	0.22880	670109	B154875	1666409	0.50754	0.02500	11.85
6	0.22928	671860	8152891	1663137	0.50688	0.02500	11.88
7	0.22943	672386	8152295	1662155	0.50668	0.02500	11.91
8	0.22947	672544	8152116	1661860	0.50662	0.02500	11.92
9	0.22949	672591	8152062	1661772	0.50660	0.02500	11.94
10	0.22949	672605	B152046	1661745	0.50660	0.02500	11.95
11	0.22949	672610	8152041	1661737	0.50659	0.02500	11.96
12	0.22949	672611	8152040	1661735	0.50659	0.02500	11.96
13	0.22949	672611	8152040	1661734	0.50659	0.02500	11.97
14	0.22949	672611	8152039	1661734	0.50659	0.02500 (11.78
15	0.22949	672611	8152039	1661734	0.50659	0.02500	11.98
16	0.22949	672611	8152039	1661734	0.50659	0.02500	11.98
17	0.22949	672611	8152039	1661734	0.50659	0.02500	11.99
18	0.22949	672611	8152039	1661734	0.50659	0.02500	11.99
19	0.22949	672611	8152039	1661734	0.50659	0.02500	11.99
20	0.22949	672611	8152039	1661734	0.50659	0.02500	11.99
21	0.22949	672611	8152039	1661734	0.50659	0.02500	11.99
22	0.22949	672611	8152039	1661734	0.50659	0.02500	12.00
23	0.22949	672611	8152039	1661734	0.50659	0.02500	12.00
24	0.22949	672611	8152039	1661734	0.50659	0.02500	12.00
25	0.22949	672611	8152039	1661734	0.50659	0.02500	12.00

One does not need the Insurer Surplus Model to find the yield for T=1.

> $u_0 + r_1/(1+i) = (u_1+d_1)/(1+i) \Rightarrow$ $(1+i) \cdot u_0 = u_1 + d_1 - r_1 = E[V_1] \Rightarrow$ $i = E[V_1]/u_0 - 1.$

Now:

 $E[V_1] = u_0(1+i_u) + P \cdot L_1.$ $i = i_u + P \cdot L_1 / u_0.$ Thus: (4.1)

Note that Equation 4.1 can also be derived from Equation 1.1.

5. Uncertainty in Loss Reserves

The time t=0 does not have to be the date the insurance company begins operation. The old advertising jingle "Today is the first day of the rest of your life" applies also to insurance companies. Applying the above approach to an ongoing insurance company presents a special problem which is discussed here.

Probably the largest and most uncertain liability for a property and casualty insurance firm is the loss reserve. This creates uncertainty in the initial surplus, U₀. We attempt to model this by making the additional assumption:

 U_{Ω} has a normal distribution with known mean and variance.

The debate concerning the variability of loss reserves has taken on new life within the last few years. Publications by the Casualty Actuarial Society Committee on the Theory of Risk [1986], De Jong and Zehnwirth [1983] and Taylor [1986] deal with this problem extensively. Even so, the author considers the problem far from being solved.

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In our example, the ABC Insurance Company, we will use \$1,790,035 as the standard deviation of the loss reserve i.e. the initial surplus. This figure was derived from the following assumptions.

1. The claim severity distribution is known.

- Claims are paid out over a period of eight years. The paid to ultimate ratios are .05, .20, .40, .60, .75, .90, .96 and 1.00 respectively.
- 3. The smallest claims are settled first.

The details of this derivation are in the appendix.

Using the Insurer Surplus we calculate that a value of \$9,340,000 for u_O and Umax will result in a yield of 12% if all other inputs remain the same. Table 3 contains the output.

