APPROXIMATIONS TO RISK THEORY'S F(x, t)BY MEANS OF THE GAMMA DISTRIBUTION

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It seems that there are people who are prepared to accept what the numerical analyst would regard as a shockingly poor approximation to F(x, t), the distribution function of aggregate claims in the interval of time (o, t), provided it can be quickly produced on a desk or pocket computer with the use of standard statistical tables. The so-called NP (Normal Power) approximation has acquired an undeserved reputation for accuracy among the various possibilities and we propose to show why it should be abandoned in favour of a simple gamma function approximation.

Discounting encomiums on the NP method such as Bühlmann's (1974): "Everybody known to me who has worked with it has been surprised by its unexpectedly good accuracy", we believe there are only three sources of original published material on the approximation, namely Kauppi *et al* (1969), Pesonen (1969) and Berger (1972). Only the last two authors calculated values of F(x, t) by the NP method and compared them with "true" four or five decimal values obtained by inverting the characteristic function of F(x, t) on an electronic computer.

Briefly, the NP method for approximating F(x, t) consists of calculating y from the quadratic (NP2) or cubic (NP3) equation

$$\frac{x-t}{\sqrt{x_2}} = y + \frac{x_3/x_2^{3/2}}{3!} (y^2 - 1) + \frac{x_4/x_2^2}{4!} (y^3 - 3y) - \frac{x_3^2/x_2^3}{(3!)^2} (2y^3 - 5y)$$
(1)

where the kappas are the cumulants of $F(\cdot, t)$, and treating the result as a standardized Normal variate so that

$$F(x,t) \simeq \Phi(y) \equiv \frac{1}{\sqrt{2\pi}} \int_{-\pi}^{\pi} e^{-z^{2}/2} dz \qquad (2)$$

Berger (*loc. cit.*) found that the use of x_4 and the inclusion of the last two terms of the foregoing equation in y "does not generally produce better results than NP2". In our view, the necessity of solving a cubic equation and, possibly, choosing the appropriate root (Berger, 1972) removes the "second approximation" from the list of simple procedures.

Among the "short cut methods" of approximating F(x, t) tried by Bohman and Esscher in their classic 1963-64 paper was the gamma distribution with density

$$\frac{\Gamma}{\Gamma(\alpha)} e^{-y} y^{x-1} \qquad 0 \le y < \infty$$

so that

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$$F(t+z/\overline{x_2},t) \simeq \frac{1}{\Gamma(\alpha)} \int_{0}^{\alpha+i\sqrt{\alpha}} e^{-y} y^{\alpha-1} \equiv P(\alpha, \alpha+z/\overline{\alpha})$$
(3)

where the *P*-notation for the incomplete gamma ratio is now standard (see, e.g., Magnus *et al*, 1966) and α is to be determined from

$$\alpha = \frac{4}{\varkappa_3^2/\varkappa_2^3} \equiv \frac{4}{\gamma_1^2} \tag{4}$$

The joint authors reported that "the method has an astonishing accuracy in large parts of the field investigated" and one wonders why it has not been used more widely. The tables of Khamis-Rudert (1965) allow the approximation to be made with facility. It is mentioned, however, that what we write as P(a, x) is called by Khamis P(a, 2x); this must be watched when using the tables.

Let us therefore compare the published NP2 and NP3 approximations to F(x, t) with those obtained from (3) and (4). In the appended table the first four *t*-values come from Berger's (1972) Table 2, the next is from Pesonen (1969) and the last two are from Berger's (*loc. cit.*) Table 3. There are 38 values of I - F(x, t)shown in the Table and the gamma approximation (which is overloaded with decimals in the Table) is better than NP2 in 27 of them. It is better than NP3 in 27 also. What is more important is that the gamma approximation is better than NP2 in 9 of the 12 cases where deviations from the mean are 4, 5 or 6 standard deviations; the corresponding number among the dozen similar NP3 cases is

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also 9—but not the same 9/ Furthermore, the superiority of the gamma approximation does not seem to depend on the size of α , large values of which are supposed to ameliorate the accuracy of the NP method. Surely here is a case for discarding the Normal Power method altogether.

To conclude, it is mentioned that just as the NP method can be extended to provide stop loss premiums (Pesonen, 1969) the same is true of the gamma approximation. The stop loss premium at priority x can be shown to be

$$\int_{a}^{a} (u-x)d_{u} F(u,t) \propto \sqrt{\alpha x_{2}} \frac{q^{x}e^{-q}}{\Gamma(\alpha+1)} + (x-t) P(\alpha,q) - (x-t)$$

where $q = \alpha + \sqrt{\alpha/x_{2}} (x-t)$.

No calculations of this quantity were made as it was not thought that any different conclusions would have been drawn.

Individual claim distribution (Bohman-Esscher)	Negative binomial index	• t	$z = \frac{x - t}{\sqrt{x_2}}$	α
Non-industrial fire	ω	1000	0 I 2 3 4 6	2.7147
Non-industrial fire	20	1000	0 1 2 3 4 6	6.0741
Non-industrial fire	. ∞	100	0 I 2 3 4 6	0.27148
Non-industrial fire	20	100	o 1 2 3 4 6	0.32569
Life B	ω	1000	0 I 2 3 4 6	2.7056
Non-industrial fire	I	1000	0 I 3 5	0.9901
Non-industrial fire	I	100	0 I 3 5	0.5854

* The values in this panel were calculated by the author.

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$\alpha + z \psi \alpha$	I - F(x, t) "exact"	Gamma approx.	NP2	NP3
2.7147	.4265	.4193	.4228	.4131
4.3623	.1364	.1483	.1587	.1425
6.0100	.04523	.04481	.04938	.04497
7.6376	.01401	.01234	.01348	.01387
9.3052	.00352	.00319	.00333	.00428
12.6005	.00022	.00019	.00164	.00042
6.0741	.4476	.4460	.4472	4444
8.5387	.1502	.1535	.1587	.1509
11.0032	.03968	.03977	.04179	.04000
13.4678	.00892	.00849	.00881	.00920
15.9324	.00177	.001 58	.00157	.00195
20.8615	.00005	.00004	.00003	.00008
0.27148	·3743	.2639	.3129	.1641
0.79252	.0947	.1027	.1587	.0827
1.31355	.03450	.04783	.08152	.04827
1.83459	.01709	.02383	.04195	.03016
2.35563	.00893	.01232	.02156	.01967
3.39770	.00378	.00351	.00565	80000.
0.32569	.3801	.2805	.3226	.1795
0.89638	1006	.1083	.1587	.0827
1.46708	.03521	.04892	.07856	.0488
2.03777	.01680	.02350	0.3880	.0298
2.60846	.00855	.01168	.01907	.01897
3.74985	.00365	.00306	.00454	.00843
2.7056	.3992	.4191	.4227 *	.4194
4.3505	.1562	.1482	.1587	.1510
5.9953	.04569	.04483	.04947	.04531
7.6402	.01258	.01236	.01350	.01201
9.2851	.00281	.00320	.00334	.00291
12.5748	.00012	.00019	.00016	.00014
0.9901	.3671	.3672	.3805	.3593
1.9851	.1353	.1352	.1587	.1347
3.9752	.0184	.0184	.0229	.0194
5.9653	.0025	.0025	.0028	.0029
0.5854	.3448	.3299	.3540	.3040
1.3505	.1226	.1242	.1587	.1189
2.8807	.0198	.0213	.0297	.0238
4.4110	.0046	.0040	.0051	.0056

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