Table 3 Insurer Surplus Model Uncertain Initial Surplus

t	Pt(R)	۲t	^u t	dt	Pt(D)	۴t	Yield
1	0.16524	458453	8444587	2414266	0.61053	0.02500	11.35
2	0.20838	596836	8244475	1803622	0.53504	0.02500	11.61
3	0.22284	648744	8184631	1703257	0.51494	0.02500	11.74
4	0.22722	664559	B166623	1673646	0.50893	0.02500	11.81
5	0.22854	669323	B161202	1664741	0.50712	0.02500	11.86
6	0.22893	670757	8159570	1662061	0.50658	0.02500	11.89
7	0.22905	671188	8159079	1661254	0.50641	0.02500	11.91
8	0.22909	671318	8158931	1661011	0.50636	0.02500	11.93
9	0.22910	671357	8158886	1660938	0.50635	0.02500	11.94
10	0.22910	671369	8158873	1660916	0.50634	0.02500	11.95
11	0.22910	671373	8158869	1660909	0.50634	0.02500	11.96
12	0.22910	671374	8158868	1660907	0.50634	0.02500	11,97
13	0.22910	671374	8158867	1660907	0.50634	0.02500	11,97
14	0.22911	671374	8159867	1660906	0.50634	0.02500	11.98
15	0.22911	671374	8156867	1660906	0.50634	0.02500	11.98
16	0.22911	671374	8158867	1660906	0.50634	0.02500	11.98
17	0.22911	671374	8158867	1660906	0.50634	0.02500	11.99
18	0.22911	671374	8158867	1660906	0.50634	0.02500	11.99
19	0.22911	671374	8158867	1660906	0.50634	0.02500	11.99
20	0.22911	671374	8158867	1660906	0.50634	0.02500	11.99
21	0.22911	671374	B158867	1660906	0.50634	0.02500	12.00
22	0.22911	671374	8158867	1660906	0.50634	0.02500	12.00
23	0.22911	671374	8158867	1660906	0.50634	0.02500	12.00
24	0.22911	671374	8158867	1660906	0.50634	0.02500	12.00
25	0.22911	671374	8158867	1660906	0.50634	0.02500	12.00

This example suggests that the uncertainty in loss reserves has little effect on surplus levels from the investor's point of view. More will said about this later.

6. The Underwriting Cycle

We now consider the case when the security loading varies from year to year in a cyclic maner. This is a well established phenomenon in casualty insurance which is felt, at least by the author, to be caused by intense competition from within the insurance industry. Berger [1986] proposes a model whereby the underwriting cycle results from (1) the desire to maximize profits and (2) aversion to

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bankruptcy.

To model the underwriting cycle we assume that

$$L_t = L_0 + A \cdot \sin(\omega \cdot (t-1) + \phi).$$

This is a special case of the AR(2) model considered by Beard, Pentikainen and Pesonen [1984, p. 202 and p. 388] for cyclic variation.

To demonstrate the effects of the underwriting cycle on the ABC Insurance Company we set $L_0 = .025$, A = .02394 and $\omega = \pi/4$. These parameters will produce an eight year cycle with a reasonable amount of variation..

We first consider what happens when we catch the cycle on the way up. If we set $\phi \approx 0$ along with the assumptions stated immediately above and in Section 4, we calculate that a value of \$10,600,000 for u₀ and Umax will result in a yield of 12%. The results of the Insurer Surplus Model for this case are in Table 4.

Table 4 Insurer Surplus Model Underwriting Cycle on the Way Up

t	Pt ^(R)	rt	ut	dt	Pt(D)	Ľŧ	Yield
1	0.09049	215329	9643356	2307973	0,63153	0.02500	10.72
2	0.13201	355329	9381137	2034662	0.56686	0.04193	12.50
3	0.14135	387016	9323605	1986147	0.55316	0.04894	13.58
4	0.15099	418743	9258197	1882081	0.53551	0.04193	13.85
5	0.17115	486594	9124016	1676267	0.49982	0.02500	13.50
6	0.19766	579358	8956122	1456180	0.45728	0.00807	12.85
7	0.21712	649582	8838445	1325894	0.42894	0.00106	12.23
8	0.21700	648298	8839775	1338762	0.42969	0.00807	11.89
9	0.19766	577898	8958231	1489828	0.45896	0.02500	11.87
10	0.17253	490525	9118200	1706563	0.49970	0.04193	12.07
11	0.15641	436580	9224283	1856320	0.52740	0.04894	12.29
12	0.15683	438201	9220174	1834282	0.52578	0.04193	12.40
13	0.17355	494854	9108970	1659268	0.49611	0.02500	12.37
14	0.19868	582988	8949991	1449991	0.45582	0.00807	12.23
15	0.21755	651133	8835919	1323475	0.42835	0.00106	12.08
16	0.21718	648916	8838742	1337734	0.42944	0,00807	11.98
17	0.19773	578126	8957820	1489373	0.45886	0.02500	11.97
18	0.17256	490607	9118042	1706367	0.49966	0.04193	12.03
19	0.15642	436610	9224224	1856242	0,52739	0.04894	12.11
20	0.15684	438213	9220151	1834253	0.52577	0.04193	12.15
21	0.17355	494859	9108961	1659258	0.49610	0.02500	12.14
22	0.19868	582990	8949987	1449987	0.45582	0.00807	12.09
23	0.21755	651134	8835917	1323473	0.42835	0.00106	12.04
24	0.21718	648916	8838741	1337733	0.42944	0.00807	12.00
25	0.19773	578126	8957820	1489372	0.45886	0.02500	12.00

Let us next consider what happens when we catch the cycle on the way down. If we set $\phi = \pi$ along with the assumptions stated immediately above and in Section 4, we calculate that a value of \$7,975,000 for u₀ and Umax will result in a yield of 12%. The results of the Insurer Surplus Model for this case are in Table 5.

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Table 5 Insurer Surplus Model Underwriting Cycle on the Way Down

t	Pt(R)	rt	чt	ďt	p _t (D)	Ľt	Yield
1	0.22753	636215	7380114	2209600	0.61761	0.02500	12.27
2	0.30233	922729	7214041	1693095	0.52639	0.00807	10.41
3	0.32807	1026142	7158032	1536263	0.49623	0.00106	9.20
4	0.32173	1000594	7171898	1577697	0.50374	0.00807	9.00
5	0.29409	891738	7232678	1761271	0.53683	0.02500	9.64
6	0.26412	778592	7299746	1983999	0.57398	0.04193	10.64
7	0.24901	723287	7334002	2105747	0.59320	0.04894	11.49
8	0.25610	748834	7317772	2043617	0.58400	0.04193	11.95
9	0.28192	844836	7259561	1842113	0.55153	0.02500	12.01
10	0.31286	964985	7191167	1630440	0.51409	0.00807	11.81
11	0.33011	1034457	7153645	1524718	0.49389	0.00106	11.55
12	0.32211	1002136	7171065	1575421	0.50329	0.00807	11.42
13	0.29416	892006	7232524	1760810	0.53674	0.02500	11.47
14	0.26413	778637	7299718	1983909	0.57396	0.04193	11.66
15	0.24901	723295	7333997	2105730	0.59319	0.04894	11.86
16	0.25610	748835	7317771	2043614	0.58400	0.04193	11.98
17	0.28192	844836	7259561	1842112	0.55153	0.02500	12.00
18	0.31286	964985	7191166	1630440	0.51409	0.00807	11.94
19	0.33011	1034457	7153645	1524718	0.49389	0.00106	11.85
20	0.32211	1002136	7171065	1575421	0.50329	0.00807	11.81
21	0.29416	892006	7232524	1760810	0.53674	0.02500	11.82
22	0.26413	778637	7299718	1983909	0.57396	0.04193	11.88
23	0.24901	723295	7333997	2105730	0.59319	0.04894	11.95
24	0,25610	748835	7317771	2043614	0.58400	0.04193	11.99
25	0.28192	844836	7259561	1842112	0.55153	0.02500	12.00

7. Ruin Theory

Thus far our assumption has been that the investors in an insurance company will adjust the surplus so that the expected yield will be constant. An alternative to this assumption is provided by ruin theory. Ruin theory² makes the assumption that the investors in an insurance company will adjust the surplus so that the probability of insolvency (i.e. the probability of ruin) will remain constant. In this section we shall demonstrate that the two assumptions imply quite different results.

2. See, for example, Beard, Pentikainen and Pesonen [1984 ch. 4].

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It is sufficient to consider the probability of ruin for a one year time span. Let ϵ be the selected probability of ruin. We have:

$$Pr\{U_1 < 0\} = \epsilon$$
 if and only if $u_0(1+i_u) + P(1+L_1) = x_{1-\epsilon}$,

where $x_{1-\epsilon}$ is the $1-\epsilon^{th}$ percentile of the random loss X. If ϵ is fixed, it can be seen that a reduction in L_1 should be accompanied by a corresponding increase in u_0 , and conversely an increase in L_1 should be accompanied by a corresponding decrease in u_0 .

Equation 4.1 indicates the opposite behavior. If i is fixed, it can be seen that L_1 and U_0 move in the same direction. This behavior also holds in the multiyear analysis of the underwriting cycle. If the cycle is on the way down, Umax also goes down and the insurance company's surplus is reduced. The opposite happens when the cycle is on the way up.

The two assumptions have different implications when we consider uncertainty in loss reserves. It was demonstrated in the example above that uncertainty in the loss reserves has little effect on the surplus. The surplus raises from \$9,330,000 to \$9,340,000. Suppose we are satisfied with the probability of ruin for the standard assumptions (Table 2). Using the Insurer Surplus Model with Umin = 0, we calculate that the probability of ruin after one year is .0152. If the standard deviation of the loss reserve is \$1,790,035, as in Table 3, it requires a surplus of \$10,045,000 to maintain the probability of ruin of .0152 for the first year.

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8. The Cost of Regulation

It is the regulator's job to impose standards which keep the number of insolvencies to a minimum. One way of doing this is to impose a minimum surplus so that the probability of ruin is acceptably low. It was demonstrated in the last section that such a regulatory strategy may not be in accordance with the wishes of insurance company owners.

The owners don't have any choice in the matter. The regulators set the standards and the insurance companies comply with them. A higher minimum standard will result in a higher level of surplus in the industry as a whole, and a higher profit loading will be demanded. The purpose of this section is find this additional cost of solvency regulation to insurance consumers.

Let us consider the example in Table 2. We will vary the minimum surplus and calculate the security loading that will result in a yield rate of 12% after 25 years. The results are in Table 6.

Table 6 The Cost of Regulation

Minimum Surplus	Security Loading (%)	Security Loading (\$)	Additional Security Loading
\$6,000,000	2.500%	\$500,000	<u></u>
7,000,000	2.583	516,600	\$16,600
8,000,000	2.673	534,600	18,000
9,000,000	2.767	553,400	18,800
10,000,000	3.000	600,000	46,600
11,000,000	3.300	660,000	60,000

Note that if the minimum surplus goes above \$9,330,000 the minimum surplus becomes the maximum surplus, and the security loading can

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be obtained by solving Equation 1.1.

The changes in the market conditions brought on by increasing the minimum surplus are clearly more complex than is assumed by the above example. However, this may be an indication that the cost of regulation is small if the minimum surplus is not too high.

Concluding Remarks

This paper has attempted to analyze the capital structure of an insurance company in a way that:

(1) viewed the insurance company as an ongoing enterprise; and

(2) allowed for the stochastic nature of insurance business.

When one attempts a simple one year deterministic analysis as was done in the introduction, it is possible to comprehend the implications instantly. However, when given a complex computer program like the Insurer Surplus Model, the best one can do is try some examples and draw tentative conclusions. This paper represents one such attempt. The main conclusions are listed below.

 The underwriting cycle has a major effect on the amount of capital that will be invested in an insurance company. For example, an insurance company should lower its surplus in the down part of the cycle. In our examples, the goal was to obtain an expected yield of 12% over a 25 year period. One should not view this strategy as being shortsighted.

The uncertainty in loss reserves has little effect from the

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investor's point of view. However, it can have a substantial effect from the regulator's point of view.

3. Whether the investors like it or not, the regulators may require a minimum surplus. If this minimum is below what the investors would voluntarily allow, the cost to the policyholders is relatively small. As this regulatory minimum increases, the cost to the policyholders becomes substantial.

There are several items that should enter this analysis, but did not. A discussion of some of these items follows.

We assumed that that the investor would seek the same expected yield in all circumstances. One could reasonably argue that the investor should seek a higher yield when the surplus is low because of the increased variability of the return. This is debatable. It is unlikely that the investor would invest all his/her assets in a single enterprise, and so the investor's risk aversion should not be much of a factor. However, the author would like to keep the debate open.

The issue of asset risk has been omitted from this entire discussion. It could very well be as important as any of the items mentioned above. Any analysis of asset risk must include strategies for asset/liability matching. A good place for casualty actuaries to start would be the paper "Duration" by Ronald E. Ferguson [1983]. Further research needs to be done in order to integrate asset risk into the above approach for analyzing the capital structure of an insurance company.

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	Exhibit I		Collectiv	e Ŕisk M	odel	ABC I	nsurance	Company
Line	Expected Loss	Claim Se Distrib	verity ution	Contagi Paramet	on er	Claim Cou Mean	nt Clai St	m Count d Dev
1	20000000	Pare	to	0.0400		203	41	0.401
Mixing Aggrega Aggrega	parameter te mean te std dev	0.0000 20000000 4147667						
Aggr	egate	Entry	Cumulati	ve	Exces	s Pure	Excess	Pure
Loss	Amount	Ratio	Probabil:	ity	Pre	mium	Premium	Ratio
10000	000-00	0.5000	0.001	8	100013	41.98	0.50	01
11000	000.00	0.5500	0.005	6	90047	91.48	0.45	02
12000	000.00	0.6000	0.014	3	80142	34.53	0.40	07
13000	000.00	0.6500	0.031	5	70362	73,36	0.35	18
14000	000.00	0.7000	0.060	3	60812	25.87	0.30	41
15000	000.00	0.7500	0.105	5	51629	BO.37	0.25	81
16000	00.00	0.8000	0.166	7	42977	50.83	0.21	49
17000	000.00	0.8500	0.2440	2	35019	85.98	0.17	51
18000	00.00	0.9000	0.3333	3	27897	31.74	0.13	95
19000	00.00	0.9500	0.429	5	21708	38.94	0.10	85
20000	00.00	1.0000	0.5268	3	16491	58.71	0.08	25
21000	000.00	1.0500	0.619	5	122290	02.24	0.06	11
22000	000.00	1.1000	0.703	3	88523	20.10	0.04	43
23000	00.00	1.1500	0.7756	5	62570	08.80	0.03	13
240000	00.00	1.2000	0.8350)	4320	75.05	0.02	16
250000	00.00	1.2500	0.882	L	2916	55.47	0.01	46
260000	00.00	1.3000	0.9180)	1925	75.14	0.00	76
270000	00.00	1.3500	0.9444	1	1244	71.15	0.00	62
280000	00.00	1.4000	0.963	2	788:	4.92	0.00	39
290000	00.00	1.4500	0.9762	2	4892	28.01	0.00	24
300000	00.00	1.5000	0.9849	2	2980	02.53	0.00	15
310000	000.00	1.5500	0.9907		178	24.96	0.00	09
320000	000.00	1.6000	0.9943	5	1047	76.32	0.00	05
330000		1.6500	0.9966		60:	54.93	0.00	33
340000	00.00	1.7000	0.9980	2	344	13.74	0.00	52
30000		1 8000	0.7787		192		0.00	71
370000		1 0500	0.9994		100		0.00	21
370000		1 9000	0.7776	2	20	0 04	0.00	00
30000	00.00	1 9500	0.7770		دد.	4 40	0.00	20
370000		2 0000	0.7777	,	10	37.00 34 10	0.000	50
410000		2.0500	1 0000	,	2	10.10	0.00	00
410000	////	2.0000	1.0000		•	TT . J/	0.00	<i></i>

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Appendix The Variability of Loss Reserves

In Section 5 we studied how the variability of loss reserves affected the surplus. We assumed that the loss reserves were normally distributed with a standard deviation of \$1,790,035. In this appendix we show how the standard deviation was derived.

Three assumptions were made.

- 1. The claim severity distribution is known.
- 2. Claims are paid out over a period of eight years. The paid to ultimate ratios are .05, .20, .40, .60, .75, .90, .96 and 1.00 respectively.
- 3. The smallest claims are settled first.

We used the Pareto distribution for the claim severity. The c.d.f. is given by:

$$S(z) = 1 - (a/(a+z))^{\alpha}$$

with a = 10000 and α = 2.

Let: c(i) = maximum claim size settled in the ith prior year; and <math>n(i) = number of claims remaining to be settled.

We have: c(i) ∫ z•dS(z) <u>0</u> = paid to ultimate ratio for prior year i; and E[Z]

 $n(i) = (1 - S(c(i))) \cdot 2039.544.$

Recall that 2039.544 is the annual expected number of claims for the ABC Insurance Company.

We then calculate the following values.

<u>i</u>	c(i)	n(i)
1	2844	1236
2	7947	633
3	16754	285
4	32912	111
5	60172	41
6	154844	8
7	276340	2

For prior year i, n(i) claims are selected at random from the claim severity distribution, S(z), conditioned on each claim being above c(i), The loss reserve is the total amount generated by this process. The distribution of loss reserves can be calculated by CRIMCALC, a computer program for the Heckman/Meyers algorithm. Exhibit II gives the output for CRIMCALC and Exhibit III shows that the distribution of loss reserves can be approximated by a normal distribution.

Reserve Risk

Line	Expected Loss	Claim Severity Distribution	Contagion Parameter	Claim Count Mean	Claim Count Std Dev
1	18991244	Prior year 1	-0.0008	1236.000	0.000
2	15991483	Prior year 2	-0.0016	633.000	0.000
3	12001336	Prior year 3	-0.0035	285,000	0.000
4	8018018	Prior year 4	-0.0090	111.000	0.000
5	4949249	Prior year 5	-0.0244	41.000	0.000
6	2131540	Prior year 6	-0.1250	8.000	0.000
7	803876	Prior year 7	-0,5000	2.000	0.000

Mixing par	ameter	0.0000
Aggregate	mean	62886746
Aggregate	std dev	1790035

Aggregate	Entry	Cumulative	Excess Pure	Excess Pure
Loss Amount.	Ratio	Probability	Premium	Premium Ratio
5600000.00	0.8905	0.0000	6886755.06	0.1095
56500000.00	0.8984	0.0001	6386768.85	0.1016
57000000.00	0.9064	0.0002	5886823.18	0.0936
57500000.00	0.9143	0.0006	5387011.46	0.0857
5800000.00	0.9223	0.0019	4887592.52	0.0777
58500000.00	0.9302	0.0050	4389195.07	0.0698
59000000.00	0.9382	0.0117	3893164,53	0.0619
59500000.00	0.9461	0.0252	3402048.32	0.0541
6000000.00	0.9541	0.0492	2920121.58	0.0464
6050000.00	0.9620	0.0881	2453739.46	0.0390
61000000.00	0.9700	0.1452	2011241.67	0.0320
61500000.00	0.9779	0.2219	1602207.69	0.0255
6200000.00	0.9859	0.3162	1236084.24	0.0197
62500000.00	0.9939	0.4229	920494.91	0.0146
63000000.00	1.0018	0.5341	659738.75	0.0105
43500000.00	1.0098	0.6415	453969.36	0.0072
64000000.00	1.0177	0.7375	299305.76	0.0048
64500000.00	1.0257	0.8175	188784.16	0.0030
4500000.00	1.0336	0.8796	113786.09	0.0018
65500000.00	1.0416	0.9246	65488.00	0.0010
6600000.00	1.0495	0.9552	35975.26	0.0006
66500000.00	1.0575	0.9748	18860.86	0.0003
6700000.00	1.0654	0.9865	9438.13	0.0002
67500000.00	1.0734	0.9931	4509.60	0.0001
68000000.00	1.0813	0.9967	2059.08	0.0000
68500000.00	1.0893	0.9985	900.03	0.0000
6900000.00	1.0972	0,9993	377.99	0.0000
69500000.00	1.1052	0.9997	153.68	0.0000
7000000.00	1.1131	0.9999	61.63	0.0000

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